Calculus Review

$\frac{d(e^{cx})}{dx} = ce^{cx}$.

Grows proportional to what you have! $e = (1+1/n)^n$.

$\frac{d(x^2)}{dx} = 2x$.

$\frac{(x+\delta)^2 - x^2}{\delta} = \frac{2x\delta + \delta^2}{\delta} = 2x + \delta$.

$\int x dx = \frac{x^2}{2} + c$.

Fundamental Theorem. or Triangle: width x, height x has area $\frac{x^2}{2}$.

$\frac{d(\ln x)}{dx} = \frac{1}{x}$

$x = e^y \implies 1 = e^y \frac{dy}{dx} \implies \frac{dy}{dx} = \frac{1}{e^y} = \frac{1}{x}$.

Flipping x and y axis, flips slope and function and argument.

Chain Rule: $\frac{d(f(g(x))}{dx} = f'(g(x))g'(x)dx$

Slope of $g(\cdot)$ times slope of $f(\cdot)$ at appropriate values.

Product Rule:

$(f(x)g(x))' = f'(x)g(x) + f(x)g'(x)$.

$d(uv) = udv + vdu$

Cuz:$d(uv) = uv - (u+du)(v+dv) = udv + vdu + dudv$.

Integration by Parts: $\int udv = uv - \int vdu$.

Summary

Continuous Probability 1

1. pdf: $Pr[X \in (x, x+\delta]] \approx f_X(x)\delta$.
2. CDF: $Pr[X \leq x] = F_X(x) = lim_{\delta \to 0} \sum_i f_X(x_i)\delta = \int_{-\infty}^{x} f_X(y)dy$.
3. $X \sim U[a,b]$: $f_X(x) = \frac{1}{b-a} 1\{a \leq x \leq b\}; F_X(x) = \frac{x-a}{b-a}$ for $a \leq x \leq b$.
4. $X \sim Expo(\lambda)$:
   $f_X(x) = \lambda \exp\{-\lambda x\} 1\{x \geq 0\}; F_X(x) = 1 - \exp\{-\lambda x\}$ for $x \leq 0$.
5. Target: $f_X(x) = 2x \cdot 1\{0 \leq x \leq 1\}; F_X(x) = x^2$ for $0 \leq x \leq 1$.
6. Joint pdf: $Pr[X \in (x, x+\delta), Y = (y, y+\delta)) = f_{X,Y}(x,y)\delta^2$.
   6.1 Conditional Distribution: $f_{X|Y}(x,y) = \frac{f_{X,Y}(x,y)}{f_Y(y)}$.
   6.2 Independence: $f_{X|Y}(x,y) = f_X(x)$

Poll

What is true?
X has CDF $F(x)$ and PDF $f(x)$.

(A) $Pr[X > t] = 1 - Pr[X \le t]$.
Event $X > t$ is the event that X is not $\le t$.

(B) $S(t) = Pr[X > t] = 1 - F(t)$.
Definition of CDF.

(C) $Y = 2X$, $f_Y(y) = 2f(y)$.
False. Confuses density of outcome with value oof outcome.

(D) $Y = 2X$, $F_Y(y) = F(y/2)$.
Event $Y > y$ is event $X > y/2$.

(E) $Y = 2X$, $f_Y(y) = \frac{1}{2}f(y/2)$.
Spreads out density of X over twice the range.
Chain rule from (D).

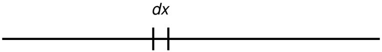
(A), (B), (D) think events, (E) think event and density.

(C) confuses probability density of outcome with value of outcome.

Discrete/Continuous

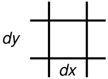
Discrete: Probability of outcome $\rightarrow$ random variables, events.

Continuous: "outcome" is real number.
Probability: Events is interval.
Density: $Pr[X \in [x, x+dx]] = f(x)dx$



$Pr[X \in [x, x+dx]] \approx f(x)dx$

Joint Continuous in d variables: "outcome" is $\in R^d$.
Probability: Events is block.
Density: $Pr[(X,Y) \in ([x, x+dx], [y, y+dx])] = f(x,y)dxdy$



$Pr[(X,Y) \in ([x, x+dx], [y, y+dy])] \approx f(x,y)dxdy$

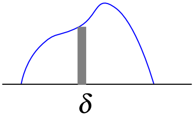
Probability

Probability!
Challenges us. But really neat.
At times, continuous.
At others, discrete.

Sample Space: Ω, $Pr[\omega]$.
Event: $Pr[A] = \sum_{\omega \in A} Pr[\omega]$
$\sum_{\omega} Pr[\omega] = 1$.
Random variables: $X(\omega)$.
Distribution: $Pr[X = x]$
$\sum_x Pr[X = x] = 1$.

Random Variable: X, Range is reals.
Event: $A = [a, b]$, $Pr[X \in A]$,
CDF: $F(x) = Pr[X \leq x]$.
PDF: $f(x) = \frac{dF(x)}{dx}$.
$\int_{-\infty}^{\infty} f(x) = 1$.

Continuous as Discrete.
$Pr[X \in [x, x+\delta]] \approx f(x)\delta$



Probability Rules are all good.

Conditional Probability.

Events: A, B

Discrete: "Heads", "Tails", $X = 1$, $Y = 5$.

Continuous: X in $[.2, .3]$. $X \in [.2, .3]$ or $X \in [.4, .6]$.

Conditional Probability: $Pr[A|B] = \frac{Pr[A \cap B]}{Pr[B]}$

Pr["Second Heads"|"First Heads"],
$Pr[X \in [.2, .3] | X \in [.2, .3]$ or $X \in [.5, .6]]$.

Total Probability Rule: $Pr[A] = Pr[A \cap B] + Pr[A \cap \overline{B}]$
Pr["Second Heads"] $= Pr[HH] + Pr[TH]$
B is First coin heads.
$Pr[X \in [.45, .55]] = Pr[X \in [.45, .50]] + Pr[X \in (.50, .55]]$
B is $X \in [0, .5]$

Product Rule: $Pr[A \cap B] = Pr[A|B]Pr[B]$.
Bayes Rule: $Pr[A|B] = Pr[B|A]Pr[A]/Pr[B]$.

All work for continuous with intervals as events.

Conditional density.

Conditional Density: $f_{X|Y}(x,y)$.

Conditional Probability: $Pr[X \in A | Y \in B] = \frac{Pr[X \in A, Y \in B]}{Pr[Y \in B]}$

$Pr[X \in [x, x+dx] | Y \in [y, y+dy]] = \frac{f_{X,Y}(x,y)dxdy}{f_Y(y)dy}$

$f_{X|Y}(x,y) = \frac{f_{X,Y}(x,y)}{f_Y(y)} = \frac{f_{X,Y}(x,y)}{\int_{-\infty}^{+\infty} f_{X,Y}(x,y)dx}$

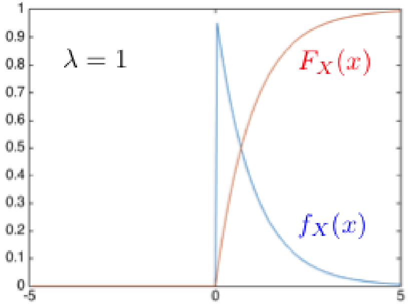
Corollary: For independent random variables, $f_{X|Y}(x,y) = f_X(x)$.

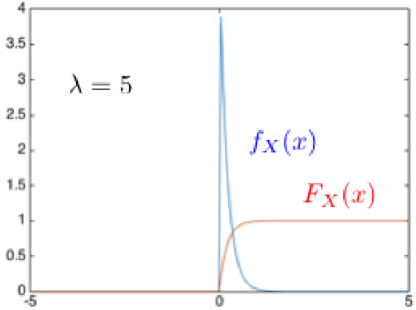
$Expo(\lambda)$

The exponential distribution with parameter $\lambda > 0$ is defined by

$$f_X(x) = \lambda e^{-\lambda x} 1\{x \geq 0\}$$

$$F_X(x) = \begin{cases} 0, & \text{if } x < 0 \\ 1 - e^{-\lambda x}, & \text{if } x \geq 0. \end{cases}$$





Note that $Pr[X > t] = e^{-\lambda t}$ for $t > 0$.

Some Properties

1. ***Expo*** **is memoryless.** Let $X = Expo(\lambda)$. Then, for $s, t > 0$,

$$\begin{aligned} Pr[X > t + s \mid X > s] &= \frac{Pr[X > t + s]}{Pr[X > s]} \\ &= \frac{e^{-\lambda(t+s)}}{e^{-\lambda s}} = e^{-\lambda t} \\ &= Pr[X > t]. \end{aligned}$$

'Used is as good as new.'

2. Scaling ***Expo*****.** Let $X = Expo(\lambda)$ and $Y = aX$ for some $a > 0$. Then

$$\begin{aligned} Pr[Y > t] &= Pr[aX > t] = Pr[X > t/a] \\ &= e^{-\lambda(t/a)} = e^{-(\lambda/a)t} = Pr[Z > t] \text{ for } Z = Expo(\lambda/a). \end{aligned}$$

Thus, $a \times Expo(\lambda) = Expo(\lambda/a)$.

Also, $Expo(\lambda) = \frac{1}{\lambda} Expo(1)$.

More Properties

3. Scaling Uniform. Let $X = U[0,1]$ and $Y = a + bX$ where $b > 0$. Then,

$$\begin{aligned} Pr[Y \in (y, y+\delta)] &= Pr[a + bX \in (y, y+\delta)] = Pr[X \in (\frac{y-a}{b}, \frac{y+\delta-a}{b})] \\ &= Pr[X \in (\frac{y-a}{b}, \frac{y-a}{b} + \frac{\delta}{b})] = \frac{1}{b}\delta, \text{ for } 0 < \frac{y-a}{b} < 1 \\ &= \frac{1}{b}\delta, \text{ for } a < y < a + b. \end{aligned}$$

Thus, $f_Y(y) = \frac{1}{b}$ for $a < y < a + b$. Hence, $Y = U[a, a+b]$.

Replace b by $b - a$, use $X = U[0,1]$, then $Y = a + (b-a)X$ is $U[a,b]$.

Some More Properties

4. Scaling pdf. Let $f_X(x)$ be the pdf of X and $Y = a + bX$ where $b > 0$. Then

$$
\begin{aligned}
Pr[Y \in (y, y+\delta)] &= Pr[a + bX \in (y, y+\delta)] = Pr[X \in (\frac{y-a}{b}, \frac{y+\delta-a}{b})] \\
&= Pr[X \in (\frac{y-a}{b}, \frac{y-a}{b} + \frac{\delta}{b})] = f_X(\frac{y-a}{b})\frac{\delta}{b}.
\end{aligned}
$$

Now, the left-hand side is $f_Y(y)\delta$. Hence,

$$f_Y(y) = \frac{1}{b} f_X(\frac{y-a}{b}).$$

Expectation

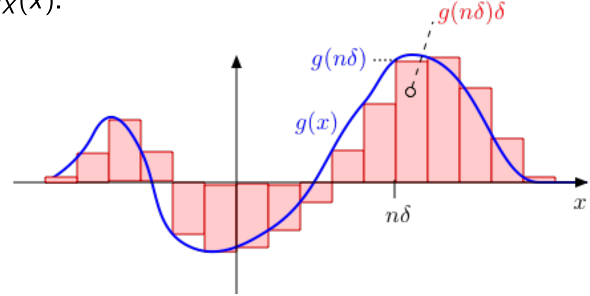
Definition: The **expectation** of a random variable X with pdf $f(x)$ is *defined* as

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx.$$

Justification: Say $X = n\delta$ w.p. $f_X(n\delta)\delta$ for $n \in \mathbb{Z}$. Then,

$$E[X] = \sum_n (n\delta) Pr[X = n\delta] = \sum_n (n\delta) f_X(n\delta)\delta = \int_{-\infty}^{\infty} x f_X(x) dx.$$

Indeed, for any g, one has $\int g(x)dx \approx \sum_n g(n\delta)\delta$. Choose $g(x) = x f_X(x)$.



Examples of Expectation

1. $X = U[0,1]$. Then, $f_X(x) = 1\{0 \le x \le 1\}$. Thus,

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_0^1 x.1 dx = [\frac{x^2}{2}]_0^1 = \frac{1}{2}.$$

2. $X =$ distance to 0 of dart shot uniformly in unit circle. Then $f_X(x) = 2x1\{0 \le x \le 1\}$. Thus,

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_0^1 x \cdot 2x dx = [\frac{2x^3}{3}]_0^1 = \frac{2}{3}.$$

Examples of Expectation

3. $X = Expo(\lambda)$. Then, $f_X(x) = \lambda e^{-\lambda x} 1\{x \geq 0\}$. Thus,

$$E[X] = \int_0^\infty x\lambda e^{-\lambda x} dx = -\int_0^\infty x de^{-\lambda x}.$$

Recall the integration by parts formula:

$$\begin{aligned}\int_a^b u(x)dv(x) &= [u(x)v(x)]_a^b - \int_a^b v(x)du(x) \\ &= u(b)v(b) - u(a)v(a) - \int_a^b v(x)du(x).\end{aligned}$$

Thus,

$$\begin{aligned}\int_0^\infty x de^{-\lambda x} &= [xe^{-\lambda x}]_0^\infty - \int_0^\infty e^{-\lambda x} dx \\ &= 0 - 0 + \frac{1}{\lambda}\int_0^\infty de^{-\lambda x} = -\frac{1}{\lambda}.\end{aligned}$$

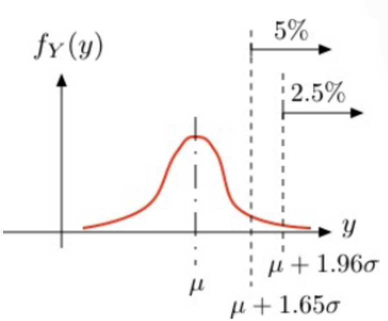
Hence, $E[X] = \frac{1}{\lambda}$.

Normal (Gaussian) Distribution.

For any μ and σ, a **normal** (aka **Gaussian**) random variable Y, which we write as $Y = \mathcal{N}(\mu, \sigma^2)$, has pdf

$$f_Y(y) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(y-\mu)^2/2\sigma^2}.$$

Standard normal has $\mu = 0$ and $\sigma = 1$.



Note: $Pr[|Y - \mu| > 1.65\sigma] \approx 10\%$; $Pr[|Y - \mu| > 2\sigma] \approx 5\%$.

Scaling and Shifting and properties

Theorem Let $X = \mathcal{N}(0,1)$ and $Y = \mu + \sigma X$. Then

$$Y = \mathcal{N}(\mu, \sigma^2).$$

Theorem If $Y = \mathcal{N}(\mu, \sigma^2)$, then

$$E[Y] = \mu \text{ and } var[Y] = \sigma^2.$$

Review: Law of Large Numbers.

Theorem: Independent identically distributed random variables, X_i,

$A_n = \frac{1}{n}\sum X_i$ "tends to the mean."

Each X_i, has $\mu = E(X_i)$ and $\sigma^2 = \text{var}(X_i)$.

Mean of A_n is μ, and variance is σ^2/n.

Used Chebyshev.

$$Pr[|A_n - \mu| > \varepsilon] \leq \frac{var[A_n]}{\varepsilon^2} = \frac{\sigma^2}{n\varepsilon} \to 0.$$

Central Limit Theorem

Central Limit Theorem

Let $X_1, X_2, \ldots$ be i.i.d. with $E[X_1] = \mu$ and $var(X_1) = \sigma^2$. Define

$$S_n := \frac{A_n - \mu}{\sigma/\sqrt{n}} = \frac{X_1 + \cdots + X_n - n\mu}{\sigma\sqrt{n}}.$$

Then,

$$S_n \to \mathcal{N}(0,1), \text{as } n \to \infty.$$

That is,

$$Pr[S_n \leq \alpha] \to \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\alpha} e^{-x^2/2} dx.$$

Proof: See EE126.

Note:

$$E(S_n) = \frac{1}{\sigma/\sqrt{n}}(E(A_n) - \mu) = 0$$

$$Var(S_n) = \frac{1}{\sigma^2/n} Var(A_n) = 1.$$

Confidence Intervals.

Recall: $A_n = \frac{1}{n}\sum X_i$, for X_i identical and independent.

For $\mu = E(X_i)$ and variance σ^2. Mean of A_n is μ, and variance is σ^2/n.

Recall Chebyshev: $Pr[|A_n - \mu| > \varepsilon] \leq \frac{var[A_n]}{\varepsilon^2}$

Implies to get confidence $1-\delta$ we need

$$\frac{var A_n}{\varepsilon^2} = \frac{1}{n}\frac{\sigma^2}{\varepsilon^2} \leq \delta \text{ or } n \geq \frac{\sigma^2}{\varepsilon^2}\frac{1}{\delta}$$

Central Limit Theorem:

$$Pr[|A_n - \mu| > \varepsilon] \leq C\int_{x\geq\varepsilon}^{\infty} e^{-\frac{x^2}{2 var A_n}} \leq Ce^{-\frac{\varepsilon^2}{2 var A_n}}$$

for $\varepsilon > \sqrt{Var A_n}$ (C is roughly $2/\sqrt{2\pi}$)

Implies to get confidence $1 - C\delta$ we need

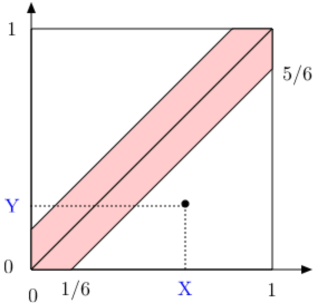
$$e^{-\frac{\varepsilon^2}{2 var A}} \leq \delta \implies \quad -\frac{n\varepsilon^2}{2\sigma^2} \leq \log\delta \implies n \geq \frac{2\sigma^2}{\varepsilon^2}\log\frac{1}{\delta}.$$

Examples: Meeting at a Restaurant

Two friends go to a restaurant independently uniformly at random between noon and 1pm.

They agree they will wait for 10 minutes.

What is the probability they meet?



Here, (X, Y) are the times when the friends reach the restaurant.

The shaded area are the pairs where $|X - Y| < 1/6$, i.e., such that they meet.

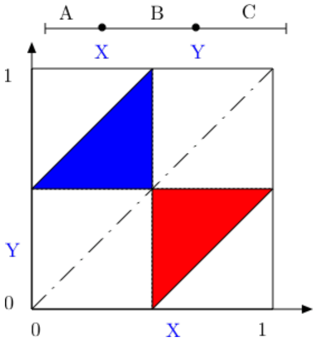
The complement is the sum of two rectangles. When you put them together, they form a square with sides $5/6$.

Thus, $Pr[\text{meet}] = 1 - (\frac{5}{6})^2 = \frac{11}{36}$.

Breaking a Stick

You break a stick at two points chosen independently uniformly at random.

What is the probability you can make a triangle with the three pieces?



Let X, Y be the two break points along the $[0,1]$ stick.

A triangle if
$A < B + C, B < A + C$, and $C < A + B$.

If $X < Y$, this means
$X < 0.5$, $Y < X + .5$, $Y > 0.5$.
This is the blue triangle.

If $X > Y$, get red triangle, by symmetry.

Thus, $Pr[\text{make triangle}] = 1/4$.

Maximum of Two Exponentials

Let $X = Expo(\lambda)$ and $Y = Expo(\mu)$ be independent.
Define $Z = \max\{X, Y\}$.

Calculate $E[Z]$.

We compute f_Z, then integrate.

One has

$$\begin{aligned} Pr[Z < z] &= Pr[X < z, Y < z] = Pr[X < z]Pr[Y < z] \\ &= (1 - e^{-\lambda z})(1 - e^{-\mu z}) = 1 - e^{-\lambda z} - e^{-\mu z} + e^{-(\lambda+\mu)z} \end{aligned}$$

Thus,

$$f_Z(z) = \lambda e^{-\lambda z} + \mu e^{-\mu z} - (\lambda + \mu)e^{-(\lambda+\mu)z}, \forall z > 0.$$

Since, $\int_0^\infty x\lambda e^{-\lambda x}dx = \lambda[-\frac{xe^{-\lambda x}}{\lambda} - \frac{e^{-\lambda x}}{\lambda^2}]_0^\infty = \frac{1}{\lambda}$.

$$E[Z] = \int_0^\infty z f_Z(z)dz = \frac{1}{\lambda} + \frac{1}{\mu} - \frac{1}{\lambda + \mu}.$$

Minimum of n i.i.d. Exponentials.

Let $X_1, \ldots, X_n$ be i.i.d. *Expo*(1). Define $Z = \min\{X_1, X_2, \ldots, X_n\}$.

What is true?

(A) Z is exponential.
(B) Parameter is n.
(C) $lim_{N \to \infty}(1 - n/N)^N \to e^{-n}$
(D) $E[Z] = 1/n$.

(C) is an argument for (A), (B) and (D).

Maximum of n i.i.d. Exponentials

Let $X_1, \ldots, X_n$ be i.i.d. *Expo*(1). Define $Z = \max\{X_1, X_2, \ldots, X_n\}$.

Calculate $E[Z]$.

We use a recursion. The key idea is as follows:

$$Z = \min\{X_1, \ldots, X_n\} + \max Y_1, \ldots, Y_{n-1}. \qquad Y_i \sim Expo(1).$$

From memoryless property of the exponential.

Let then $A_n = E[Z]$. We see that

$$\begin{aligned} A_n &= E[\min\{X_1, \ldots, X_n\}] + A_{n-1} \\ &= \frac{1}{n} + A_{n-1} \end{aligned}$$

because the minimum of *Expo* is *Expo* with the sum of the rates.

Hence,

$$E[Z] = A_n = 1 + \frac{1}{2} + \cdots + \frac{1}{n} = H(n).$$

Quantization Noise

In digital video and audio, one represents a continuous value by a finite number of bits.

This introduces an error perceived as noise: the quantization noise. What is the power of that noise?

Model: $X = U[0,1]$ is the continuous value. Y - closest multiple of 2^{-n} to X. Represent Y with n bits. The error is $Z := X - Y$.

The power of the noise is $E[Z^2]$.

Analysis: We see that Z is uniform in $[-a, a = 2^{-(n)}]$.

Thus,

$$E[Z^2] = \frac{a^2}{3} = \frac{1}{3} 2^{-2(n+1)}.$$

The power of the signal X is $E[X^2] = \frac{1}{3}$.

Quantization Noise

We saw that $E[Z^2] = \frac{1}{3}2^{-2(n+1)}$ and $E[X^2] = \frac{1}{3}$.

The **signal to noise ratio** (SNR) is the power of the signal divided by the power of the noise.

Thus,

$$SNR = 2^{2(n+1)}.$$

Expressed in decibels, one has

$$SNR(dB) = 10\log_{10}(SNR) = 20(n+1)\log_{10}(2) \approx 6(n+1).$$

For instance, if $n = 16$, then $SNR(dB) \approx 112dB$.

Expected Squared Distance

Problem 1: Pick two points X and Y independently and uniformly at random in $[0,1]$.

What is $E[(X-Y)^2]$?

Analysis: One has

$$\begin{aligned} E[(X-Y)^2] &= E[X^2+Y^2-2XY] \\ &= \frac{1}{3}+\frac{1}{3}-2\frac{1}{2}\frac{1}{2} \\ &= \frac{2}{3}-\frac{1}{2}=\frac{1}{6}. \end{aligned}$$

Problem 2: What about in a unit square?

Analysis: One has

$$\begin{aligned} E[||\mathbf{X}-\mathbf{Y}||^2] &= E[(X_1-Y_1)^2]+E[(X_2-Y_2)^2] \\ &= 2\times\frac{1}{6}. \end{aligned}$$

Problem 3: What about in n dimensions? $\frac{n}{6}$.

Summary

Continuous Probability

- Continuous RVs are similar to discrete RVs
- Think that $X \in [x, x+\varepsilon]$ with probability $f_X(x)\varepsilon$
- Sums become integrals, ...
- The exponential distribution is magical: memoryless.