

# CS70: Lecture 34.

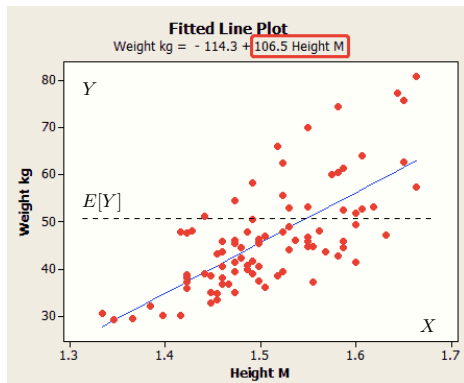
## Linear Regression (LR)

1. Motivation for Linear Regression (LR)
2. Minimum Mean Squared Error: Discussion
3. Covariance: Definition and Properties
4. Linear Regression (LR): Non-Bayesian vs. Bayesian (LLSE)
5. Derivation and Illustration

# Linear Regression: Motivation

Example 1: 100 people.

Let  $(X_n, Y_n) = (\text{height}, \text{weight})$  of person  $n$ , for  $n = 1, \dots, 100$ :



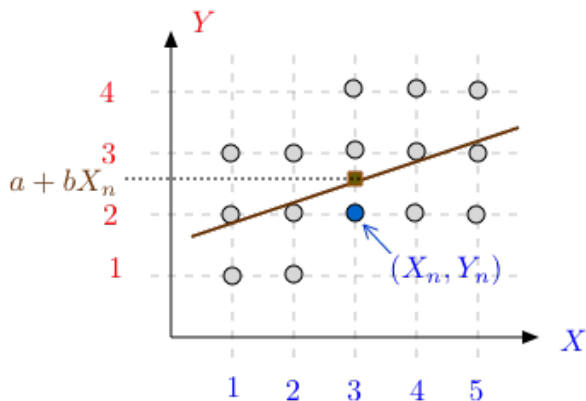
The blue line is  $Y = -114.3 + 106.5X$ . ( $X$  in meters,  $Y$  in kg.)

Best linear fit: [Linear Regression](#).

# Motivation

Example 2: 15 people.

We look at two attributes:  $(X_n, Y_n)$  of person  $n$ , for  $n = 1, \dots, 15$ :



The line  $Y = a + bX$  is the linear regression.

## Linear Regression: Discussion

If we want to guess the value of a random variable  $Y$ , and know nothing more than its distribution, what's our best guess?

Depends on how we measure the 'goodness' of our guess.

Say we use the **expected squared error between  $Y$  and our guess** as the "error" measure. Then? Answer is:  $E[Y]$ .

More precisely, the value of  $a$  that minimizes  $E[(Y - a)^2]$  is  $a = E[Y]$ .

**Proof:**

Let  $\hat{Y} := Y - E[Y]$ . Then,  $E[\hat{Y}] = 0$ . So,  $E[\hat{Y}c] = 0, \forall c$ . Now,

$$\begin{aligned} E[(Y - a)^2] &= E[(Y - E[Y] + E[Y] - a)^2] \\ &= E[(\hat{Y} + c)^2] \text{ with } c = E[Y] - a \\ &= E[\hat{Y}^2 + 2\hat{Y}c + c^2] = E[\hat{Y}^2] + 2E[\hat{Y}c] + c^2 \\ &= E[\hat{Y}^2] + 0 + c^2 \geq E[\hat{Y}^2]. \end{aligned}$$

Hence,  $E[(Y - a)^2] \geq E[(Y - E[Y])^2], \forall a$ . □

## Linear Regression: Discussion

Thus, if we want to guess the value of  $Y$ , we choose  $E[Y]$ .

Now assume we make some observation  $X$  related to  $Y$ .

How do we use that observation to improve our guess about  $Y$ ?

Idea: use a function  $g(X)$  of the observation to estimate  $Y$ .

The simplest  $g(X)$  is a constant that does not depend on  $X$ .

The next simplest function is linear:  $g(X) = a + bX$ .

What is the best linear function? That is our next topic.

(We can also consider a general function  $g(X)$ . Any guess on what is the best function to use? Answer:  $E[Y|X]$ .)

# Covariance

**Definition** The covariance of  $X$  and  $Y$  is

$$\text{cov}(X, Y) := E[(X - E[X])(Y - E[Y])].$$

**Fact**

$$\text{cov}(X, Y) = E[XY] - E[X]E[Y].$$

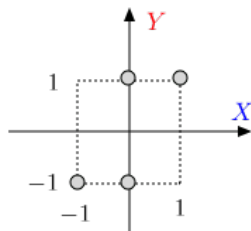
**Proof:**

$$\begin{aligned} E[(X - E[X])(Y - E[Y])] &= E[XY - E[X]Y - XE[Y] + E[X]E[Y]] \\ &= E[XY] - E[X]E[Y] - E[X]E[Y] + E[X]E[Y] \\ &= E[XY] - E[X]E[Y]. \end{aligned}$$

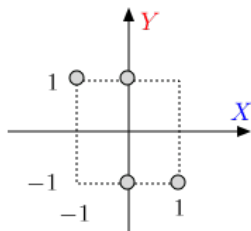


# Examples of Covariance

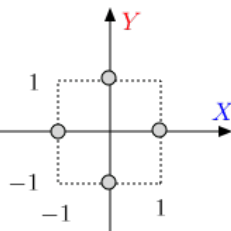
Four equally likely pairs of values



$$\text{cov}(X, Y) = 1/2$$



$$\text{cov}(X, Y) = -1/2$$



$$\text{cov}(X, Y) = 0$$

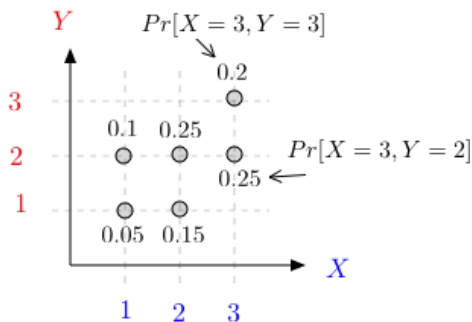
Note that  $E[X] = 0$  and  $E[Y] = 0$  in these examples. Then  $\text{cov}(X, Y) = E[XY]$ .

When  $\text{cov}(X, Y) > 0$ , the RVs  $X$  and  $Y$  tend to be large or small together.  $X$  and  $Y$  are said to be **positively correlated**.

When  $\text{cov}(X, Y) < 0$ , when  $X$  is larger,  $Y$  tends to be smaller.  $X$  and  $Y$  are said to be **negatively correlated**.

When  $\text{cov}(X, Y) = 0$ , we say that  $X$  and  $Y$  are **uncorrelated**.

## Examples of Covariance



$$E[X] = 1 \times 0.15 + 2 \times 0.4 + 3 \times 0.45 = 1.9$$

$$E[X^2] = 1^2 \times 0.15 + 2^2 \times 0.4 + 3^2 \times 0.45 = 5.8$$

$$E[Y] = 1 \times 0.2 + 2 \times 0.6 + 3 \times 0.2 = 2$$

$$E[XY] = 1 \times 0.05 + 1 \times 2 \times 0.1 + \dots + 3 \times 3 \times 0.2 = 4.85$$

$$\text{cov}(X, Y) = E[XY] - E[X]E[Y] = 1.05$$

$$\text{var}[X] = E[X^2] - E[X]^2 = 2.19.$$



# Properties of Covariance

$$\text{cov}(X, Y) = E[(X - E[X])(Y - E[Y])] = E[XY] - E[X]E[Y].$$

## Fact

(a)  $\text{var}[X] = \text{cov}(X, X)$

(b)  $X, Y$  independent  $\Rightarrow \text{cov}(X, Y) = 0$

(c)  $\text{cov}(a + X, b + Y) = \text{cov}(X, Y)$

(d)  $\text{cov}(aX + bY, cU + dV) = ac.\text{cov}(X, U) + ad.\text{cov}(X, V) + bc.\text{cov}(Y, U) + bd.\text{cov}(Y, V).$

## Proof:

Prove (a),(b),(c) yourself to check your understanding.

(d) In view of (c), one can subtract the means and assume that the RVs are zero-mean. Then,

$$\begin{aligned}\text{cov}(aX + bY, cU + dV) &= E[(aX + bY)(cU + dV)] \\ &= ac.E[XU] + ad.E[XV] + bc.E[YU] + bd.E[YV] \\ &= ac.\text{cov}(X, U) + ad.\text{cov}(X, V) + bc.\text{cov}(Y, U) + bd.\text{cov}(Y, V).\end{aligned}$$



# Linear Regression: Non-Bayesian

## Definition

Given the samples  $\{(X_n, Y_n), n = 1, \dots, N\}$ , the **Linear Regression** of  $Y$  over  $X$  is

$$\hat{Y} = a + bX$$

where  $(a, b)$  minimize

$$\sum_{n=1}^N (Y_n - a - bX_n)^2.$$

Thus,  $\hat{Y}_n = a + bX_n$  is our guess about  $Y_n$  given  $X_n$ . The squared error is  $(Y_n - \hat{Y}_n)^2$ . The LR minimizes the sum of the squared errors.

Note: This is a **non-Bayesian** formulation: there is no prior.

# Linear Least Squares Estimate

## Definition

Given two RVs  $X$  and  $Y$  with known distribution  $Pr[X = x, Y = y]$ , the **Linear Least Squares Estimate** of  $Y$  given  $X$  is

$$\hat{Y} = a + bX =: L[Y|X]$$

where  $(a, b)$  minimize

$$g(a, b) := E[(Y - a - bX)^2].$$

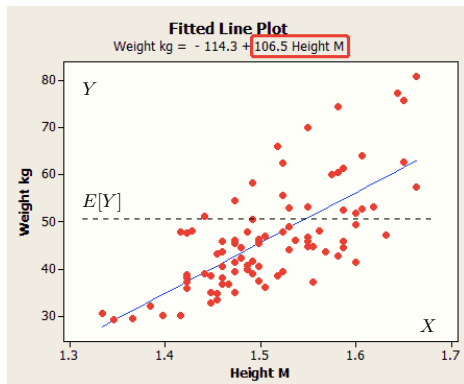
Thus,  $\hat{Y} = a + bX$  is our guess about  $Y$  given  $X$ . The squared error is  $(Y - \hat{Y})^2$ . The LLSE minimizes the expected value of the squared error.

Note: This is a **Bayesian** formulation: there is a prior.

# Linear Regression: Example

Example 1: 100 people.

Let  $(X_n, Y_n) = (\text{height, weight})$  of person  $n$ , for  $n = 1, \dots, 100$ :



The blue line is  $Y = -114.3 + 106.5X$ . ( $X$  in meters,  $Y$  in kg.) Best linear fit: [Linear Regression](#).

## LR: Non-Bayesian or Uniform?

Observe that

$$\frac{1}{N} \sum_{n=1}^N (Y_n - a - bX_n)^2 = E[(Y - a - bX)^2]$$

where one assumes that

$$(X, Y) = (X_n, Y_n), \text{ w.p. } \frac{1}{N} \text{ for } n = 1, \dots, N.$$

That is, the non-Bayesian LR is equivalent to the Bayesian LLSE that assumes that  $(X, Y)$  is uniform on the set of observed samples.

Thus, we can study the two cases LR and LLSE in one shot.

However, the interpretations are different!

# LLSE

## Theorem

Consider two RVs  $X, Y$  with a given distribution

$Pr[X = x, Y = y]$ . Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

## Proof 1:

$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X])$ . Hence,  $E[Y - \hat{Y}] = 0$ .

Also,  $E[(Y - \hat{Y})X] = 0$ , after a bit of algebra. (See next slide.)

Hence, by combining the two brown equalities,

$E[(Y - \hat{Y})(c + dX)] = 0$ . Then,  $E[(Y - \hat{Y})(\hat{Y} - a - bX)] = 0, \forall a, b$ .

Indeed:  $\hat{Y} = \alpha + \beta X$  for some  $\alpha, \beta$ , so that  $\hat{Y} - a - bX = c + dX$  for some  $c, d$ . Now,

$$\begin{aligned} E[(Y - a - bX)^2] &= E[(Y - \hat{Y} + \hat{Y} - a - bX)^2] \\ &= E[(Y - \hat{Y})^2] + E[(\hat{Y} - a - bX)^2] + 0 \geq E[(Y - \hat{Y})^2]. \end{aligned}$$

This shows that  $E[(Y - \hat{Y})^2] \leq E[(Y - a - bX)^2]$ , for all  $(a, b)$ .

Thus  $\hat{Y}$  is the LLSE. □

## A Bit of Algebra

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]).$$

Hence,  $E[Y - \hat{Y}] = 0$ . We want to show that  $E[(Y - \hat{Y})X] = 0$ .

Note that

$$E[(Y - \hat{Y})X] = E[(Y - \hat{Y})(X - E[X])],$$

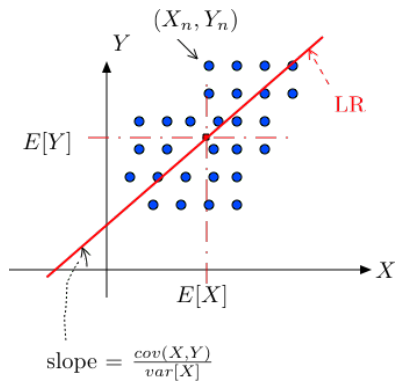
because  $E[(Y - \hat{Y})E[X]] = 0$ .

Now,

$$\begin{aligned} & E[(Y - \hat{Y})(X - E[X])] \\ &= E[(Y - E[Y])(X - E[X])] - \frac{\text{cov}(X, Y)}{\text{var}[X]} E[(X - E[X])(X - E[X])] \\ &=^{(*)} \text{cov}(X, Y) - \frac{\text{cov}(X, Y)}{\text{var}[X]} \text{var}[X] = 0. \quad \square \end{aligned}$$

(\*) Recall that  $\text{cov}(X, Y) = E[(X - E[X])(Y - E[Y])]$  and  $\text{var}[X] = E[(X - E[X])^2]$ .

## LR: Illustration



Note that

- ▶ the LR line goes through  $(E[X], E[Y])$
- ▶ its slope is  $\frac{\text{cov}(X, Y)}{\text{var}(X)}$ .



# Summary

## Linear Regression

1. Covariance:  $\text{cov}(X, Y) := E[(X - E[X])(Y - E[Y])]$ .
2. Linear Regression:  $L[Y|X] = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X])$
3. Non-Bayesian: minimize  $\sum_n (Y_n - a - bX_n)^2$
4. Bayesian: minimize  $E[(Y - a - bX)^2]$