

Breaking a Stick

Breaking a Stick

You break a stick at two points chosen independently uniformly at random.

Breaking a Stick

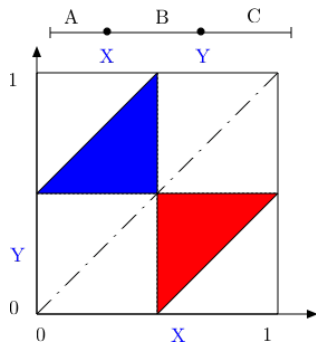
You break a stick at two points chosen independently uniformly at random.

What is the probability you can make a triangle with the three pieces?

Breaking a Stick

You break a stick at two points chosen independently uniformly at random.

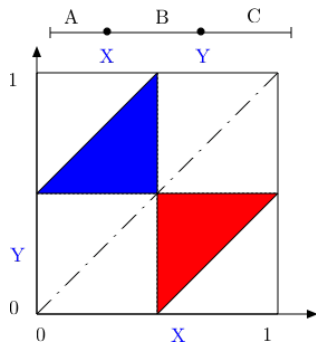
What is the probability you can make a triangle with the three pieces?



Breaking a Stick

You break a stick at two points chosen independently uniformly at random.

What is the probability you can make a triangle with the three pieces?

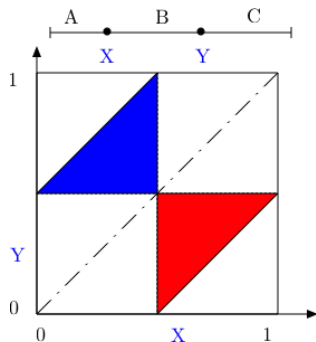


Let X, Y be the two break points along the $[0, 1]$ stick.

Breaking a Stick

You break a stick at two points chosen independently uniformly at random.

What is the probability you can make a triangle with the three pieces?



Let X, Y be the two break points along the $[0, 1]$ stick.

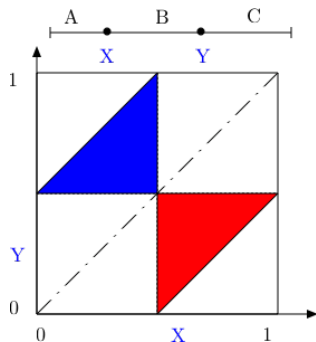
A triangle if

$A < B + C, B < A + C,$ and $C < A + B.$

Breaking a Stick

You break a stick at two points chosen independently uniformly at random.

What is the probability you can make a triangle with the three pieces?



Let X, Y be the two break points along the $[0, 1]$ stick.

A triangle if

$A < B + C, B < A + C,$ and $C < A + B.$

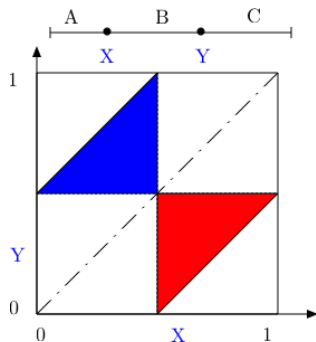
If $X < Y,$ this means

$X < 0.5,$

Breaking a Stick

You break a stick at two points chosen independently uniformly at random.

What is the probability you can make a triangle with the three pieces?



Let X, Y be the two break points along the $[0, 1]$ stick.

A triangle if

$A < B + C, B < A + C,$ and $C < A + B.$

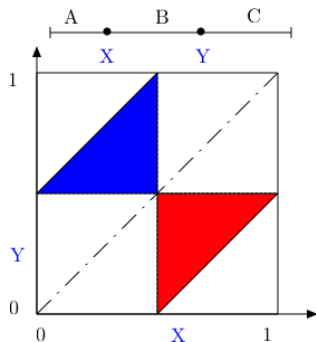
If $X < Y,$ this means

$X < 0.5, Y < X + .5,$

Breaking a Stick

You break a stick at two points chosen independently uniformly at random.

What is the probability you can make a triangle with the three pieces?



Let X, Y be the two break points along the $[0, 1]$ stick.

A triangle if

$A < B + C, B < A + C,$ and $C < A + B.$

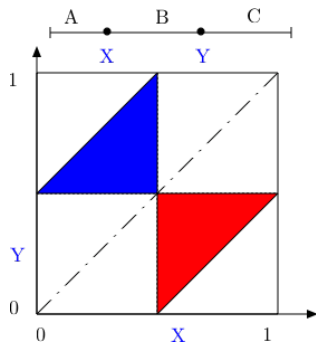
If $X < Y$, this means

$X < 0.5, Y < X + .5, Y > 0.5.$

Breaking a Stick

You break a stick at two points chosen independently uniformly at random.

What is the probability you can make a triangle with the three pieces?



Let X, Y be the two break points along the $[0, 1]$ stick.

A triangle if

$A < B + C, B < A + C,$ and $C < A + B.$

If $X < Y$, this means

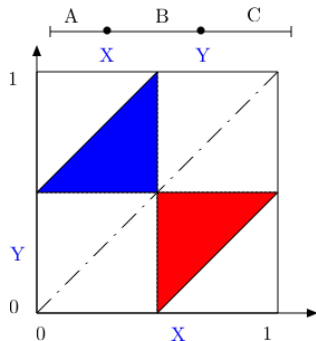
$X < 0.5, Y < X + .5, Y > 0.5.$

This is the blue triangle.

Breaking a Stick

You break a stick at two points chosen independently uniformly at random.

What is the probability you can make a triangle with the three pieces?



Let X, Y be the two break points along the $[0, 1]$ stick.

A triangle if

$A < B + C, B < A + C,$ and $C < A + B.$

If $X < Y$, this means

$X < 0.5, Y < X + .5, Y > 0.5.$

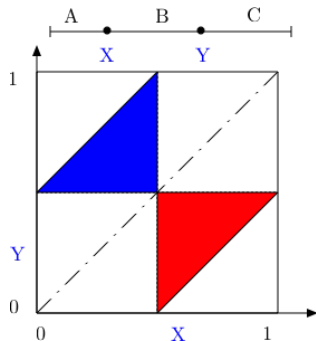
This is the blue triangle.

If $X > Y$, get red triangle, by symmetry.

Breaking a Stick

You break a stick at two points chosen independently uniformly at random.

What is the probability you can make a triangle with the three pieces?



Let X, Y be the two break points along the $[0, 1]$ stick.

A triangle if

$A < B + C, B < A + C,$ and $C < A + B.$

If $X < Y$, this means

$X < 0.5, Y < X + .5, Y > 0.5.$

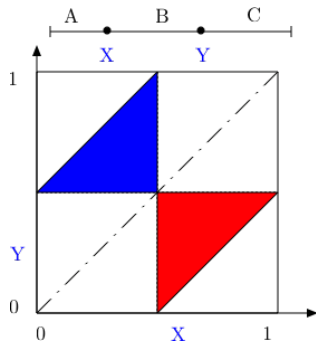
This is the blue triangle.

If $X > Y$, get red triangle, by symmetry.

Breaking a Stick

You break a stick at two points chosen independently uniformly at random.

What is the probability you can make a triangle with the three pieces?



Let X, Y be the two break points along the $[0, 1]$ stick.

A triangle if

$A < B + C, B < A + C,$ and $C < A + B.$

If $X < Y$, this means

$X < 0.5, Y < X + .5, Y > 0.5.$

This is the blue triangle.

If $X > Y$, get red triangle, by symmetry.

Thus, $Pr[\text{make triangle}] = 1/4.$

Maximum of Two Exponentials

Maximum of Two Exponentials

Let $X = \text{Expo}(\lambda)$ and $Y = \text{Expo}(\mu)$ be independent.

Maximum of Two Exponentials

Let $X = \text{Expo}(\lambda)$ and $Y = \text{Expo}(\mu)$ be independent.

Define $Z = \max\{X, Y\}$.

Maximum of Two Exponentials

Let $X = \text{Expo}(\lambda)$ and $Y = \text{Expo}(\mu)$ be independent.

Define $Z = \max\{X, Y\}$.

Calculate $E[Z]$.

Maximum of Two Exponentials

Let $X = \text{Expo}(\lambda)$ and $Y = \text{Expo}(\mu)$ be independent.

Define $Z = \max\{X, Y\}$.

Calculate $E[Z]$.

We compute f_Z , then integrate.

Maximum of Two Exponentials

Let $X = \text{Expo}(\lambda)$ and $Y = \text{Expo}(\mu)$ be independent.

Define $Z = \max\{X, Y\}$.

Calculate $E[Z]$.

We compute f_Z , then integrate.

One has

Maximum of Two Exponentials

Let $X = \text{Expo}(\lambda)$ and $Y = \text{Expo}(\mu)$ be independent.

Define $Z = \max\{X, Y\}$.

Calculate $E[Z]$.

We compute f_Z , then integrate.

One has

$$\Pr[Z < z] = \Pr[X < z, Y < z]$$

Maximum of Two Exponentials

Let $X = \text{Expo}(\lambda)$ and $Y = \text{Expo}(\mu)$ be independent.

Define $Z = \max\{X, Y\}$.

Calculate $E[Z]$.

We compute f_Z , then integrate.

One has

$$\Pr[Z < z] = \Pr[X < z, Y < z] = \Pr[X < z]\Pr[Y < z]$$

Maximum of Two Exponentials

Let $X = \text{Expo}(\lambda)$ and $Y = \text{Expo}(\mu)$ be independent.

Define $Z = \max\{X, Y\}$.

Calculate $E[Z]$.

We compute f_Z , then integrate.

One has

$$\begin{aligned} \Pr[Z < z] &= \Pr[X < z, Y < z] = \Pr[X < z]\Pr[Y < z] \\ &= (1 - e^{-\lambda z})(1 - e^{-\mu z}) = \end{aligned}$$

Maximum of Two Exponentials

Let $X = \text{Expo}(\lambda)$ and $Y = \text{Expo}(\mu)$ be independent.

Define $Z = \max\{X, Y\}$.

Calculate $E[Z]$.

We compute f_Z , then integrate.

One has

$$\begin{aligned} \Pr[Z < z] &= \Pr[X < z, Y < z] = \Pr[X < z]\Pr[Y < z] \\ &= (1 - e^{-\lambda z})(1 - e^{-\mu z}) = 1 - e^{-\lambda z} - e^{-\mu z} + e^{-(\lambda+\mu)z} \end{aligned}$$

Maximum of Two Exponentials

Let $X = \text{Expo}(\lambda)$ and $Y = \text{Expo}(\mu)$ be independent.

Define $Z = \max\{X, Y\}$.

Calculate $E[Z]$.

We compute f_Z , then integrate.

One has

$$\begin{aligned} \Pr[Z < z] &= \Pr[X < z, Y < z] = \Pr[X < z]\Pr[Y < z] \\ &= (1 - e^{-\lambda z})(1 - e^{-\mu z}) = 1 - e^{-\lambda z} - e^{-\mu z} + e^{-(\lambda+\mu)z} \end{aligned}$$

Thus,

$$f_Z(z) = \lambda e^{-\lambda z} + \mu e^{-\mu z} - (\lambda + \mu)e^{-(\lambda+\mu)z}, \forall z > 0.$$

Maximum of Two Exponentials

Let $X = \text{Expo}(\lambda)$ and $Y = \text{Expo}(\mu)$ be independent.

Define $Z = \max\{X, Y\}$.

Calculate $E[Z]$.

We compute f_Z , then integrate.

One has

$$\begin{aligned} \Pr[Z < z] &= \Pr[X < z, Y < z] = \Pr[X < z]\Pr[Y < z] \\ &= (1 - e^{-\lambda z})(1 - e^{-\mu z}) = 1 - e^{-\lambda z} - e^{-\mu z} + e^{-(\lambda+\mu)z} \end{aligned}$$

Thus,

$$f_Z(z) = \lambda e^{-\lambda z} + \mu e^{-\mu z} - (\lambda + \mu)e^{-(\lambda+\mu)z}, \forall z > 0.$$

Since, $\int_0^\infty x \lambda e^{-\lambda x} dx = \lambda \left[-\frac{x e^{-\lambda x}}{\lambda} - \frac{e^{-\lambda x}}{\lambda^2} \right]_0^\infty = \frac{1}{\lambda}$.

$$E[Z] = \int_0^\infty z f_Z(z) dz =$$

Maximum of Two Exponentials

Let $X = \text{Exp}(\lambda)$ and $Y = \text{Exp}(\mu)$ be independent.

Define $Z = \max\{X, Y\}$.

Calculate $E[Z]$.

We compute f_Z , then integrate.

One has

$$\begin{aligned}Pr[Z < z] &= Pr[X < z, Y < z] = Pr[X < z]Pr[Y < z] \\ &= (1 - e^{-\lambda z})(1 - e^{-\mu z}) = 1 - e^{-\lambda z} - e^{-\mu z} + e^{-(\lambda+\mu)z}\end{aligned}$$

Thus,

$$f_Z(z) = \lambda e^{-\lambda z} + \mu e^{-\mu z} - (\lambda + \mu)e^{-(\lambda+\mu)z}, \forall z > 0.$$

Since, $\int_0^\infty x \lambda e^{-\lambda x} dx = \lambda \left[-\frac{x e^{-\lambda x}}{\lambda} - \frac{e^{-\lambda x}}{\lambda^2} \right]_0^\infty = \frac{1}{\lambda}$.

$$E[Z] = \int_0^\infty z f_Z(z) dz = \frac{1}{\lambda} + \frac{1}{\mu} - \frac{1}{\lambda + \mu}.$$

Maximum of Two Exponentials

Let $X = \text{Exp}(\lambda)$ and $Y = \text{Exp}(\mu)$ be independent.

Define $Z = \max\{X, Y\}$.

Calculate $E[Z]$.

We compute f_Z , then integrate.

One has

$$\begin{aligned} \Pr[Z < z] &= \Pr[X < z, Y < z] = \Pr[X < z]\Pr[Y < z] \\ &= (1 - e^{-\lambda z})(1 - e^{-\mu z}) = 1 - e^{-\lambda z} - e^{-\mu z} + e^{-(\lambda+\mu)z} \end{aligned}$$

Thus,

$$f_Z(z) = \lambda e^{-\lambda z} + \mu e^{-\mu z} - (\lambda + \mu)e^{-(\lambda+\mu)z}, \forall z > 0.$$

Since, $\int_0^\infty x \lambda e^{-\lambda x} dx = \lambda \left[-\frac{x e^{-\lambda x}}{\lambda} - \frac{e^{-\lambda x}}{\lambda^2} \right]_0^\infty = \frac{1}{\lambda}$.

$$E[Z] = \int_0^\infty z f_Z(z) dz = \frac{1}{\lambda} + \frac{1}{\mu} - \frac{1}{\lambda + \mu}.$$

Minimum of n i.i.d. Exponentials.

Let X_1, \dots, X_n be i.i.d. $\text{Expo}(1)$.

Minimum of n i.i.d. Exponentials.

Let X_1, \dots, X_n be i.i.d. $\text{Expo}(1)$. Define $Z = \min\{X_1, X_2, \dots, X_n\}$.

Minimum of n i.i.d. Exponentials.

Let X_1, \dots, X_n be i.i.d. $\text{Expo}(1)$. Define $Z = \min\{X_1, X_2, \dots, X_n\}$.

What is true?

Minimum of n i.i.d. Exponentials.

Let X_1, \dots, X_n be i.i.d. $\text{Expo}(1)$. Define $Z = \min\{X_1, X_2, \dots, X_n\}$.

What is true?

- (A) Z is exponential.
- (B) Parameter is n .
- (C) $\lim_{N \rightarrow \infty} (1 - n/N)^N \rightarrow e^{-n}$
- (D) $E[Z] = 1/n$.

Minimum of n i.i.d. Exponentials.

Let X_1, \dots, X_n be i.i.d. $\text{Expo}(1)$. Define $Z = \min\{X_1, X_2, \dots, X_n\}$.

What is true?

(A) Z is exponential.

(B) Parameter is n .

(C) $\lim_{N \rightarrow \infty} (1 - n/N)^N \rightarrow e^{-n}$

(D) $E[Z] = 1/n$.

(C) is an argument for (A), (B) and (D).

Maximum of n i.i.d. Exponentials

Maximum of n i.i.d. Exponentials

Let X_1, \dots, X_n be i.i.d. $\text{Expo}(1)$.

Maximum of n i.i.d. Exponentials

Let X_1, \dots, X_n be i.i.d. $\text{Expo}(1)$. Define $Z = \max\{X_1, X_2, \dots, X_n\}$.

Maximum of n i.i.d. Exponentials

Let X_1, \dots, X_n be i.i.d. $\text{Expo}(1)$. Define $Z = \max\{X_1, X_2, \dots, X_n\}$.

Calculate $E[Z]$.

Maximum of n i.i.d. Exponentials

Let X_1, \dots, X_n be i.i.d. $\text{Expo}(1)$. Define $Z = \max\{X_1, X_2, \dots, X_n\}$.

Calculate $E[Z]$.

We use a recursion.

Maximum of n i.i.d. Exponentials

Let X_1, \dots, X_n be i.i.d. $\text{Expo}(1)$. Define $Z = \max\{X_1, X_2, \dots, X_n\}$.

Calculate $E[Z]$.

We use a recursion. The key idea is as follows:

Maximum of n i.i.d. Exponentials

Let X_1, \dots, X_n be i.i.d. $\text{Expo}(1)$. Define $Z = \max\{X_1, X_2, \dots, X_n\}$.

Calculate $E[Z]$.

We use a recursion. The key idea is as follows:

$$Z = \min\{X_1, \dots, X_n\} + \max Y_1, \dots, Y_{n-1}.$$

Maximum of n i.i.d. Exponentials

Let X_1, \dots, X_n be i.i.d. $\text{Expo}(1)$. Define $Z = \max\{X_1, X_2, \dots, X_n\}$.

Calculate $E[Z]$.

We use a recursion. The key idea is as follows:

$$Z = \min\{X_1, \dots, X_n\} + \max Y_1, \dots, Y_{n-1}. \quad Y_i \sim \text{Expo}(1).$$

Maximum of n i.i.d. Exponentials

Let X_1, \dots, X_n be i.i.d. $\text{Expo}(1)$. Define $Z = \max\{X_1, X_2, \dots, X_n\}$.

Calculate $E[Z]$.

We use a recursion. The key idea is as follows:

$$Z = \min\{X_1, \dots, X_n\} + \max Y_1, \dots, Y_{n-1}. \quad Y_i \sim \text{Expo}(1).$$

From memoryless property of the exponential.

Maximum of n i.i.d. Exponentials

Let X_1, \dots, X_n be i.i.d. $\text{Expo}(1)$. Define $Z = \max\{X_1, X_2, \dots, X_n\}$.

Calculate $E[Z]$.

We use a recursion. The key idea is as follows:

$$Z = \min\{X_1, \dots, X_n\} + \max Y_1, \dots, Y_{n-1}. \quad Y_i \sim \text{Expo}(1).$$

From memoryless property of the exponential.

Let then $A_n = E[Z]$.

Maximum of n i.i.d. Exponentials

Let X_1, \dots, X_n be i.i.d. $\text{Expo}(1)$. Define $Z = \max\{X_1, X_2, \dots, X_n\}$.

Calculate $E[Z]$.

We use a recursion. The key idea is as follows:

$$Z = \min\{X_1, \dots, X_n\} + \max Y_1, \dots, Y_{n-1}. \quad Y_i \sim \text{Expo}(1).$$

From memoryless property of the exponential.

Let then $A_n = E[Z]$. We see that

$$A_n = E[\min\{X_1, \dots, X_n\}] + A_{n-1}$$

Maximum of n i.i.d. Exponentials

Let X_1, \dots, X_n be i.i.d. $\text{Expo}(1)$. Define $Z = \max\{X_1, X_2, \dots, X_n\}$.

Calculate $E[Z]$.

We use a recursion. The key idea is as follows:

$$Z = \min\{X_1, \dots, X_n\} + \max Y_1, \dots, Y_{n-1}. \quad Y_i \sim \text{Expo}(1).$$

From memoryless property of the exponential.

Let then $A_n = E[Z]$. We see that

$$\begin{aligned} A_n &= E[\min\{X_1, \dots, X_n\}] + A_{n-1} \\ &= \frac{1}{n} + A_{n-1} \end{aligned}$$

Maximum of n i.i.d. Exponentials

Let X_1, \dots, X_n be i.i.d. $\text{Expo}(1)$. Define $Z = \max\{X_1, X_2, \dots, X_n\}$.

Calculate $E[Z]$.

We use a recursion. The key idea is as follows:

$$Z = \min\{X_1, \dots, X_n\} + \max Y_1, \dots, Y_{n-1}. \quad Y_i \sim \text{Expo}(1).$$

From memoryless property of the exponential.

Let then $A_n = E[Z]$. We see that

$$\begin{aligned} A_n &= E[\min\{X_1, \dots, X_n\}] + A_{n-1} \\ &= \frac{1}{n} + A_{n-1} \end{aligned}$$

because the minimum of Expo is Expo with the sum of the rates.

Maximum of n i.i.d. Exponentials

Let X_1, \dots, X_n be i.i.d. $\text{Expo}(1)$. Define $Z = \max\{X_1, X_2, \dots, X_n\}$. Calculate $E[Z]$.

We use a recursion. The key idea is as follows:

$$Z = \min\{X_1, \dots, X_n\} + \max Y_1, \dots, Y_{n-1}. \quad Y_i \sim \text{Expo}(1).$$

From memoryless property of the exponential.

Let then $A_n = E[Z]$. We see that

$$\begin{aligned} A_n &= E[\min\{X_1, \dots, X_n\}] + A_{n-1} \\ &= \frac{1}{n} + A_{n-1} \end{aligned}$$

because the minimum of Expo is Expo with the sum of the rates.

Hence,

$$E[Z] = A_n = 1 + \frac{1}{2} + \dots + \frac{1}{n} = H(n).$$

CS70: Markov Chains.

Markov Chains

CS70: Markov Chains.

Markov Chains

CS70: Markov Chains.

Markov Chains

1. Examples
2. Definition
3. Stationary Distribution
4. Periodicity.
5. Hitting Time.
6. Here before there.

Two-State Markov Chain

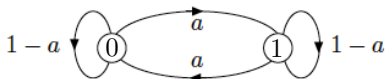
Here is a symmetric two-state Markov chain.

Two-State Markov Chain

Here is a symmetric two-state Markov chain. It describes a random motion in $\{0, 1\}$.

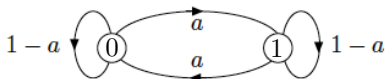
Two-State Markov Chain

Here is a symmetric two-state Markov chain. It describes a random motion in $\{0, 1\}$. Here, a is the probability that the state changes in the next step.



Two-State Markov Chain

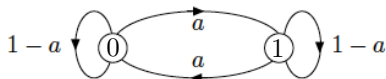
Here is a symmetric two-state Markov chain. It describes a random motion in $\{0, 1\}$. Here, a is the probability that the state changes in the next step.



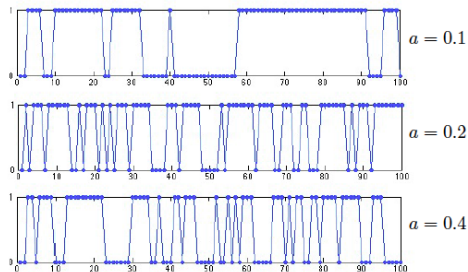
Let's simulate the Markov chain:

Two-State Markov Chain

Here is a symmetric two-state Markov chain. It describes a random motion in $\{0, 1\}$. Here, a is the probability that the state changes in the next step.

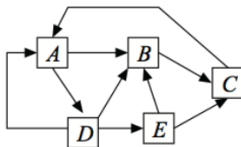


Let's simulate the Markov chain:



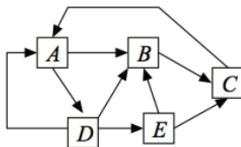
Five-State Markov Chain

At each step, the MC follows one of the outgoing arrows of the current state, with equal probabilities.



Five-State Markov Chain

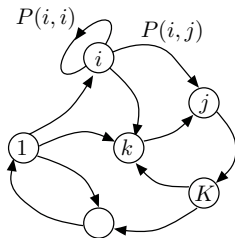
At each step, the MC follows one of the outgoing arrows of the current state, with equal probabilities.



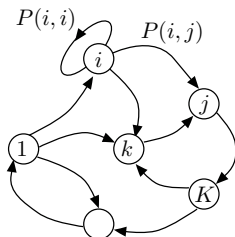
Let's simulate the Markov chain:

Finite Markov Chain: Definition

Finite Markov Chain: Definition

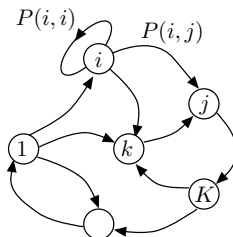


Finite Markov Chain: Definition



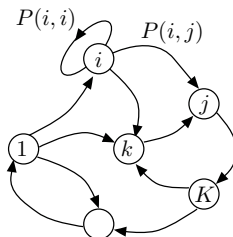
- ▶ A finite set of states: $\mathcal{X} = \{1, 2, \dots, K\}$

Finite Markov Chain: Definition



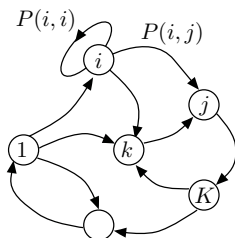
- ▶ A finite set of states: $\mathcal{X} = \{1, 2, \dots, K\}$
- ▶ A probability distribution π_0 on \mathcal{X} :

Finite Markov Chain: Definition



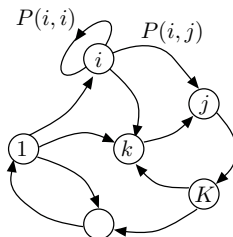
- ▶ A finite set of states: $\mathcal{X} = \{1, 2, \dots, K\}$
- ▶ A probability distribution π_0 on \mathcal{X} : $\pi_0(i) \geq 0, \sum_i \pi_0(i) = 1$

Finite Markov Chain: Definition



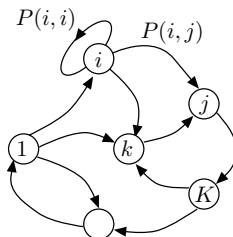
- ▶ A finite set of states: $\mathcal{X} = \{1, 2, \dots, K\}$
- ▶ A probability distribution π_0 on \mathcal{X} : $\pi_0(i) \geq 0, \sum_i \pi_0(i) = 1$
- ▶ Transition probabilities: $P(i, j)$ for $i, j \in \mathcal{X}$

Finite Markov Chain: Definition



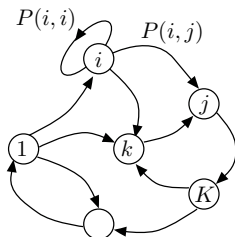
- ▶ A finite set of states: $\mathcal{X} = \{1, 2, \dots, K\}$
- ▶ A probability distribution π_0 on \mathcal{X} : $\pi_0(i) \geq 0, \sum_i \pi_0(i) = 1$
- ▶ Transition probabilities: $P(i,j)$ for $i, j \in \mathcal{X}$
 $P(i,j) \geq 0, \forall i, j;$

Finite Markov Chain: Definition



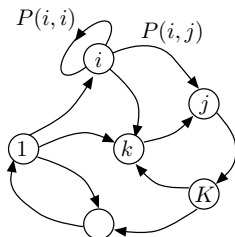
- ▶ A finite set of states: $\mathcal{X} = \{1, 2, \dots, K\}$
- ▶ A probability distribution π_0 on \mathcal{X} : $\pi_0(i) \geq 0, \sum_i \pi_0(i) = 1$
- ▶ Transition probabilities: $P(i,j)$ for $i, j \in \mathcal{X}$
 $P(i,j) \geq 0, \forall i, j; \sum_j P(i,j) = 1, \forall i$

Finite Markov Chain: Definition



- ▶ A finite set of states: $\mathcal{X} = \{1, 2, \dots, K\}$
- ▶ A probability distribution π_0 on \mathcal{X} : $\pi_0(i) \geq 0, \sum_i \pi_0(i) = 1$
- ▶ Transition probabilities: $P(i, j)$ for $i, j \in \mathcal{X}$
 $P(i, j) \geq 0, \forall i, j; \sum_j P(i, j) = 1, \forall i$
- ▶ $\{X_n, n \geq 0\}$ is defined so that

Finite Markov Chain: Definition



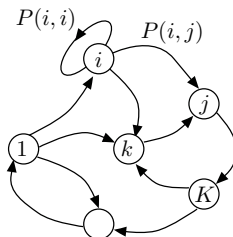
- ▶ A finite set of states: $\mathcal{X} = \{1, 2, \dots, K\}$
- ▶ A probability distribution π_0 on \mathcal{X} : $\pi_0(i) \geq 0, \sum_i \pi_0(i) = 1$
- ▶ Transition probabilities: $P(i, j)$ for $i, j \in \mathcal{X}$

$$P(i, j) \geq 0, \forall i, j; \sum_j P(i, j) = 1, \forall i$$

- ▶ $\{X_n, n \geq 0\}$ is defined so that

$$Pr[X_0 = i] = \pi_0(i), i \in \mathcal{X}$$

Finite Markov Chain: Definition



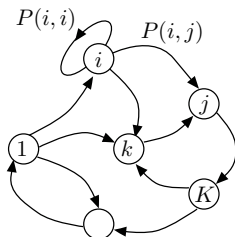
- ▶ A finite set of states: $\mathcal{X} = \{1, 2, \dots, K\}$
- ▶ A probability distribution π_0 on \mathcal{X} : $\pi_0(i) \geq 0, \sum_i \pi_0(i) = 1$
- ▶ Transition probabilities: $P(i, j)$ for $i, j \in \mathcal{X}$

$$P(i, j) \geq 0, \forall i, j; \sum_j P(i, j) = 1, \forall i$$

- ▶ $\{X_n, n \geq 0\}$ is defined so that

$$Pr[X_0 = i] = \pi_0(i), i \in \mathcal{X} \text{ (initial distribution)}$$

Finite Markov Chain: Definition



- ▶ A finite set of states: $\mathcal{X} = \{1, 2, \dots, K\}$
- ▶ A probability distribution π_0 on \mathcal{X} : $\pi_0(i) \geq 0, \sum_i \pi_0(i) = 1$
- ▶ Transition probabilities: $P(i, j)$ for $i, j \in \mathcal{X}$

$$P(i, j) \geq 0, \forall i, j; \sum_j P(i, j) = 1, \forall i$$

- ▶ $\{X_n, n \geq 0\}$ is defined so that

$$Pr[X_0 = i] = \pi_0(i), i \in \mathcal{X} \text{ (initial distribution)}$$

$$Pr[X_{n+1} = j | X_0, \dots, X_n = i] = P(i, j), i, j \in \mathcal{X}.$$

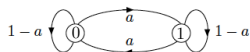
Two-State Markov Chain

Symmetric two-state Markov chain for a random motion on $\{0, 1\}$.

Two-State Markov Chain

Symmetric two-state Markov chain for a random motion on $\{0, 1\}$.

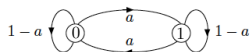
Recall a is the probability of a state change in a step.



Two-State Markov Chain

Symmetric two-state Markov chain for a random motion on $\{0, 1\}$.

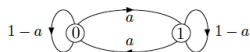
Recall a is the probability of a state change in a step.



Two-State Markov Chain

Symmetric two-state Markov chain for a random motion on $\{0, 1\}$.

Recall a is the probability of a state change in a step.

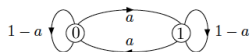


$$P = \begin{matrix} & \begin{matrix} 0 & 1 \end{matrix} \\ \begin{matrix} 0 \\ 1 \end{matrix} & \begin{pmatrix} 1-a & a \\ a & 1-a \end{pmatrix} \end{matrix}$$

Two-State Markov Chain

Symmetric two-state Markov chain for a random motion on $\{0, 1\}$.

Recall a is the probability of a state change in a step.



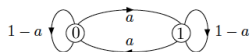
$$P = \begin{matrix} & \begin{matrix} 0 & 1 \end{matrix} \\ \begin{matrix} 0 \\ 1 \end{matrix} & \begin{pmatrix} 1-a & a \\ a & 1-a \end{pmatrix} \end{matrix}$$

Sum of row entries?

Two-State Markov Chain

Symmetric two-state Markov chain for a random motion on $\{0, 1\}$.

Recall a is the probability of a state change in a step.



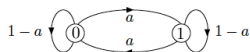
$$P = \begin{matrix} & \begin{matrix} 0 & 1 \end{matrix} \\ \begin{matrix} 0 \\ 1 \end{matrix} & \begin{pmatrix} 1-a & a \\ a & 1-a \end{pmatrix} \end{matrix}$$

Sum of row entries? 1.

Two-State Markov Chain

Symmetric two-state Markov chain for a random motion on $\{0, 1\}$.

Recall a is the probability of a state change in a step.



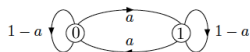
$$P = \begin{matrix} & \begin{matrix} 0 & 1 \end{matrix} \\ \begin{matrix} 0 \\ 1 \end{matrix} & \begin{pmatrix} 1-a & a \\ a & 1-a \end{pmatrix} \end{matrix}$$

Sum of row entries? 1. Always.

Two-State Markov Chain

Symmetric two-state Markov chain for a random motion on $\{0, 1\}$.

Recall a is the probability of a state change in a step.



$$P = \begin{matrix} & \begin{matrix} 0 & 1 \end{matrix} \\ \begin{matrix} 0 \\ 1 \end{matrix} & \begin{pmatrix} 1-a & a \\ a & 1-a \end{pmatrix} \end{matrix}$$

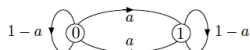
Sum of row entries? 1. Always.

Evolving distribution.

Two-State Markov Chain

Symmetric two-state Markov chain for a random motion on $\{0, 1\}$.

Recall a is the probability of a state change in a step.



$$P = \begin{matrix} & \begin{matrix} 0 & 1 \end{matrix} \\ \begin{matrix} 0 \\ 1 \end{matrix} & \begin{pmatrix} 1-a & a \\ a & 1-a \end{pmatrix} \end{matrix}$$

Sum of row entries? 1. Always.

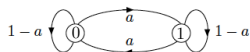
Evolving distribution.

$$\text{If } \pi_0 = [1, 0]$$

Two-State Markov Chain

Symmetric two-state Markov chain for a random motion on $\{0, 1\}$.

Recall a is the probability of a state change in a step.



$$P = \begin{matrix} & \begin{matrix} 0 & 1 \end{matrix} \\ \begin{matrix} 0 \\ 1 \end{matrix} & \begin{pmatrix} 1-a & a \\ a & 1-a \end{pmatrix} \end{matrix}$$

Sum of row entries? 1. Always.

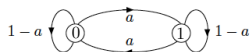
Evolving distribution.

If $\pi_0 = [1, 0]$ what is π_1 ?

Two-State Markov Chain

Symmetric two-state Markov chain for a random motion on $\{0, 1\}$.

Recall a is the probability of a state change in a step.



$$P = \begin{matrix} & \begin{matrix} 0 & 1 \end{matrix} \\ \begin{matrix} 0 \\ 1 \end{matrix} & \begin{pmatrix} 1-a & a \\ a & 1-a \end{pmatrix} \end{matrix}$$

Sum of row entries? 1. Always.

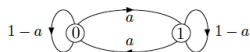
Evolving distribution.

If $\pi_0 = [1, 0]$ what is π_1 ? $\pi_1 P$

Two-State Markov Chain

Symmetric two-state Markov chain for a random motion on $\{0, 1\}$.

Recall a is the probability of a state change in a step.



$$P = \begin{matrix} & \begin{matrix} 0 & 1 \end{matrix} \\ \begin{matrix} 0 \\ 1 \end{matrix} & \begin{pmatrix} 1-a & a \\ a & 1-a \end{pmatrix} \end{matrix}$$

Sum of row entries? 1. Always.

Evolving distribution.

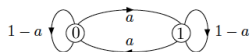
If $\pi_0 = [1, 0]$ what is π_1 ? $\pi_1 P = [1-a, a]$.

What is π_2 ?

Two-State Markov Chain

Symmetric two-state Markov chain for a random motion on $\{0, 1\}$.

Recall a is the probability of a state change in a step.



$$P = \begin{matrix} & \begin{matrix} 0 & 1 \end{matrix} \\ \begin{matrix} 0 \\ 1 \end{matrix} & \begin{pmatrix} 1-a & a \\ a & 1-a \end{pmatrix} \end{matrix}$$

Sum of row entries? 1. Always.

Evolving distribution.

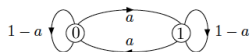
If $\pi_0 = [1, 0]$ what is π_1 ? $\pi_1 P = [1-a, a]$.

What is π_2 ? $\pi_1 P$

Two-State Markov Chain

Symmetric two-state Markov chain for a random motion on $\{0, 1\}$.

Recall a is the probability of a state change in a step.



$$P = \begin{matrix} & \begin{matrix} 0 & 1 \end{matrix} \\ \begin{matrix} 0 \\ 1 \end{matrix} & \begin{pmatrix} 1-a & a \\ a & 1-a \end{pmatrix} \end{matrix}$$

Sum of row entries? 1. Always.

Evolving distribution.

If $\pi_0 = [1, 0]$ what is π_1 ? $\pi_1 P = [1-a, a]$.

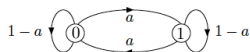
What is π_2 ? $\pi_1 P [(1-a)(1-a) + a^2, (1-a)a + a(1-a)]$

What is π_{100} ?

Two-State Markov Chain

Symmetric two-state Markov chain for a random motion on $\{0, 1\}$.

Recall a is the probability of a state change in a step.



$$P = \begin{matrix} & \begin{matrix} 0 & 1 \end{matrix} \\ \begin{matrix} 0 \\ 1 \end{matrix} & \begin{pmatrix} 1-a & a \\ a & 1-a \end{pmatrix} \end{matrix}$$

Sum of row entries? 1. Always.

Evolving distribution.

If $\pi_0 = [1, 0]$ what is π_1 ? $\pi_1 P = [1-a, a]$.

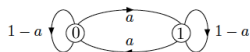
What is π_2 ? $\pi_1 P [(1-a)(1-a) + a^2, (1-a)a + a(1-a)]$

What is π_{100} ? Just guessing, but close to $[\cdot 5, \cdot 5]$.

Two-State Markov Chain

Symmetric two-state Markov chain for a random motion on $\{0, 1\}$.

Recall a is the probability of a state change in a step.



$$P = \begin{matrix} & \begin{matrix} 0 & 1 \end{matrix} \\ \begin{matrix} 0 \\ 1 \end{matrix} & \begin{pmatrix} 1-a & a \\ a & 1-a \end{pmatrix} \end{matrix}$$

Sum of row entries? 1. Always.

Evolving distribution.

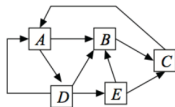
If $\pi_0 = [1, 0]$ what is π_1 ? $\pi_1 P = [1-a, a]$.

What is π_2 ? $\pi_1 P [(1-a)(1-a) + a^2, (1-a)a + a(1-a)]$

What is π_{100} ? Just guessing, but close to $[\cdot 5, \cdot 5]$. Later.

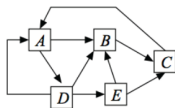
Five-State Markov Chain

MC follows each outgoing arrows of current state with equal probabilities.



Five-State Markov Chain

MC follows each outgoing arrows of current state with equal probabilities.

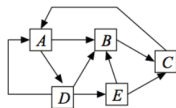


$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 0 & .5 & 0 & 0 & .5 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

Evolving distribution from $\pi_0 = [1, 0, 0, 0, 0]$?

Five-State Markov Chain

MC follows each outgoing arrows of current state with equal probabilities.



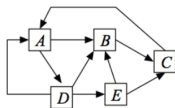
$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 0 & .5 & 0 & 0 & .5 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

Evolving distribution from $\pi_0 = [1, 0, 0, 0, 0]$?

What is π_1 ?

Five-State Markov Chain

MC follows each outgoing arrows of current state with equal probabilities.



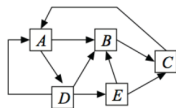
$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 0 & .5 & 0 & 0 & .5 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

Evolving distribution from $\pi_0 = [1, 0, 0, 0, 0]$?

What is π_1 ? $\pi_1 P$

Five-State Markov Chain

MC follows each outgoing arrows of current state with equal probabilities.



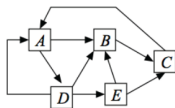
$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 0 & .5 & 0 & 0 & .5 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

Evolving distribution from $\pi_0 = [1, 0, 0, 0, 0]$?

What is π_1 ? $\pi_1 P = [0, .5, 0, .5, 0]$.

Five-State Markov Chain

MC follows each outgoing arrows of current state with equal probabilities.



$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 0 & .5 & 0 & 0 & .5 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

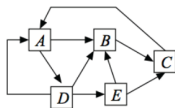
Evolving distribution from $\pi_0 = [1, 0, 0, 0, 0]$?

What is π_1 ? $\pi_1 P = [0, .5, 0, .5, 0]$.

If $\pi_t [.2, .2, .2, .2, .2]$, what is π_{t+1} ?

Five-State Markov Chain

MC follows each outgoing arrows of current state with equal probabilities.



$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 0 & .5 & 0 & 0 & .5 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

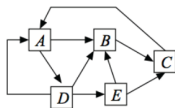
Evolving distribution from $\pi_0 = [1, 0, 0, 0, 0]$?

What is π_1 ? $\pi_1 P = [0, .5, 0, .5, 0]$.

If $\pi_t [.2, .2, .2, .2, .2]$, what is π_{t+1} ? $\pi_t P$

Five-State Markov Chain

MC follows each outgoing arrows of current state with equal probabilities.



$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 0 & .5 & 0 & 0 & .5 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

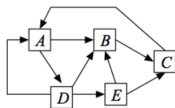
Evolving distribution from $\pi_0 = [1, 0, 0, 0, 0]$?

What is π_1 ? $\pi_1 P = [0, .5, 0, .5, 0]$.

If $\pi_t [.2, .2, .2, .2, .2]$, what is π_{t+1} ? $\pi_t P [.2, .3, .3, .1, .1]$.

Five-State Markov Chain

MC follows each outgoing arrows of current state with equal probabilities.



$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 0 & .5 & 0 & 0 & .5 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

Evolving distribution from $\pi_0 = [1, 0, 0, 0, 0]$?

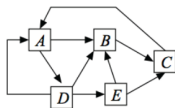
What is π_1 ? $\pi_1 P = [0, .5, 0, .5, 0]$.

If $\pi_t [.2, .2, .2, .2, .2]$, what is π_{t+1} ? $\pi_t P [.2, .3, .3, .1, .1]$.

This is just taking scaled (by .2) in-degree.

Five-State Markov Chain

MC follows each outgoing arrows of current state with equal probabilities.



$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 0 & .5 & 0 & 0 & .5 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

Evolving distribution from $\pi_0 = [1, 0, 0, 0, 0]$?

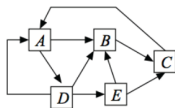
What is π_1 ? $\pi_1 P = [0, .5, 0, .5, 0]$.

If $\pi_t [.2, .2, .2, .2, .2]$, what is π_{t+1} ? $\pi_t P [.2, .3, .3, .1, .1]$.

This is just taking scaled (by .2) in-degree. Only works for uniform.

Five-State Markov Chain

MC follows each outgoing arrows of current state with equal probabilities.



$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 0 & .5 & 0 & 0 & .5 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

Evolving distribution from $\pi_0 = [1, 0, 0, 0, 0]$?

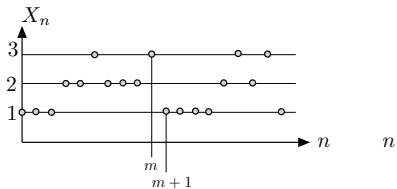
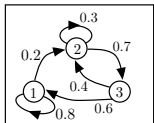
What is π_1 ? $\pi_1 P = [0, .5, 0, .5, 0]$.

If $\pi_t [.2, .2, .2, .2, .2]$, what is π_{t+1} ? $\pi_t P [.2, .3, .3, .1, .1]$.

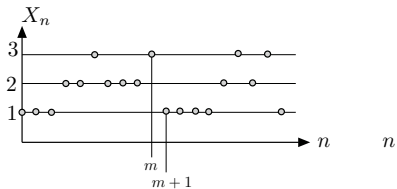
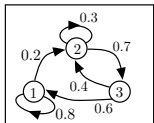
This is just taking scaled (by .2) in-degree. Only works for uniform.

What is it at π_{10000} ?

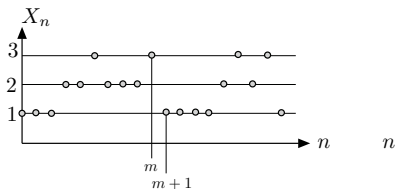
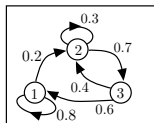
Distribution of X_n



Distribution of X_n

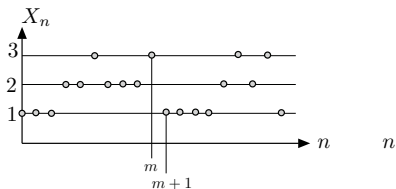
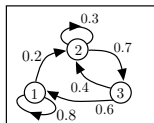


Distribution of X_n



Recall π_n is a distribution over states for X_n .

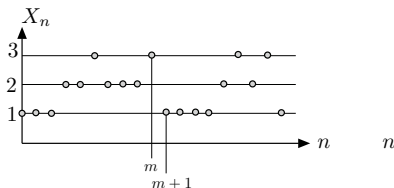
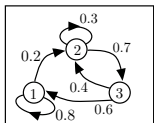
Distribution of X_n



Recall π_n is a distribution over states for X_n .

Stationary distribution: $\pi = \pi P$.

Distribution of X_n

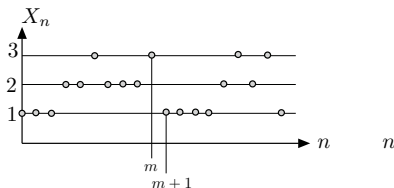
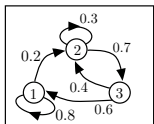


Recall π_n is a distribution over states for X_n .

Stationary distribution: $\pi = \pi P$.

Distribution over states is the same before/after transition.
probability entering i : $\sum_{j,i} P(j, i)\pi(j)$.

Distribution of X_n



Recall π_n is a distribution over states for X_n .

Stationary distribution: $\pi = \pi P$.

Distribution over states is the same before/after transition.

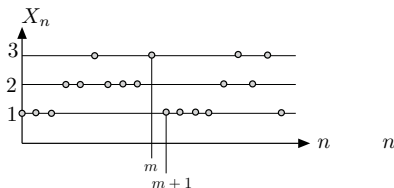
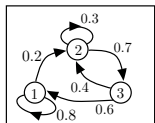
probability entering i : $\sum_{j,i} P(j, i)\pi(j)$.

probability leaving i : π_i .

are Equal!

Distribution same after one step.

Distribution of X_n



Recall π_n is a distribution over states for X_n .

Stationary distribution: $\pi = \pi P$.

Distribution over states is the same before/after transition.

probability entering i : $\sum_{j,i} P(j, i)\pi(j)$.

probability leaving i : π_i .

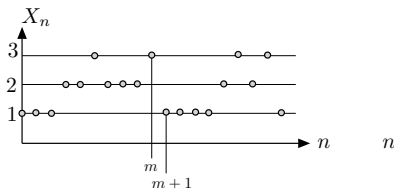
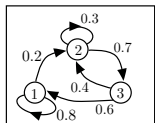
are Equal!

Distribution same after one step.

Questions? Does one exist? Is it unique?

If it exists and is unique.

Distribution of X_n



Recall π_n is a distribution over states for X_n .

Stationary distribution: $\pi = \pi P$.

Distribution over states is the same before/after transition.

probability entering i : $\sum_{j,i} P(j, i)\pi(j)$.

probability leaving i : π_i .

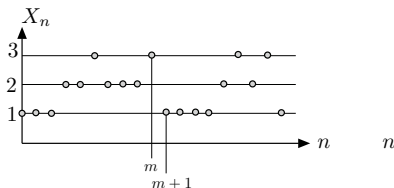
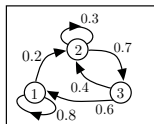
are Equal!

Distribution same after one step.

Questions? Does one exist? Is it unique?

If it exists and is unique. Then what?

Distribution of X_n



Recall π_n is a distribution over states for X_n .

Stationary distribution: $\pi = \pi P$.

Distribution over states is the same before/after transition.

probability entering i : $\sum_{j,i} P(j, i)\pi(j)$.

probability leaving i : π_i .

are Equal!

Distribution same after one step.

Questions? Does one exist? Is it unique?

If it exists and is unique. Then what?

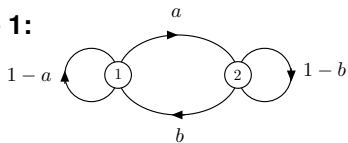
Sometimes the distribution as $n \rightarrow \infty$

Stationary: Example

Example 1:

Stationary: Example

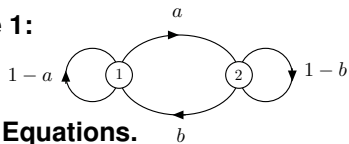
Example 1:



$$P = \begin{bmatrix} 1-a & a \\ b & 1-b \end{bmatrix}$$

Stationary: Example

Example 1:



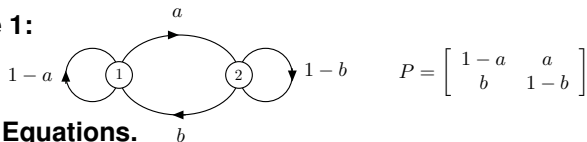
$$P = \begin{bmatrix} 1-a & a \\ b & 1-b \end{bmatrix}$$

Balance Equations.

$$\pi P = \pi$$

Stationary: Example

Example 1:



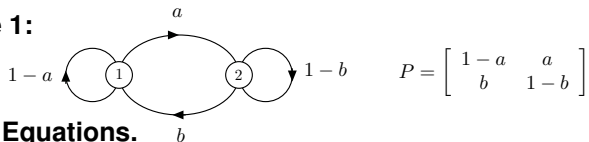
$$P = \begin{bmatrix} 1-a & a \\ b & 1-b \end{bmatrix}$$

Balance Equations.

$$\pi P = \pi \Leftrightarrow [\pi(1), \pi(2)] \begin{bmatrix} 1-a & a \\ b & 1-b \end{bmatrix} = [\pi(1), \pi(2)]$$

Stationary: Example

Example 1:



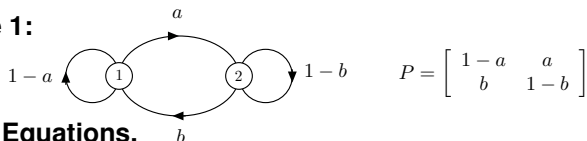
$$P = \begin{bmatrix} 1-a & a \\ b & 1-b \end{bmatrix}$$

Balance Equations.

$$\begin{aligned} \pi P = \pi &\Leftrightarrow [\pi(1), \pi(2)] \begin{bmatrix} 1-a & a \\ b & 1-b \end{bmatrix} = [\pi(1), \pi(2)] \\ &\Leftrightarrow \pi(1)(1-a) + \pi(2)b = \pi(1) \text{ and} \end{aligned}$$

Stationary: Example

Example 1:



$$P = \begin{bmatrix} 1-a & a \\ b & 1-b \end{bmatrix}$$

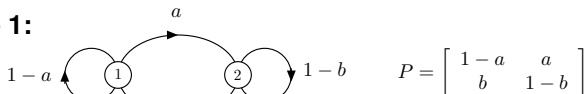
Balance Equations.

$$\pi P = \pi \Leftrightarrow [\pi(1), \pi(2)] \begin{bmatrix} 1-a & a \\ b & 1-b \end{bmatrix} = [\pi(1), \pi(2)]$$

$$\Leftrightarrow \pi(1)(1-a) + \pi(2)b = \pi(1) \text{ and } \pi(1)a + \pi(2)(1-b) = \pi(2)$$

Stationary: Example

Example 1:



$$P = \begin{bmatrix} 1-a & a \\ b & 1-b \end{bmatrix}$$

Balance Equations.

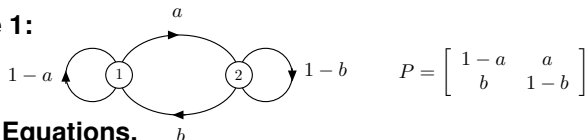
$$\pi P = \pi \Leftrightarrow [\pi(1), \pi(2)] \begin{bmatrix} 1-a & a \\ b & 1-b \end{bmatrix} = [\pi(1), \pi(2)]$$

$$\Leftrightarrow \pi(1)(1-a) + \pi(2)b = \pi(1) \text{ and } \pi(1)a + \pi(2)(1-b) = \pi(2)$$

$$\Leftrightarrow \pi(1)a = \pi(2)b.$$

Stationary: Example

Example 1:



$$P = \begin{bmatrix} 1-a & a \\ b & 1-b \end{bmatrix}$$

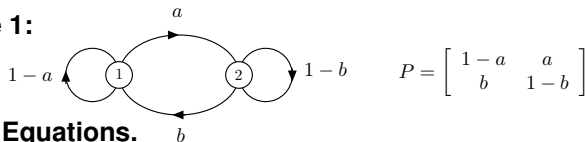
Balance Equations.

$$\begin{aligned} \pi P = \pi &\Leftrightarrow [\pi(1), \pi(2)] \begin{bmatrix} 1-a & a \\ b & 1-b \end{bmatrix} = [\pi(1), \pi(2)] \\ &\Leftrightarrow \pi(1)(1-a) + \pi(2)b = \pi(1) \text{ and } \pi(1)a + \pi(2)(1-b) = \pi(2) \\ &\Leftrightarrow \pi(1)a = \pi(2)b. \end{aligned}$$

These equations are redundant!

Stationary: Example

Example 1:



$$P = \begin{bmatrix} 1-a & a \\ b & 1-b \end{bmatrix}$$

Balance Equations.

$$\pi P = \pi \Leftrightarrow [\pi(1), \pi(2)] \begin{bmatrix} 1-a & a \\ b & 1-b \end{bmatrix} = [\pi(1), \pi(2)]$$

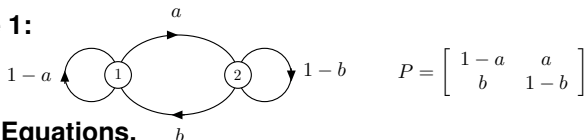
$$\Leftrightarrow \pi(1)(1-a) + \pi(2)b = \pi(1) \text{ and } \pi(1)a + \pi(2)(1-b) = \pi(2)$$

$$\Leftrightarrow \pi(1)a = \pi(2)b.$$

These equations are redundant! We have to add an equation:

Stationary: Example

Example 1:



$$P = \begin{bmatrix} 1-a & a \\ b & 1-b \end{bmatrix}$$

Balance Equations.

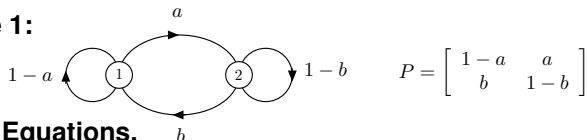
$$\begin{aligned} \pi P = \pi &\Leftrightarrow [\pi(1), \pi(2)] \begin{bmatrix} 1-a & a \\ b & 1-b \end{bmatrix} = [\pi(1), \pi(2)] \\ &\Leftrightarrow \pi(1)(1-a) + \pi(2)b = \pi(1) \text{ and } \pi(1)a + \pi(2)(1-b) = \pi(2) \\ &\Leftrightarrow \pi(1)a = \pi(2)b. \end{aligned}$$

These equations are redundant! We have to add an equation:

$$\pi(1) + \pi(2) = 1.$$

Stationary: Example

Example 1:



$$P = \begin{bmatrix} 1-a & a \\ b & 1-b \end{bmatrix}$$

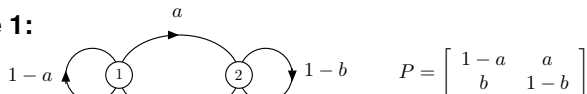
Balance Equations.

$$\begin{aligned} \pi P = \pi &\Leftrightarrow [\pi(1), \pi(2)] \begin{bmatrix} 1-a & a \\ b & 1-b \end{bmatrix} = [\pi(1), \pi(2)] \\ &\Leftrightarrow \pi(1)(1-a) + \pi(2)b = \pi(1) \text{ and } \pi(1)a + \pi(2)(1-b) = \pi(2) \\ &\Leftrightarrow \pi(1)a = \pi(2)b. \end{aligned}$$

These equations are redundant! We have to add an equation:
 $\pi(1) + \pi(2) = 1$. Then we find

Stationary: Example

Example 1:



$$P = \begin{bmatrix} 1-a & a \\ b & 1-b \end{bmatrix}$$

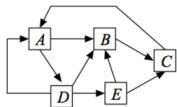
Balance Equations.

$$\begin{aligned} \pi P = \pi &\Leftrightarrow [\pi(1), \pi(2)] \begin{bmatrix} 1-a & a \\ b & 1-b \end{bmatrix} = [\pi(1), \pi(2)] \\ &\Leftrightarrow \pi(1)(1-a) + \pi(2)b = \pi(1) \text{ and } \pi(1)a + \pi(2)(1-b) = \pi(2) \\ &\Leftrightarrow \pi(1)a = \pi(2)b. \end{aligned}$$

These equations are redundant! We have to add an equation:
 $\pi(1) + \pi(2) = 1$. Then we find

$$\pi = \left[\frac{b}{a+b}, \frac{a}{a+b} \right].$$

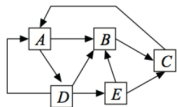
Stationary: Example 2



$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 1/3 & 1/3 & 0 & 0 & 1/3 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

Balance equations: $\pi P = \pi$.

Stationary: Example 2

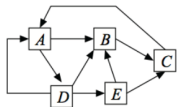


$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 1/3 & 1/3 & 0 & 0 & 1/3 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

Balance equations: $\pi P = \pi$.

$$\pi(C) + 1/3\pi(D) = \pi(A)$$

Stationary: Example 2



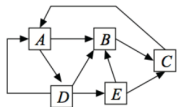
$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 1/3 & 1/3 & 0 & 0 & 1/3 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

Balance equations: $\pi P = \pi$.

$$\pi(C) + 1/3\pi(D) = \pi(A)$$

$$.5\pi(A) + 1/3\pi(D) + .5\pi(C) = \pi(B)$$

Stationary: Example 2



$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 1/3 & 1/3 & 0 & 0 & 1/3 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

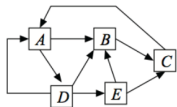
Balance equations: $\pi P = \pi$.

$$\pi(C) + 1/3\pi(D) = \pi(A)$$

$$.5\pi(A) + 1/3\pi(D) + .5\pi(C) = \pi(B)$$

$$1\pi(B) + .5\pi(E) = \pi(C)$$

Stationary: Example 2



$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 1/3 & 1/3 & 0 & 0 & 1/3 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

Balance equations: $\pi P = \pi$.

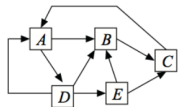
$$\pi(C) + 1/3\pi(D) = \pi(A)$$

$$.5\pi(A) + 1/3\pi(D) + .5\pi(C) = \pi(B)$$

$$1\pi(B) + .5\pi(E) = \pi(C)$$

$$.5\pi(A) = \pi(D)$$

Stationary: Example 2



$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 1/3 & 1/3 & 0 & 0 & 1/3 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

Balance equations: $\pi P = \pi$.

$$\pi(C) + 1/3\pi(D) = \pi(A)$$

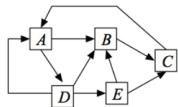
$$.5\pi(A) + 1/3\pi(D) + .5\pi(C) = \pi(B)$$

$$1\pi(B) + .5\pi(E) = \pi(C)$$

$$.5\pi(A) = \pi(D)$$

$$1/3\pi(D) = \pi(E)$$

Stationary: Example 2



$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 1/3 & 1/3 & 0 & 0 & 1/3 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

Balance equations: $\pi P = \pi$.

$$\pi(C) + 1/3\pi(D) = \pi(A)$$

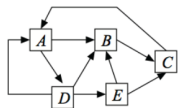
$$.5\pi(A) + 1/3\pi(D) + .5\pi(C) = \pi(B)$$

$$1\pi(B) + .5\pi(E) = \pi(C)$$

$$.5\pi(A) = \pi(D)$$

$$1/3\pi(D) = \pi(E)$$

Stationary: Example 2



$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 1/3 & 1/3 & 0 & 0 & 1/3 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

Balance equations: $\pi P = \pi$.

$$\pi(C) + 1/3\pi(D) = \pi(A)$$

$$.5\pi(A) + 1/3\pi(D) + .5\pi(C) = \pi(B)$$

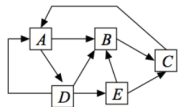
$$1\pi(B) + .5\pi(E) = \pi(C)$$

$$.5\pi(A) = \pi(D)$$

$$1/3\pi(D) = \pi(E)$$

Plus $\pi(A) + \pi(B) + \pi(C) + \pi(D) + \pi(E) = 1$.

Stationary: Example 2



$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 1/3 & 1/3 & 0 & 0 & 1/3 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

Balance equations: $\pi P = \pi$.

$$\pi(C) + 1/3\pi(D) = \pi(A)$$

$$.5\pi(A) + 1/3\pi(D) + .5\pi(C) = \pi(B)$$

$$1\pi(B) + .5\pi(E) = \pi(C)$$

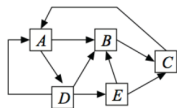
$$.5\pi(A) = \pi(D)$$

$$1/3\pi(D) = \pi(E)$$

Plus $\pi(A) + \pi(B) + \pi(C) + \pi(D) + \pi(E) = 1$.

Solution: $\frac{1}{39}[12, 9, 10, 6, 2]$.

Stationary: Example 2



$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 1/3 & 1/3 & 0 & 0 & 1/3 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

Balance equations: $\pi P = \pi$.

$$\pi(C) + 1/3\pi(D) = \pi(A)$$

$$.5\pi(A) + 1/3\pi(D) + .5\pi(C) = \pi(B)$$

$$1\pi(B) + .5\pi(E) = \pi(C)$$

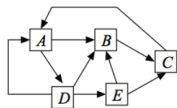
$$.5\pi(A) = \pi(D)$$

$$1/3\pi(D) = \pi(E)$$

Plus $\pi(A) + \pi(B) + \pi(C) + \pi(D) + \pi(E) = 1$.

Solution: $\frac{1}{39}[12, 9, 10, 6, 2]$. After a long time on [ChatGPT](#).

Stationary: Example 2



$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 1/3 & 1/3 & 0 & 0 & 1/3 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

Balance equations: $\pi P = \pi$.

$$\pi(C) + 1/3\pi(D) = \pi(A)$$

$$.5\pi(A) + 1/3\pi(D) + .5\pi(C) = \pi(B)$$

$$1\pi(B) + .5\pi(E) = \pi(C)$$

$$.5\pi(A) = \pi(D)$$

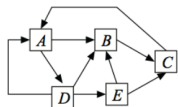
$$1/3\pi(D) = \pi(E)$$

Plus $\pi(A) + \pi(B) + \pi(C) + \pi(D) + \pi(E) = 1$.

Solution: $\frac{1}{39}[12, 9, 10, 6, 2]$. After a long time on [ChatGPT](#).

Verify: adds to 1.

Stationary: Example 2



$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 1/3 & 1/3 & 0 & 0 & 1/3 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

Balance equations: $\pi P = \pi$.

$$\pi(C) + 1/3\pi(D) = \pi(A)$$

$$.5\pi(A) + 1/3\pi(D) + .5\pi(C) = \pi(B)$$

$$1\pi(B) + .5\pi(E) = \pi(C)$$

$$.5\pi(A) = \pi(D)$$

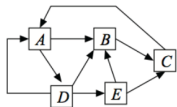
$$1/3\pi(D) = \pi(E)$$

Plus $\pi(A) + \pi(B) + \pi(C) + \pi(D) + \pi(E) = 1$.

Solution: $\frac{1}{39}[12, 9, 10, 6, 2]$. After a long time on [ChatGPT](#).

Verify: adds to 1. $\pi(A)$

Stationary: Example 2



$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 1/3 & 1/3 & 0 & 0 & 1/3 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

Balance equations: $\pi P = \pi$.

$$\pi(C) + 1/3\pi(D) = \pi(A)$$

$$.5\pi(A) + 1/3\pi(D) + .5\pi(C) = \pi(B)$$

$$1\pi(B) + .5\pi(E) = \pi(C)$$

$$.5\pi(A) = \pi(D)$$

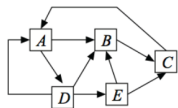
$$1/3\pi(D) = \pi(E)$$

Plus $\pi(A) + \pi(B) + \pi(C) + \pi(D) + \pi(E) = 1$.

Solution: $\frac{1}{39}[12, 9, 10, 6, 2]$. After a long time on [ChatGPT](#).

Verify: adds to 1. $\pi(A) = \pi(C) + 1/3\pi(D) \propto_{39} 10 + 1/3 \times 6$

Stationary: Example 2



$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 1/3 & 1/3 & 0 & 0 & 1/3 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

Balance equations: $\pi P = \pi$.

$$\pi(C) + 1/3\pi(D) = \pi(A)$$

$$.5\pi(A) + 1/3\pi(D) + .5\pi(C) = \pi(B)$$

$$1\pi(B) + .5\pi(E) = \pi(C)$$

$$.5\pi(A) = \pi(D)$$

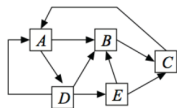
$$1/3\pi(D) = \pi(E)$$

Plus $\pi(A) + \pi(B) + \pi(C) + \pi(D) + \pi(E) = 1$.

Solution: $\frac{1}{39}[12, 9, 10, 6, 2]$. After a long time on [ChatGPT](#).

Verify: adds to 1. $\pi(A) = \pi(C) + 1/3\pi(D) \propto_{39} 10 + 1/3 \times 6 = 12$.

Stationary: Example 2



$$P = \begin{matrix} & \begin{matrix} A & B & C & D & E \end{matrix} \\ \begin{matrix} A \\ B \\ C \\ D \\ E \end{matrix} & \begin{pmatrix} 0 & .5 & 0 & .5 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 1/3 & 1/3 & 0 & 0 & 1/3 \\ 0 & .5 & .5 & 0 & 0 \end{pmatrix} \end{matrix}$$

Balance equations: $\pi P = \pi$.

$$\pi(C) + 1/3\pi(D) = \pi(A)$$

$$.5\pi(A) + 1/3\pi(D) + .5\pi(C) = \pi(B)$$

$$1\pi(B) + .5\pi(E) = \pi(C)$$

$$.5\pi(A) = \pi(D)$$

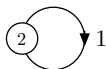
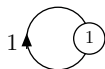
$$1/3\pi(D) = \pi(E)$$

Plus $\pi(A) + \pi(B) + \pi(C) + \pi(D) + \pi(E) = 1$.

Solution: $\frac{1}{39}[12, 9, 10, 6, 2]$. After a long time on [ChatGPT](#).

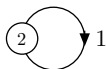
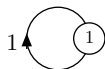
Verify: adds to 1. $\pi(A) = \pi(C) + 1/3\pi(D) \propto_{39} 10 + 1/3 \times 6 = 12$

Stationary distributions: Example 3



$$P = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

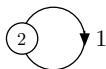
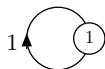
Stationary distributions: Example 3



$$P = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

$$\pi P = \pi$$

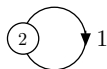
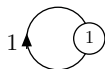
Stationary distributions: Example 3



$$P = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

$$\pi P = \pi \Leftrightarrow [\pi(1), \pi(2)] \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} = [\pi(1), \pi(2)]$$

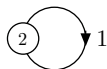
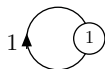
Stationary distributions: Example 3



$$P = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

$$\pi P = \pi \Leftrightarrow [\pi(1), \pi(2)] \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} = [\pi(1), \pi(2)] \Leftrightarrow \pi(1) = \pi(1) \text{ and}$$

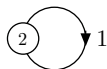
Stationary distributions: Example 3



$$P = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

$$\pi P = \pi \Leftrightarrow [\pi(1), \pi(2)] \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} = [\pi(1), \pi(2)] \Leftrightarrow \pi(1) = \pi(1) \text{ and } \pi(2) = \pi(2).$$

Stationary distributions: Example 3

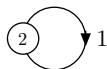


$$P = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

$$\pi P = \pi \Leftrightarrow [\pi(1), \pi(2)] \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} = [\pi(1), \pi(2)] \Leftrightarrow \pi(1) = \pi(1) \text{ and } \pi(2) = \pi(2).$$

Every distribution is invariant for this Markov chain.

Stationary distributions: Example 3

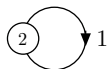
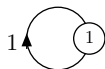


$$P = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

$$\pi P = \pi \Leftrightarrow [\pi(1), \pi(2)] \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} = [\pi(1), \pi(2)] \Leftrightarrow \pi(1) = \pi(1) \text{ and } \pi(2) = \pi(2).$$

Every distribution is invariant for this Markov chain. Since $X_n = X_0$ for all n .

Stationary distributions: Example 3

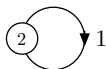
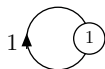


$$P = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

$$\pi P = \pi \Leftrightarrow [\pi(1), \pi(2)] \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} = [\pi(1), \pi(2)] \Leftrightarrow \pi(1) = \pi(1) \text{ and } \pi(2) = \pi(2).$$

Every distribution is invariant for this Markov chain. Since $X_n = X_0$ for all n . Hence, $Pr[X_n = i] = Pr[X_0 = i], \forall (i, n)$.

Stationary distributions: Example 3



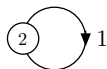
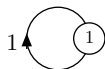
$$P = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

$$\pi P = \pi \Leftrightarrow [\pi(1), \pi(2)] \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} = [\pi(1), \pi(2)] \Leftrightarrow \pi(1) = \pi(1) \text{ and } \pi(2) = \pi(2).$$

Every distribution is invariant for this Markov chain. Since $X_n = X_0$ for all n . Hence, $Pr[X_n = i] = Pr[X_0 = i], \forall (i, n)$.

Discussion.

Stationary distributions: Example 3



$$P = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

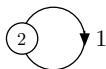
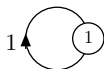
$$\pi P = \pi \Leftrightarrow [\pi(1), \pi(2)] \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} = [\pi(1), \pi(2)] \Leftrightarrow \pi(1) = \pi(1) \text{ and } \pi(2) = \pi(2).$$

Every distribution is invariant for this Markov chain. Since $X_n = X_0$ for all n . Hence, $Pr[X_n = i] = Pr[X_0 = i], \forall (i, n)$.

Discussion.

We have seen a chain with one stationary,

Stationary distributions: Example 3



$$P = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

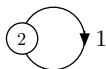
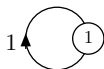
$$\pi P = \pi \Leftrightarrow [\pi(1), \pi(2)] \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} = [\pi(1), \pi(2)] \Leftrightarrow \pi(1) = \pi(1) \text{ and } \pi(2) = \pi(2).$$

Every distribution is invariant for this Markov chain. Since $X_n = X_0$ for all n . Hence, $Pr[X_n = i] = Pr[X_0 = i], \forall (i, n)$.

Discussion.

We have seen a chain with one stationary,
and a chain with many.

Stationary distributions: Example 3



$$P = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

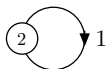
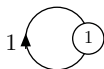
$$\pi P = \pi \Leftrightarrow [\pi(1), \pi(2)] \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} = [\pi(1), \pi(2)] \Leftrightarrow \pi(1) = \pi(1) \text{ and } \pi(2) = \pi(2).$$

Every distribution is invariant for this Markov chain. Since $X_n = X_0$ for all n . Hence, $Pr[X_n = i] = Pr[X_0 = i], \forall (i, n)$.

Discussion.

We have seen a chain with one stationary,
and a chain with many.

Stationary distributions: Example 3



$$P = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

$$\pi P = \pi \Leftrightarrow [\pi(1), \pi(2)] \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} = [\pi(1), \pi(2)] \Leftrightarrow \pi(1) = \pi(1) \text{ and } \pi(2) = \pi(2).$$

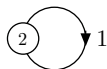
Every distribution is invariant for this Markov chain. Since $X_n = X_0$ for all n . Hence, $Pr[X_n = i] = Pr[X_0 = i], \forall (i, n)$.

Discussion.

We have seen a chain with one stationary,
and a chain with many.

When is there just one?

Stationary distributions: Example 3



$$P = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

$$\pi P = \pi \Leftrightarrow [\pi(1), \pi(2)] \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} = [\pi(1), \pi(2)] \Leftrightarrow \pi(1) = \pi(1) \text{ and } \pi(2) = \pi(2).$$

Every distribution is invariant for this Markov chain. Since $X_n = X_0$ for all n . Hence, $Pr[X_n = i] = Pr[X_0 = i], \forall (i, n)$.

Discussion.

We have seen a chain with one stationary,
and a chain with many.

When is there just one? When is a stationary distribution unique?

Irreducibility.

Definition A Markov chain is **irreducible** if it can go from every state i to every state j

Irreducibility.

Definition A Markov chain is **irreducible** if it can go from every state i to every state j (possibly in multiple steps).

Irreducibility.

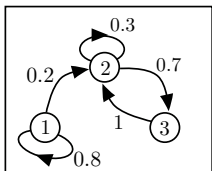
Definition A Markov chain is **irreducible** if it can go from every state i to every state j (possibly in multiple steps).

Examples:

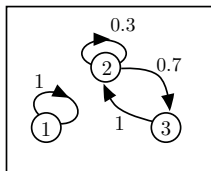
Irreducibility.

Definition A Markov chain is **irreducible** if it can go from every state i to every state j (possibly in multiple steps).

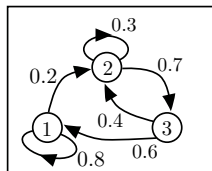
Examples:



[A]



[B]

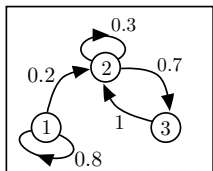


[C]

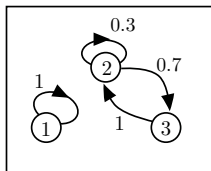
Irreducibility.

Definition A Markov chain is **irreducible** if it can go from every state i to every state j (possibly in multiple steps).

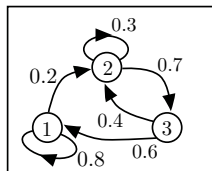
Examples:



[A]



[B]



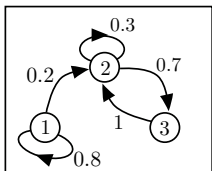
[C]

[A] is

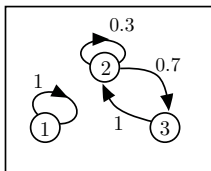
Irreducibility.

Definition A Markov chain is **irreducible** if it can go from every state i to every state j (possibly in multiple steps).

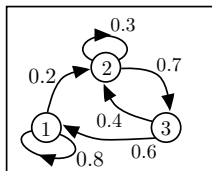
Examples:



[A]



[B]



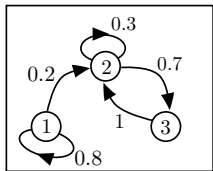
[C]

[A] is **not irreducible**.

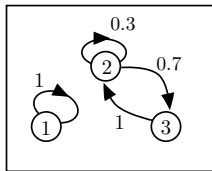
Irreducibility.

Definition A Markov chain is **irreducible** if it can go from every state i to every state j (possibly in multiple steps).

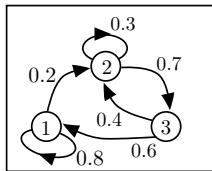
Examples:



[A]



[B]



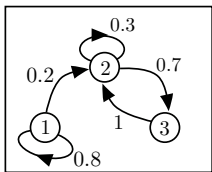
[C]

[A] is **not irreducible**. It cannot go from (2) to (1).

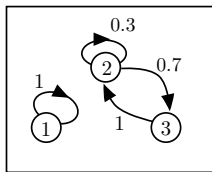
Irreducibility.

Definition A Markov chain is **irreducible** if it can go from every state i to every state j (possibly in multiple steps).

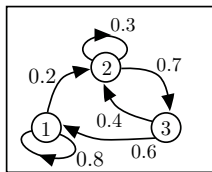
Examples:



[A]



[B]



[C]

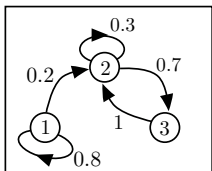
[A] is **not irreducible**. It cannot go from (2) to (1).

[B] is

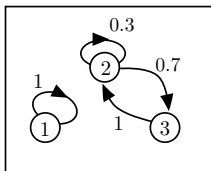
Irreducibility.

Definition A Markov chain is **irreducible** if it can go from every state i to every state j (possibly in multiple steps).

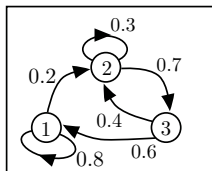
Examples:



[A]



[B]



[C]

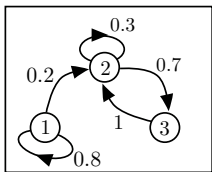
[A] is **not irreducible**. It cannot go from (2) to (1).

[B] is **not irreducible**.

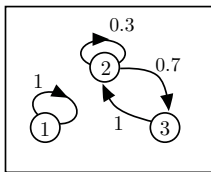
Irreducibility.

Definition A Markov chain is **irreducible** if it can go from every state i to every state j (possibly in multiple steps).

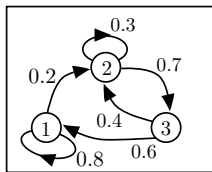
Examples:



[A]



[B]



[C]

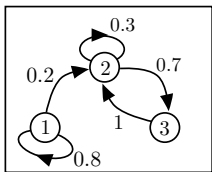
[A] is **not irreducible**. It cannot go from (2) to (1).

[B] is **not irreducible**. It cannot go from (2) to (1).

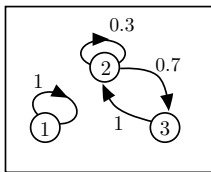
Irreducibility.

Definition A Markov chain is **irreducible** if it can go from every state i to every state j (possibly in multiple steps).

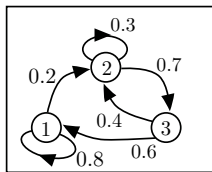
Examples:



[A]



[B]



[C]

[A] is **not irreducible**. It cannot go from (2) to (1).

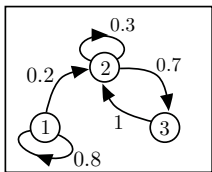
[B] is **not irreducible**. It cannot go from (2) to (1).

[C] is

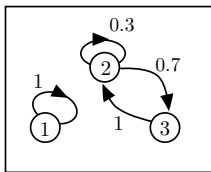
Irreducibility.

Definition A Markov chain is **irreducible** if it can go from every state i to every state j (possibly in multiple steps).

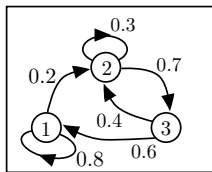
Examples:



[A]



[B]



[C]

[A] is **not irreducible**. It cannot go from (2) to (1).

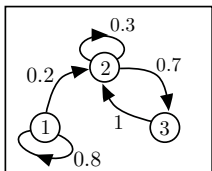
[B] is **not irreducible**. It cannot go from (2) to (1).

[C] is **irreducible**.

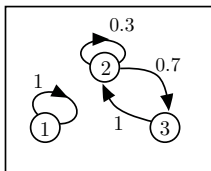
Irreducibility.

Definition A Markov chain is **irreducible** if it can go from every state i to every state j (possibly in multiple steps).

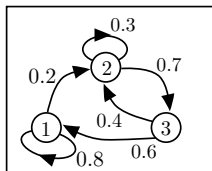
Examples:



[A]



[B]



[C]

[A] is **not irreducible**. It cannot go from (2) to (1).

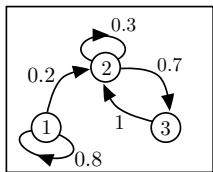
[B] is **not irreducible**. It cannot go from (2) to (1).

[C] is **irreducible**. It can go from every i to every j .

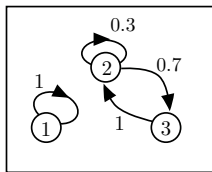
Irreducibility.

Definition A Markov chain is **irreducible** if it can go from every state i to every state j (possibly in multiple steps).

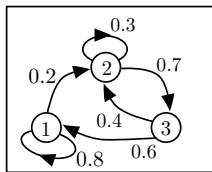
Examples:



[A]



[B]



[C]

[A] is **not irreducible**. It cannot go from (2) to (1).

[B] is **not irreducible**. It cannot go from (2) to (1).

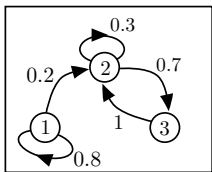
[C] is **irreducible**. It can go from every i to every j .

If you consider the graph with arrows when $P(i,j) > 0$,

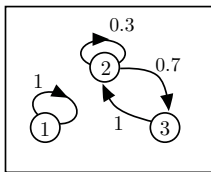
Irreducibility.

Definition A Markov chain is **irreducible** if it can go from every state i to every state j (possibly in multiple steps).

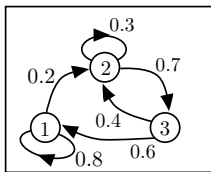
Examples:



[A]



[B]



[C]

[A] is **not irreducible**. It cannot go from (2) to (1).

[B] is **not irreducible**. It cannot go from (2) to (1).

[C] is **irreducible**. It can go from every i to every j .

If you consider the graph with arrows when $P(i,j) > 0$, irreducible means that there is a single (strongly) connected component.

Existence and uniqueness of Invariant Distribution

Existence and uniqueness of Invariant Distribution

Theorem A finite irreducible Markov chain has one and only one invariant distribution.

Existence and uniqueness of Invariant Distribution

Theorem A finite irreducible Markov chain has one and only one invariant distribution.

That is, there is a unique positive vector $\pi = [\pi(1), \dots, \pi(K)]$ such that $\pi P = \pi$ and $\sum_k \pi(k) = 1$.

Existence and uniqueness of Invariant Distribution

Theorem A finite irreducible Markov chain has one and only one invariant distribution.

That is, there is a unique positive vector $\pi = [\pi(1), \dots, \pi(K)]$ such that $\pi P = \pi$ and $\sum_k \pi(k) = 1$.

Existence and uniqueness of Invariant Distribution

Theorem A finite irreducible Markov chain has one and only one invariant distribution.

That is, there is a unique positive vector $\pi = [\pi(1), \dots, \pi(K)]$ such that $\pi P = \pi$ and $\sum_k \pi(k) = 1$.

Ok.

Existence and uniqueness of Invariant Distribution

Theorem A finite irreducible Markov chain has one and only one invariant distribution.

That is, there is a unique positive vector $\pi = [\pi(1), \dots, \pi(K)]$ such that $\pi P = \pi$ and $\sum_k \pi(k) = 1$.

Ok. Now.

Existence and uniqueness of Invariant Distribution

Theorem A finite irreducible Markov chain has one and only one invariant distribution.

That is, there is a unique positive vector $\pi = [\pi(1), \dots, \pi(K)]$ such that $\pi P = \pi$ and $\sum_k \pi(k) = 1$.

Ok. Now.

Only one stationary distribution if irreducible (or connected.)

Long Term Fraction of Time in States

Long Term Fraction of Time in States

Theorem Let X_n be an irreducible Markov chain with invariant distribution π .

Long Term Fraction of Time in States

Theorem Let X_n be an irreducible Markov chain with invariant distribution π .

Then, for all i ,

$$\frac{1}{n} \sum_{m=0}^{n-1} 1\{X_m = i\} \rightarrow \pi(i), \text{ as } n \rightarrow \infty.$$

Long Term Fraction of Time in States

Theorem Let X_n be an irreducible Markov chain with invariant distribution π .

Then, for all i ,

$$\frac{1}{n} \sum_{m=0}^{n-1} 1\{X_m = i\} \rightarrow \pi(i), \text{ as } n \rightarrow \infty.$$

The left-hand side is the fraction of time that $X_m = i$ during steps $0, 1, \dots, n-1$.

Long Term Fraction of Time in States

Theorem Let X_n be an irreducible Markov chain with invariant distribution π .

Then, for all i ,

$$\frac{1}{n} \sum_{m=0}^{n-1} 1\{X_m = i\} \rightarrow \pi(i), \text{ as } n \rightarrow \infty.$$

The left-hand side is the fraction of time that $X_m = i$ during steps $0, 1, \dots, n-1$. Thus, this fraction of time approaches $\pi(i)$.

Long Term Fraction of Time in States

Theorem Let X_n be an irreducible Markov chain with invariant distribution π .

Then, for all i ,

$$\frac{1}{n} \sum_{m=0}^{n-1} 1\{X_m = i\} \rightarrow \pi(i), \text{ as } n \rightarrow \infty.$$

The left-hand side is the fraction of time that $X_m = i$ during steps $0, 1, \dots, n-1$. Thus, this fraction of time approaches $\pi(i)$.

Proof: Lecture note 21 gives a plausibility argument.



Long Term Fraction of Time in States

Theorem Let X_n be an irreducible Markov chain with invariant distribution π . Then, for all i , $\frac{1}{n} \sum_{m=0}^{n-1} \mathbf{1}\{X_m = i\} \rightarrow \pi(i)$, as $n \rightarrow \infty$.

Long Term Fraction of Time in States

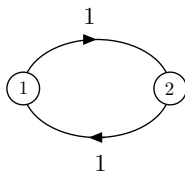
Theorem Let X_n be an irreducible Markov chain with invariant distribution π . Then, for all i , $\frac{1}{n} \sum_{m=0}^{n-1} \mathbf{1}\{X_m = i\} \rightarrow \pi(i)$, as $n \rightarrow \infty$.

Example 1:

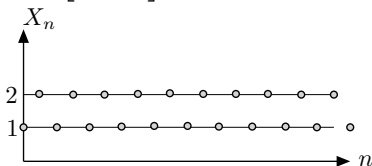
Long Term Fraction of Time in States

Theorem Let X_n be an irreducible Markov chain with invariant distribution π . Then, for all i , $\frac{1}{n} \sum_{m=0}^{n-1} 1\{X_m = i\} \rightarrow \pi(i)$, as $n \rightarrow \infty$.

Example 1:



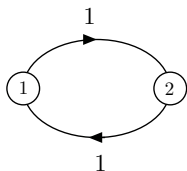
$$P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix} \quad \pi P = \pi \Rightarrow \pi = [1/2, 1/2]$$



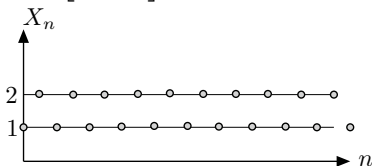
Long Term Fraction of Time in States

Theorem Let X_n be an irreducible Markov chain with invariant distribution π . Then, for all i , $\frac{1}{n} \sum_{m=0}^{n-1} 1\{X_m = i\} \rightarrow \pi(i)$, as $n \rightarrow \infty$.

Example 1:



$$P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix} \quad \pi P = \pi \Rightarrow \pi = [1/2, 1/2]$$

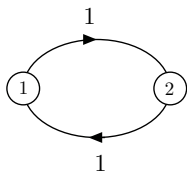


The fraction of time in state 1

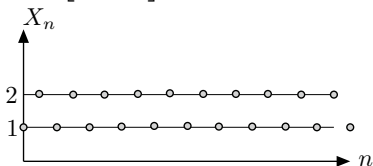
Long Term Fraction of Time in States

Theorem Let X_n be an irreducible Markov chain with invariant distribution π . Then, for all i , $\frac{1}{n} \sum_{m=0}^{n-1} \mathbf{1}\{X_m = i\} \rightarrow \pi(i)$, as $n \rightarrow \infty$.

Example 1:



$$P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix} \quad \pi P = \pi \Rightarrow \pi = [1/2, 1/2]$$

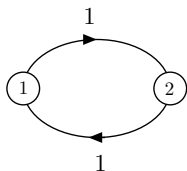


The fraction of time in state 1 converges to $1/2$,

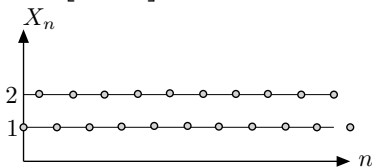
Long Term Fraction of Time in States

Theorem Let X_n be an irreducible Markov chain with invariant distribution π . Then, for all i , $\frac{1}{n} \sum_{m=0}^{n-1} \mathbf{1}\{X_m = i\} \rightarrow \pi(i)$, as $n \rightarrow \infty$.

Example 1:



$$P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix} \quad \pi P = \pi \Rightarrow \pi = [1/2, 1/2]$$



The fraction of time in state 1 converges to $1/2$, which is $\pi(1)$.

Long Term Fraction of Time in States

Theorem Let X_n be an irreducible Markov chain with invariant distribution π . Then, for all i , $\frac{1}{n} \sum_{m=0}^{n-1} \mathbf{1}\{X_m = i\} \rightarrow \pi(i)$, as $n \rightarrow \infty$.

Long Term Fraction of Time in States

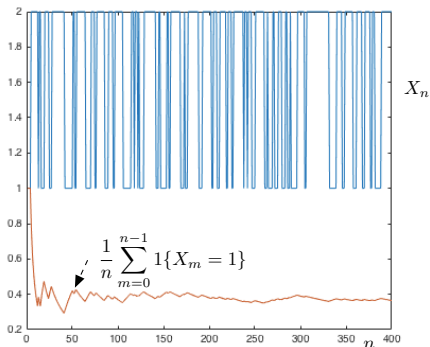
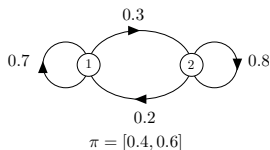
Theorem Let X_n be an irreducible Markov chain with invariant distribution π . Then, for all i , $\frac{1}{n} \sum_{m=0}^{n-1} \mathbf{1}\{X_m = i\} \rightarrow \pi(i)$, as $n \rightarrow \infty$.

Example 2:

Long Term Fraction of Time in States

Theorem Let X_n be an irreducible Markov chain with invariant distribution π . Then, for all i , $\frac{1}{n} \sum_{m=0}^{n-1} 1\{X_m = i\} \rightarrow \pi(i)$, as $n \rightarrow \infty$.

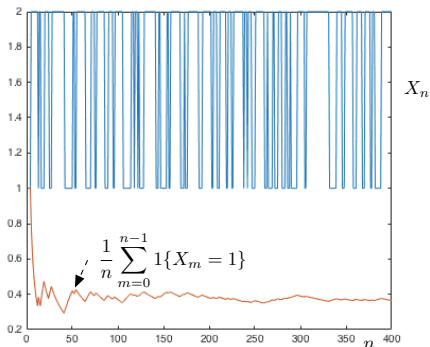
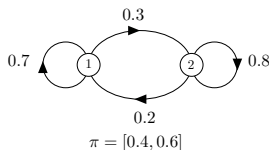
Example 2:



Long Term Fraction of Time in States

Theorem Let X_n be an irreducible Markov chain with invariant distribution π . Then, for all i , $\frac{1}{n} \sum_{m=0}^{n-1} 1\{X_m = i\} \rightarrow \pi(i)$, as $n \rightarrow \infty$.

Example 2:



Convergence to Invariant Distribution

Convergence to Invariant Distribution

Question: Assume that the MC is irreducible.

Convergence to Invariant Distribution

Question: Assume that the MC is irreducible. Does π_n approach the unique invariant distribution π ?

Convergence to Invariant Distribution

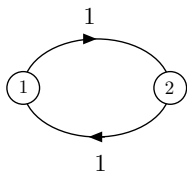
Question: Assume that the MC is irreducible. Does π_n approach the unique invariant distribution π ?

Answer: Not necessarily. Here is an example:

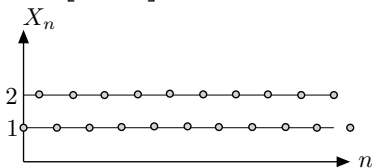
Convergence to Invariant Distribution

Question: Assume that the MC is irreducible. Does π_n approach the unique invariant distribution π ?

Answer: Not necessarily. Here is an example:



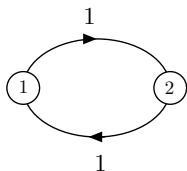
$$P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix} \quad \pi P = \pi \Rightarrow \pi = [1/2, 1/2]$$



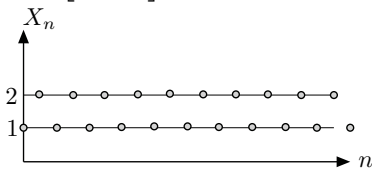
Convergence to Invariant Distribution

Question: Assume that the MC is irreducible. Does π_n approach the unique invariant distribution π ?

Answer: Not necessarily. Here is an example:



$$P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix} \quad \pi P = \pi \Rightarrow \pi = [1/2, 1/2]$$

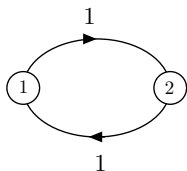


Assume $X_0 = 1$.

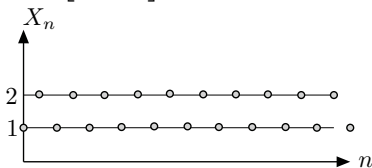
Convergence to Invariant Distribution

Question: Assume that the MC is irreducible. Does π_n approach the unique invariant distribution π ?

Answer: Not necessarily. Here is an example:



$$P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix} \quad \pi P = \pi \Rightarrow \pi = [1/2, 1/2]$$

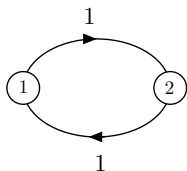


Assume $X_0 = 1$. Then $X_1 = 2$,

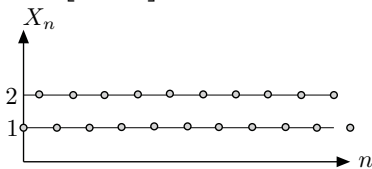
Convergence to Invariant Distribution

Question: Assume that the MC is irreducible. Does π_n approach the unique invariant distribution π ?

Answer: Not necessarily. Here is an example:



$$P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix} \quad \pi P = \pi \Rightarrow \pi = [1/2, 1/2]$$

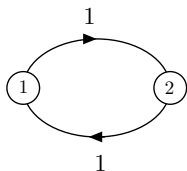


Assume $X_0 = 1$. Then $X_1 = 2, X_2 = 1,$

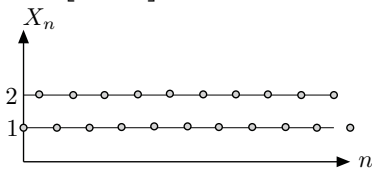
Convergence to Invariant Distribution

Question: Assume that the MC is irreducible. Does π_n approach the unique invariant distribution π ?

Answer: Not necessarily. Here is an example:



$$P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix} \quad \pi P = \pi \Rightarrow \pi = [1/2, 1/2]$$

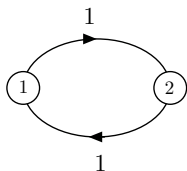


Assume $X_0 = 1$. Then $X_1 = 2, X_2 = 1, X_3 = 2, \dots$

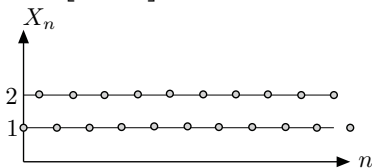
Convergence to Invariant Distribution

Question: Assume that the MC is irreducible. Does π_n approach the unique invariant distribution π ?

Answer: Not necessarily. Here is an example:



$$P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix} \quad \pi P = \pi \Rightarrow \pi = [1/2, 1/2]$$



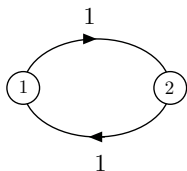
Assume $X_0 = 1$. Then $X_1 = 2, X_2 = 1, X_3 = 2, \dots$

Thus, if $\pi_0 = [1, 0]$,

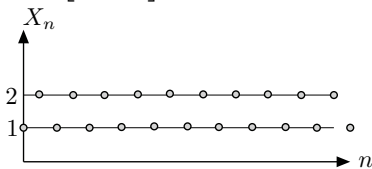
Convergence to Invariant Distribution

Question: Assume that the MC is irreducible. Does π_n approach the unique invariant distribution π ?

Answer: Not necessarily. Here is an example:



$$P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix} \quad \pi P = \pi \Rightarrow \pi = [1/2, 1/2]$$



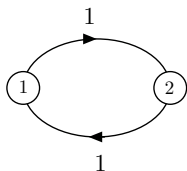
Assume $X_0 = 1$. Then $X_1 = 2, X_2 = 1, X_3 = 2, \dots$

Thus, if $\pi_0 = [1, 0], \pi_1 = [0, 1],$

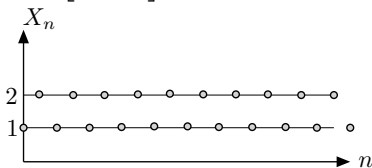
Convergence to Invariant Distribution

Question: Assume that the MC is irreducible. Does π_n approach the unique invariant distribution π ?

Answer: Not necessarily. Here is an example:



$$P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix} \quad \pi P = \pi \Rightarrow \pi = [1/2, 1/2]$$



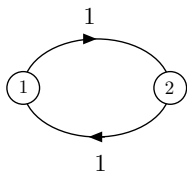
Assume $X_0 = 1$. Then $X_1 = 2, X_2 = 1, X_3 = 2, \dots$

Thus, if $\pi_0 = [1, 0], \pi_1 = [0, 1], \pi_2 = [1, 0],$

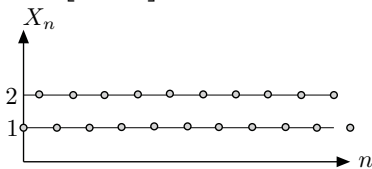
Convergence to Invariant Distribution

Question: Assume that the MC is irreducible. Does π_n approach the unique invariant distribution π ?

Answer: Not necessarily. Here is an example:



$$P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix} \quad \pi P = \pi \Rightarrow \pi = [1/2, 1/2]$$



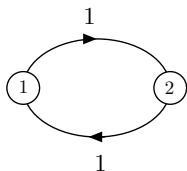
Assume $X_0 = 1$. Then $X_1 = 2, X_2 = 1, X_3 = 2, \dots$

Thus, if $\pi_0 = [1, 0], \pi_1 = [0, 1], \pi_2 = [1, 0], \pi_3 = [0, 1], \dots$, etc.

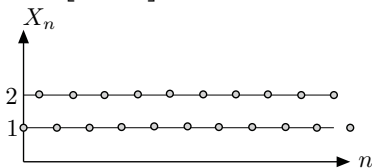
Convergence to Invariant Distribution

Question: Assume that the MC is irreducible. Does π_n approach the unique invariant distribution π ?

Answer: Not necessarily. Here is an example:



$$P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix} \quad \pi P = \pi \Rightarrow \pi = [1/2, 1/2]$$



Assume $X_0 = 1$. Then $X_1 = 2, X_2 = 1, X_3 = 2, \dots$

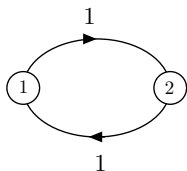
Thus, if $\pi_0 = [1, 0], \pi_1 = [0, 1], \pi_2 = [1, 0], \pi_3 = [0, 1], \dots$, etc.

Hence, π_n does not converge to $\pi = [1/2, 1/2]$.

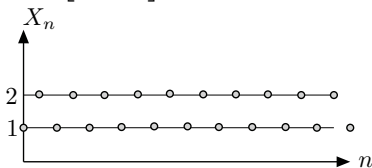
Convergence to Invariant Distribution

Question: Assume that the MC is irreducible. Does π_n approach the unique invariant distribution π ?

Answer: Not necessarily. Here is an example:



$$P = \begin{bmatrix} 0 & 1 \\ 1 & 0 \end{bmatrix} \quad \pi P = \pi \Rightarrow \pi = [1/2, 1/2]$$



Assume $X_0 = 1$. Then $X_1 = 2, X_2 = 1, X_3 = 2, \dots$

Thus, if $\pi_0 = [1, 0], \pi_1 = [0, 1], \pi_2 = [1, 0], \pi_3 = [0, 1], \dots$, etc.

Hence, π_n does not converge to $\pi = [1/2, 1/2]$.

Notice, all cycles or closed walks have even length.

Convergence to stationary distribution.

Theorem Let X_n be an irreducible Markov chain with invariant distribution π . Then, for all i , $\frac{1}{n} \sum_{m=0}^{n-1} \mathbf{1}\{X_m = i\} \rightarrow \pi(i)$, as $n \rightarrow \infty$.

Convergence to stationary distribution.

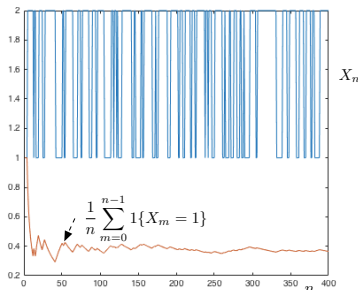
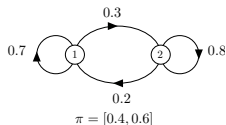
Theorem Let X_n be an irreducible Markov chain with invariant distribution π . Then, for all i , $\frac{1}{n} \sum_{m=0}^{n-1} \mathbf{1}\{X_m = i\} \rightarrow \pi(i)$, as $n \rightarrow \infty$.

Example 2:

Convergence to stationary distribution.

Theorem Let X_n be an irreducible Markov chain with invariant distribution π . Then, for all i , $\frac{1}{n} \sum_{m=0}^{n-1} 1\{X_m = i\} \rightarrow \pi(i)$, as $n \rightarrow \infty$.

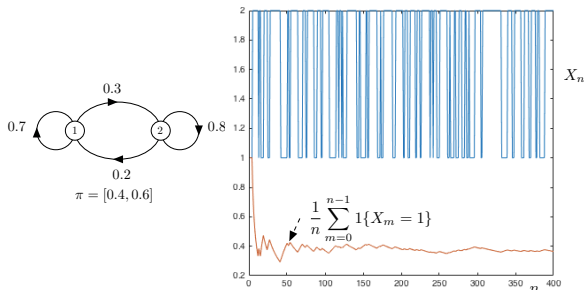
Example 2:



Convergence to stationary distribution.

Theorem Let X_n be an irreducible Markov chain with invariant distribution π . Then, for all i , $\frac{1}{n} \sum_{m=0}^{n-1} 1\{X_m = i\} \rightarrow \pi(i)$, as $n \rightarrow \infty$.

Example 2:

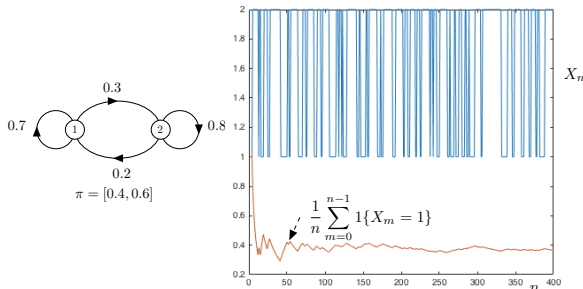


As n gets large the probability of being in state 1 approaches 0.4.
(The stationary distribution.)

Convergence to stationary distribution.

Theorem Let X_n be an irreducible Markov chain with invariant distribution π . Then, for all i , $\frac{1}{n} \sum_{m=0}^{n-1} 1\{X_m = i\} \rightarrow \pi(i)$, as $n \rightarrow \infty$.

Example 2:



As n gets large the probability of being in state 1 approaches 0.4. (The stationary distribution.) Notice cycles of length 1 and 2.

Periodicity

Periodicity

Definition: Periodicity is gcd of the lengths of all closed walks in irreducible chain.

Periodicity

Definition: Periodicity is gcd of the lengths of all closed walks in irreducible chain. Previous example: 2.

Periodicity

Definition: Periodicity is gcd of the lengths of all closed walks in irreducible chain. Previous example: 2.

Definition

Periodicity

Definition: Periodicity is gcd of the lengths of all closed walks in irreducible chain. Previous example: 2.

Definition If periodicity is 1, Markov chain is said to be **aperiodic**.

Periodicity

Definition: Periodicity is gcd of the lengths of all closed walks in irreducible chain. Previous example: 2.

Definition If periodicity is 1, Markov chain is said to be **aperiodic**. Otherwise, it is periodic.

Periodicity

Definition: Periodicity is gcd of the lengths of all closed walks in irreducible chain. Previous example: 2.

Definition If periodicity is 1, Markov chain is said to be **aperiodic**. Otherwise, it is periodic.

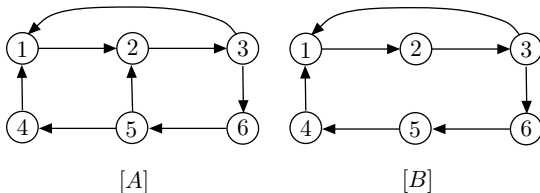
Example

Periodicity

Definition: Periodicity is gcd of the lengths of all closed walks in irreducible chain. Previous example: 2.

Definition If periodicity is 1, Markov chain is said to be **aperiodic**. Otherwise, it is periodic.

Example

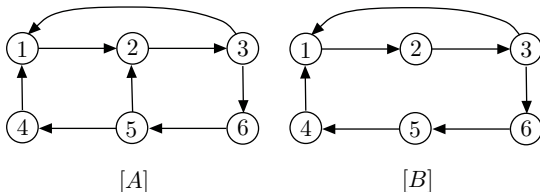


Periodicity

Definition: Periodicity is gcd of the lengths of all closed walks in irreducible chain. Previous example: 2.

Definition If periodicity is 1, Markov chain is said to be **aperiodic**. Otherwise, it is periodic.

Example



Which one converges to stationary?

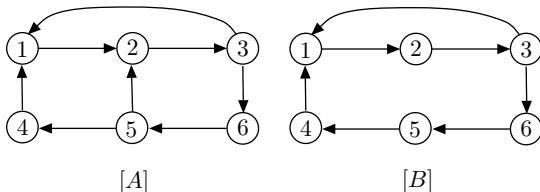
- (A) [A]
- (B) [B]
- (C) both
- (D) neither.

Periodicity

Definition: Periodicity is gcd of the lengths of all closed walks in irreducible chain. Previous example: 2.

Definition If periodicity is 1, Markov chain is said to be **aperiodic**. Otherwise, it is periodic.

Example



Which one converges to stationary?

- (A) [A]
- (B) [B]
- (C) both
- (D) neither.

(A).

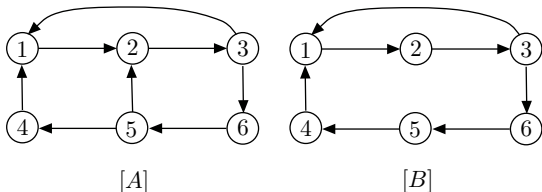
[A]:

Periodicity

Definition: Periodicity is gcd of the lengths of all closed walks in irreducible chain. Previous example: 2.

Definition If periodicity is 1, Markov chain is said to be **aperiodic**. Otherwise, it is periodic.

Example



Which one converges to stationary?

- (A) [A]
 - (B) [B]
 - (C) both
 - (D) neither.
- (A).

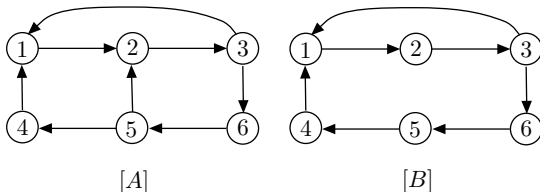
[A]: Closed walks of length 3 and length 4

Periodicity

Definition: Periodicity is gcd of the lengths of all closed walks in irreducible chain. Previous example: 2.

Definition If periodicity is 1, Markov chain is said to be **aperiodic**. Otherwise, it is periodic.

Example



Which one converges to stationary?

- (A) [A]
 - (B) [B]
 - (C) both
 - (D) neither.
- (A).

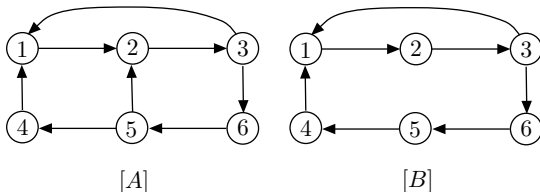
[A]: Closed walks of length 3 and length 4 \implies periodicity = 1.

Periodicity

Definition: Periodicity is gcd of the lengths of all closed walks in irreducible chain. Previous example: 2.

Definition If periodicity is 1, Markov chain is said to be **aperiodic**. Otherwise, it is periodic.

Example



Which one converges to stationary?

- (A) [A]
 - (B) [B]
 - (C) both
 - (D) neither.
- (A).

[A]: Closed walks of length 3 and length 4 \implies periodicity = 1.

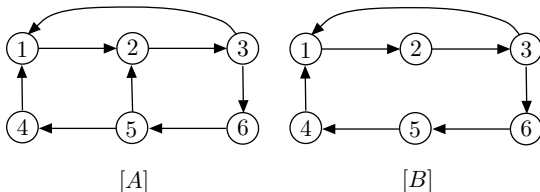
[B]:

Periodicity

Definition: Periodicity is gcd of the lengths of all closed walks in irreducible chain. Previous example: 2.

Definition If periodicity is 1, Markov chain is said to be **aperiodic**. Otherwise, it is periodic.

Example



Which one converges to stationary?

- (A) [A]
 - (B) [B]
 - (C) both
 - (D) neither.
- (A).

[A]: Closed walks of length 3 and length 4 \implies periodicity = 1.

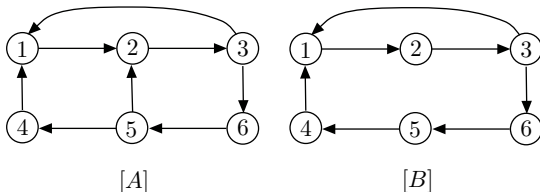
[B]: All closed walks multiple of 3

Periodicity

Definition: Periodicity is gcd of the lengths of all closed walks in irreducible chain. Previous example: 2.

Definition If periodicity is 1, Markov chain is said to be **aperiodic**. Otherwise, it is periodic.

Example



Which one converges to stationary?

- (A) [A]
 - (B) [B]
 - (C) both
 - (D) neither.
- (A).

[A]: Closed walks of length 3 and length 4 \implies periodicity = 1.

[B]: All closed walks multiple of 3 \implies periodicity = 3 .

Convergence of π_n

Convergence of π_n

Theorem Let X_n be an irreducible and aperiodic Markov chain with invariant distribution π .

Convergence of π_n

Theorem Let X_n be an irreducible and aperiodic Markov chain with invariant distribution π . Then, for all $i \in \mathcal{X}$,

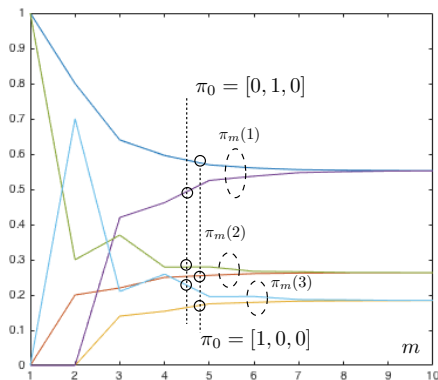
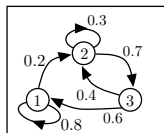
$$\pi_n(i) \rightarrow \pi(i), \text{ as } n \rightarrow \infty.$$

Convergence of π_n

Theorem Let X_n be an irreducible and aperiodic Markov chain with invariant distribution π . Then, for all $i \in \mathcal{X}$,

$$\pi_n(i) \rightarrow \pi(i), \text{ as } n \rightarrow \infty.$$

Example



Convergence of π_n

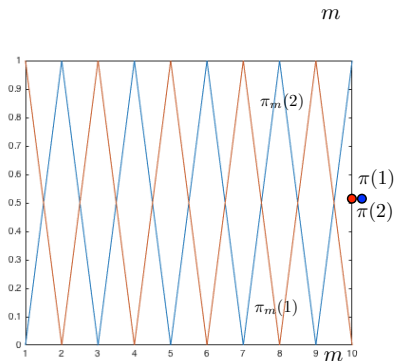
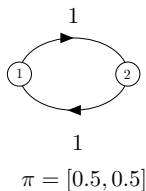
Theorem Let X_n be an irreducible and aperiodic Markov chain with invariant distribution π .

Convergence of π_n

Theorem Let X_n be an irreducible and aperiodic Markov chain with invariant distribution π . Then, for all $i \in \mathcal{X}$,

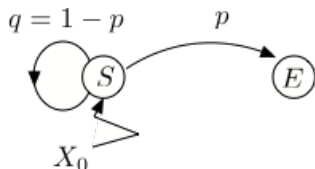
$$\pi_n(i) \rightarrow \pi(i), \text{ as } n \rightarrow \infty.$$

Non Example: periodic chain



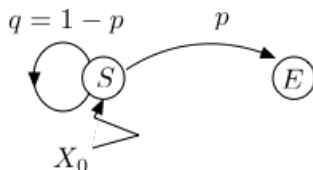
First Passage Time - Example 1. Poll

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips, on average?



First Passage Time - Example 1. Poll

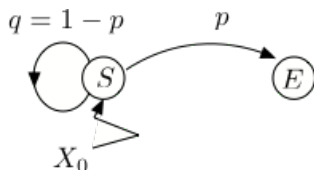
Let's flip a coin with $Pr[H] = p$ until we get H . How many flips, on average?



Let $\beta(S)$ be the average time until E , starting from S .

First Passage Time - Example 1. Poll

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips, on average?

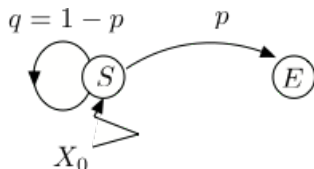


Let $\beta(S)$ be the average time until E , starting from S .

What is correct?

First Passage Time - Example 1. Poll

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips, on average?



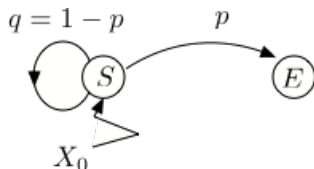
Let $\beta(S)$ be the average time until E , starting from S .

What is correct?

- (A) $\beta(S)$ is at least 1.
- (B) From S , in one step, go to S with prob. $q = 1 - p$
- (C) From S , in one step, go to E with prob. p .
- (D) If you go back to S , you are back at S .
- (D) $\beta(S) = 1 + q\beta(S) + p0$.

First Passage Time - Example 1. Poll

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips, on average?



Let $\beta(S)$ be the average time until E , starting from S .

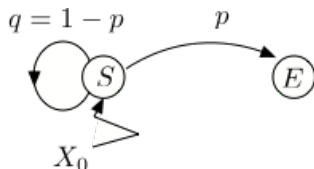
What is correct?

- (A) $\beta(S)$ is at least 1.
- (B) From S , in one step, go to S with prob. $q = 1 - p$
- (C) From S , in one step, go to E with prob. p .
- (D) If you go back to S , you are back at S .
- (D) $\beta(S) = 1 + q\beta(S) + p0$.

All are correct. (D) is the “Markov property.” Only know where you are.

First Passage Time - Example 1. Poll

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips, on average?



Let $\beta(S)$ be the average time until E , starting from S .

What is correct?

- (A) $\beta(S)$ is at least 1.
- (B) From S , in one step, go to S with prob. $q = 1 - p$
- (C) From S , in one step, go to E with prob. p .
- (D) If you go back to S , you are back at S .
- (D) $\beta(S) = 1 + q\beta(S) + p0$.

All are correct. (D) is the “Markov property.” Only know where you are.

Hitting Time - Example 1

Hitting Time - Example 1

Let's flip a coin with $\Pr[H] = p$ until we get H .

Hitting Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips,

Hitting Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips, on average (in expectation)?

Hitting Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips, on average (in expectation)?

Let's define a Markov chain:

- ▶ $X_0 = S$

Hitting Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips, on average (in expectation)?

Let's define a Markov chain:

- ▶ $X_0 = S$ (start)

Hitting Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips, on average (in expectation)?

Let's define a Markov chain:

- ▶ $X_0 = S$ (start)
- ▶ $X_n = S$ for $n \geq 1$, if last flip was T and no H yet

Hitting Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips, on average (in expectation)?

Let's define a Markov chain:

- ▶ $X_0 = S$ (start)
- ▶ $X_n = S$ for $n \geq 1$, if last flip was T and no H yet
- ▶ $X_n = E$ for $n \geq 1$, if we already got H

Hitting Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips, on average (in expectation)?

Let's define a Markov chain:

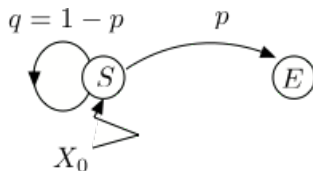
- ▶ $X_0 = S$ (start)
- ▶ $X_n = S$ for $n \geq 1$, if last flip was T and no H yet
- ▶ $X_n = E$ for $n \geq 1$, if we already got H (end)

Hitting Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips, on average (in expectation)?

Let's define a Markov chain:

- ▶ $X_0 = S$ (start)
- ▶ $X_n = S$ for $n \geq 1$, if last flip was T and no H yet
- ▶ $X_n = E$ for $n \geq 1$, if we already got H (end)

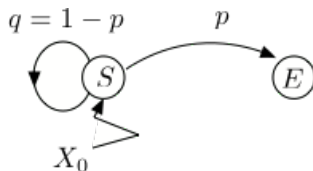


Hitting Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips, on average (in expectation)?

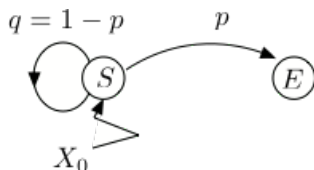
Let's define a Markov chain:

- ▶ $X_0 = S$ (start)
- ▶ $X_n = S$ for $n \geq 1$, if last flip was T and no H yet
- ▶ $X_n = E$ for $n \geq 1$, if we already got H (end)



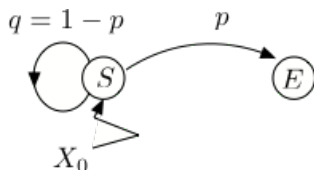
Hitting Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips, on average (in expectation)?



Hitting Time - Example 1

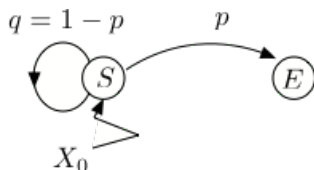
Let's flip a coin with $Pr[H] = p$ until we get H . How many flips, on average (in expectation)?



Let $\beta(S)$ be the expected time until E , starting from S .

Hitting Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips, on average (in expectation)?



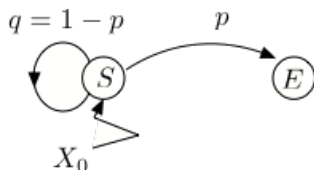
Let $\beta(S)$ be the expected time until E , starting from S .

Then,

$$\beta(S) = 1 + q\beta(S) + p \cdot 0.$$

Hitting Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips, on average (in expectation)?



Let $\beta(S)$ be the expected time until E , starting from S .

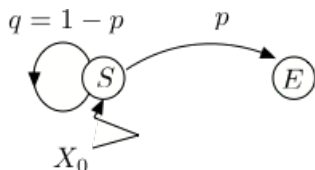
Then,

$$\beta(S) = 1 + q\beta(S) + p \cdot 0.$$

(See next slide.)

Hitting Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips, on average (in expectation)?



Let $\beta(S)$ be the expected time until E , starting from S .

Then,

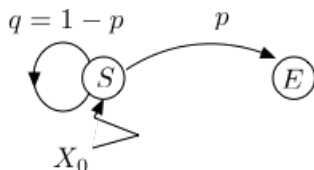
$$\beta(S) = 1 + q\beta(S) + p0.$$

(See next slide.) Hence,

$$\beta(S) = 1 + (1 - p)\beta(S) \implies p\beta(S) = 1,$$

Hitting Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips, on average (in expectation)?



Let $\beta(S)$ be the expected time until E , starting from S .

Then,

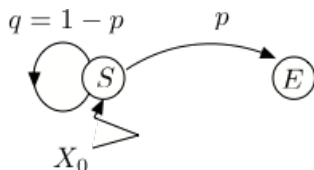
$$\beta(S) = 1 + q\beta(S) + p \cdot 0.$$

(See next slide.) Hence,

$$\beta(S) = 1 + (1 - p)\beta(S) \implies p\beta(S) = 1, \text{ so that } \beta(S) = 1/p.$$

Hitting Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips, on average (in expectation)?



Let $\beta(S)$ be the expected time until E , starting from S .

Then,

$$\beta(S) = 1 + q\beta(S) + p0.$$

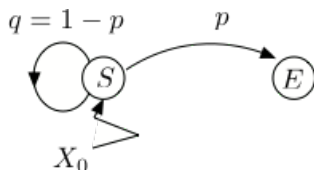
(See next slide.) Hence,

$$\beta(S) = 1 + (1 - p)\beta(S) \implies p\beta(S) = 1, \text{ so that } \beta(S) = 1/p.$$

Note: Time until E is $G(p)$.

Hitting Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips, on average (in expectation)?



Let $\beta(S)$ be the expected time until E , starting from S .

Then,

$$\beta(S) = 1 + q\beta(S) + p \cdot 0.$$

(See next slide.) Hence,

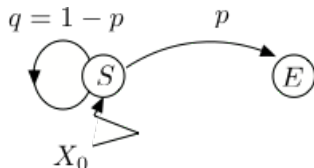
$$\beta(S) = 1 + (1 - p)\beta(S) \implies p\beta(S) = 1, \text{ so that } \beta(S) = 1/p.$$

Note: Time until E is $G(p)$.

The mean of $G(p)$ is $1/p$!!!

First Passage Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips in expectation?



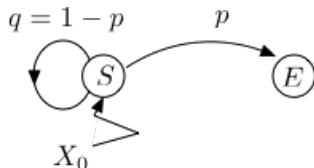
Let $\beta(S)$ be the expected time until E .

Then,

$$\beta(S) = 1 + q\beta(S) + p \cdot 0.$$

First Passage Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips in expectation?



Let $\beta(S)$ be the expected time until E .

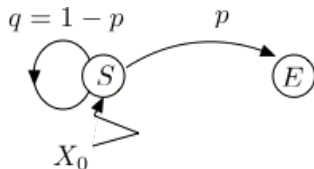
Then,

$$\beta(S) = 1 + q\beta(S) + p \cdot 0.$$

Justification:

First Passage Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips in expectation?



Let $\beta(S)$ be the expected time until E .

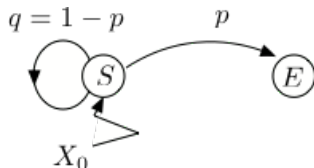
Then,

$$\beta(S) = 1 + q\beta(S) + p \cdot 0.$$

Justification: N – number of steps until E , starting from S .

First Passage Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips in expectation?



Let $\beta(S)$ be the expected time until E .

Then,

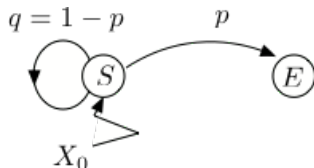
$$\beta(S) = 1 + q\beta(S) + p0.$$

Justification: N – number of steps until E , starting from S .

N' – number of steps until E , after the second visit to S .

First Passage Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips in expectation?



Let $\beta(S)$ be the expected time until E .

Then,

$$\beta(S) = 1 + q\beta(S) + p0.$$

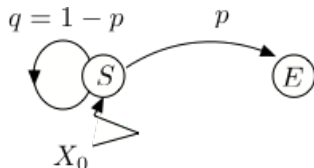
Justification: N – number of steps until E , starting from S .

N' – number of steps until E , after the second visit to S .

And $Z = 1_{\{\text{first flip} = H\}}$.

First Passage Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips in expectation?



Let $\beta(S)$ be the expected time until E .

Then,

$$\beta(S) = 1 + q\beta(S) + p \cdot 0.$$

Justification: N – number of steps until E , starting from S .

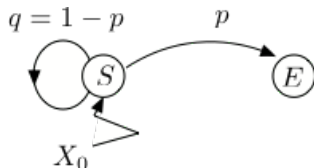
N' – number of steps until E , after the second visit to S .

And $Z = 1_{\{\text{first flip} = H\}}$. Then,

$$N = 1 + (1 - Z) \times N' + Z \times 0.$$

First Passage Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips in expectation?



Let $\beta(S)$ be the expected time until E .

Then,

$$\beta(S) = 1 + q\beta(S) + p \cdot 0.$$

Justification: N – number of steps until E , starting from S .

N' – number of steps until E , after the second visit to S .

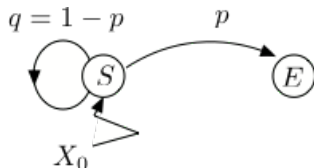
And $Z = 1_{\{\text{first flip} = H\}}$. Then,

$$N = 1 + (1 - Z) \times N' + Z \times 0.$$

Z and N' are “independent.”

First Passage Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips in expectation?



Let $\beta(S)$ be the expected time until E .

Then,

$$\beta(S) = 1 + q\beta(S) + p \cdot 0.$$

Justification: N – number of steps until E , starting from S .

N' – number of steps until E , after the second visit to S .

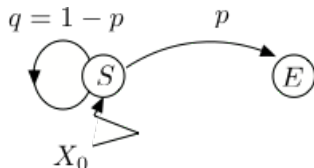
And $Z = 1_{\{\text{first flip} = H\}}$. Then,

$$N = 1 + (1 - Z) \times N' + Z \times 0.$$

Z and N' are “independent.” $E[N'] = E[N] = \beta(S)$.

First Passage Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips in expectation?



Let $\beta(S)$ be the expected time until E .

Then,

$$\beta(S) = 1 + q\beta(S) + p \cdot 0.$$

Justification: N – number of steps until E , starting from S .

N' – number of steps until E , after the second visit to S .

And $Z = 1_{\{\text{first flip} = H\}}$. Then,

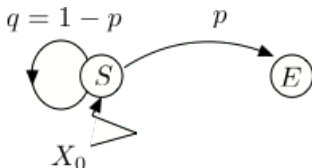
$$N = 1 + (1 - Z) \times N' + Z \times 0.$$

Z and N' are “independent.” $E[N'] = E[N] = \beta(S)$.

Hence, taking expectation,

First Passage Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips in expectation?



Let $\beta(S)$ be the expected time until E .

Then,

$$\beta(S) = 1 + q\beta(S) + p0.$$

Justification: N – number of steps until E , starting from S .

N' – number of steps until E , after the second visit to S .

And $Z = 1_{\{\text{first flip} = H\}}$. Then,

$$N = 1 + (1 - Z) \times N' + Z \times 0.$$

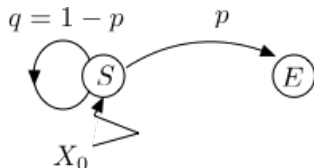
Z and N' are “independent.” $E[N'] = E[N] = \beta(S)$.

Hence, taking expectation,

$$\beta(S) = E[N] = 1 + (1 - p)E[N'] + p0$$

First Passage Time - Example 1

Let's flip a coin with $Pr[H] = p$ until we get H . How many flips in expectation?



Let $\beta(S)$ be the expected time until E .

Then,

$$\beta(S) = 1 + q\beta(S) + p0.$$

Justification: N – number of steps until E , starting from S .

N' – number of steps until E , after the second visit to S .

And $Z = 1_{\{\text{first flip} = H\}}$. Then,

$$N = 1 + (1 - Z) \times N' + Z \times 0.$$

Z and N' are “independent.” $E[N'] = E[N] = \beta(S)$.

Hence, taking expectation,

$$\beta(S) = E[N] = 1 + (1 - p)E[N'] + p0 = 1 + q\beta(S) + p0.$$

Hitting Time - Example 2

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s.

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips,

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average?

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average?

H

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average?

$H T$

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average?

H T H

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average?

H T H T

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average?

H T H T T T H

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average?

H T H T T T H T H

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average?

H T H T T T H T H T H

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average?

H T H T T T H T H T H T T H

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average?

H T H T T T H T H T H T T H T

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average?

H T H T T T H T H T H T T H T H

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average?

H T H T T T H T H T H T T H T H H

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average?

H T H T T T H T H T H T T H T H H

Let's define a Markov chain:

- ▶ $X_0 = S$

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average?

H T H T T T H T H T H T T H T H H

Let's define a Markov chain:

- ▶ $X_0 = S$ (start)
- ▶ $X_n = E$, if we already got two consecutive H s

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average?

H T H T T T H T H T H T T H T H H

Let's define a Markov chain:

- ▶ $X_0 = S$ (start)
- ▶ $X_n = E$, if we already got two consecutive H s (end)

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average?

H T H T T T H T H T H T T H T H H

Let's define a Markov chain:

- ▶ $X_0 = S$ (start)
- ▶ $X_n = E$, if we already got two consecutive H s (end)
- ▶ $X_n = T$, if last flip was T and we are not done

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average?

H T H T T T H T H T H T T H T H H

Let's define a Markov chain:

- ▶ $X_0 = S$ (start)
- ▶ $X_n = E$, if we already got two consecutive H s (end)
- ▶ $X_n = T$, if last flip was T and we are not done
- ▶ $X_n = H$, if last flip was H and we are not done

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average?

H T H T T T H T H T H T T H T H H

Let's define a Markov chain:

- ▶ $X_0 = S$ (start)
- ▶ $X_n = E$, if we already got two consecutive H s (end)
- ▶ $X_n = T$, if last flip was T and we are not done
- ▶ $X_n = H$, if last flip was H and we are not done

Hitting Time - Example 2

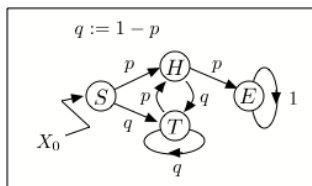
Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average?

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average? Here is a picture:

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average? Here is a picture:



S : Start

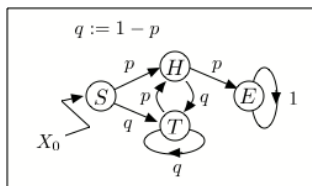
H : Last flip = H

T : Last flip = T

E : Done

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average? Here is a picture:



S : Start

H : Last flip = H

T : Last flip = T

E : Done

Which one is correct?

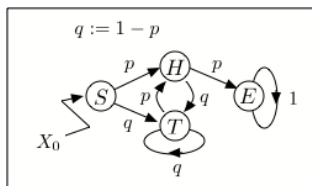
(A) $\beta(S) = 1 + p\beta(H) + q\beta(T)$

(B) $\beta(S) = p\beta(H) + q\beta(T)$

(C) $\beta(S) = \beta(S) + q\beta(T) + p\beta(H)$.

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average? Here is a picture:



S : Start

H : Last flip = H

T : Last flip = T

E : Done

Which one is correct?

(A) $\beta(S) = 1 + p\beta(H) + q\beta(T)$

(B) $\beta(S) = p\beta(H) + q\beta(T)$

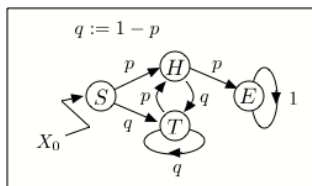
(C) $\beta(S) = \beta(S) + q\beta(T) + p\beta(H)$.

(A) Expected time from S to E .

$$\beta(S) = Pr[H]E[\beta(S)|H] + Pr[T]E[\beta(S)|T]$$

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average? Here is a picture:



S : Start

H : Last flip = H

T : Last flip = T

E : Done

Which one is correct?

(A) $\beta(S) = 1 + p\beta(H) + q\beta(T)$

(B) $\beta(S) = p\beta(H) + q\beta(T)$

(C) $\beta(S) = \beta(S) + q\beta(T) + p\beta(H)$.

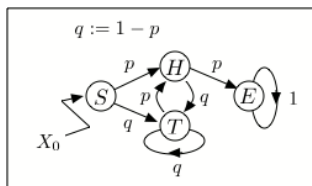
(A) Expected time from S to E .

$$\beta(S) = Pr[H]E[\beta(S)|H] + Pr[T]E[\beta(S)|T]$$

$$\beta(S) = p(1 + \beta(H)) + q(1 + \beta(T))$$

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average? Here is a picture:



S : Start

H : Last flip = H

T : Last flip = T

E : Done

Which one is correct?

(A) $\beta(S) = 1 + p\beta(H) + q\beta(T)$

(B) $\beta(S) = p\beta(H) + q\beta(T)$

(C) $\beta(S) = \beta(S) + q\beta(T) + p\beta(H)$.

(A) Expected time from S to E .

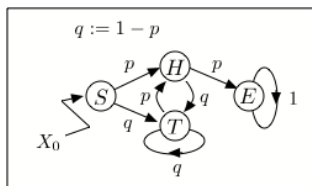
$$\beta(S) = Pr[H]E[\beta(S)|H] + Pr[T]E[\beta(S)|T]$$

$$\beta(S) = p(1 + \beta(H)) + q(1 + \beta(T))$$

$$\beta(S) = 1 + p\beta(H) + q\beta(T)$$

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average? Here is a picture:



S : Start

H : Last flip = H

T : Last flip = T

E : Done

Which one is correct?

(A) $\beta(S) = 1 + p\beta(H) + q\beta(T)$

(B) $\beta(S) = p\beta(H) + q\beta(T)$

(C) $\beta(S) = \beta(S) + q\beta(T) + p\beta(H)$.

(A) Expected time from S to E .

$$\beta(S) = Pr[H]E[\beta(S)|H] + Pr[T]E[\beta(S)|T]$$

$$\beta(S) = p(1 + \beta(H)) + q(1 + \beta(T))$$

$$\beta(S) = 1 + p\beta(H) + q\beta(T)$$

Hitting Time - Example 2

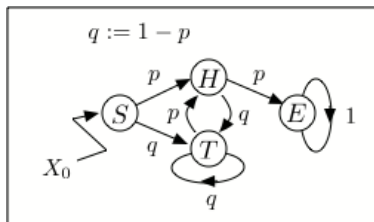
Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average?

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average? Here is a picture:

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average? Here is a picture:



S : Start

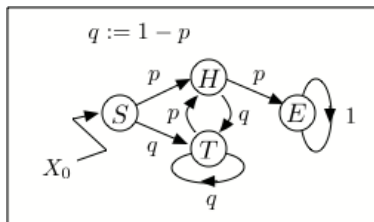
H : Last flip = H

T : Last flip = T

E : Done

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average? Here is a picture:



S : Start

H : Last flip = H

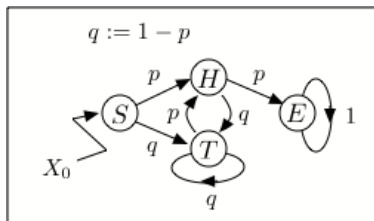
T : Last flip = T

E : Done

Let $\beta(i)$ be the average time from state i until the MC hits state E .

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average? Here is a picture:



S : Start

H : Last flip = H

T : Last flip = T

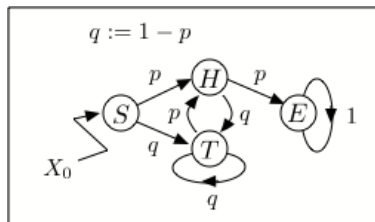
E : Done

Let $\beta(i)$ be the average time from state i until the MC hits state E .

We claim that

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average? Here is a picture:



S : Start

H : Last flip = H

T : Last flip = T

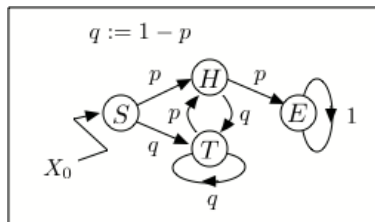
E : Done

Let $\beta(i)$ be the average time from state i until the MC hits state E .

We claim that (these are called the [first step equations](#))

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average? Here is a picture:



S : Start

H : Last flip = H

T : Last flip = T

E : Done

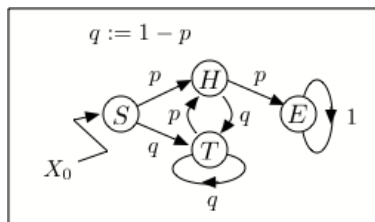
Let $\beta(i)$ be the average time from state i until the MC hits state E .

We claim that (these are called the [first step equations](#))

$$\beta(S) = 1 + p\beta(H) + q\beta(T)$$

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average? Here is a picture:



S : Start

H : Last flip = H

T : Last flip = T

E : Done

Let $\beta(i)$ be the average time from state i until the MC hits state E .

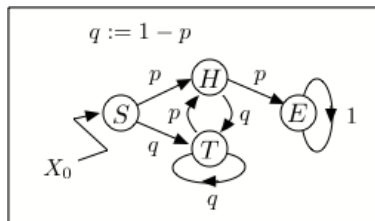
We claim that (these are called the [first step equations](#))

$$\beta(S) = 1 + p\beta(H) + q\beta(T)$$

$$\beta(H) = 1 + p\beta(H) + q\beta(T)$$

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average? Here is a picture:



S : Start

H : Last flip = H

T : Last flip = T

E : Done

Let $\beta(i)$ be the average time from state i until the MC hits state E .

We claim that (these are called the [first step equations](#))

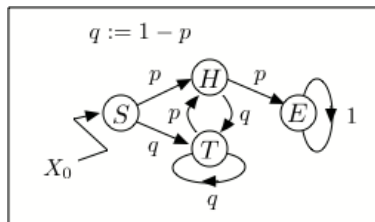
$$\beta(S) = 1 + p\beta(H) + q\beta(T)$$

$$\beta(H) = 1 + p0 + q\beta(T)$$

$$\beta(T) = 1 + p\beta(H) + q\beta(T).$$

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average? Here is a picture:



S : Start

H : Last flip = H

T : Last flip = T

E : Done

Let $\beta(i)$ be the average time from state i until the MC hits state E .

We claim that (these are called the [first step equations](#))

$$\beta(S) = 1 + p\beta(H) + q\beta(T)$$

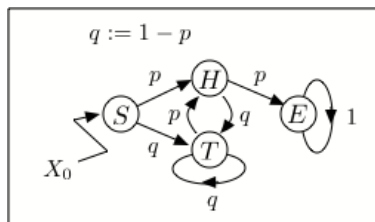
$$\beta(H) = 1 + p0 + q\beta(T)$$

$$\beta(T) = 1 + p\beta(H) + q\beta(T).$$

Solving, we find

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average? Here is a picture:



S : Start

H : Last flip = H

T : Last flip = T

E : Done

Let $\beta(i)$ be the average time from state i until the MC hits state E .

We claim that (these are called the [first step equations](#))

$$\beta(S) = 1 + p\beta(H) + q\beta(T)$$

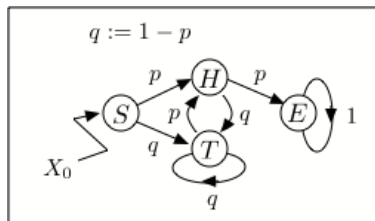
$$\beta(H) = 1 + p\beta(H) + q\beta(T)$$

$$\beta(T) = 1 + p\beta(H) + q\beta(T).$$

Solving, we find $\beta(S) = 2 + 3qp^{-1} + q^2p^{-2}$.

Hitting Time - Example 2

Let's flip a coin with $Pr[H] = p$ until we get two consecutive H s. How many flips, on average? Here is a picture:



S : Start

H : Last flip = H

T : Last flip = T

E : Done

Let $\beta(i)$ be the average time from state i until the MC hits state E .

We claim that (these are called the [first step equations](#))

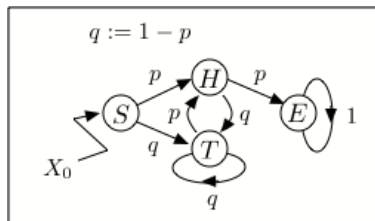
$$\beta(S) = 1 + p\beta(H) + q\beta(T)$$

$$\beta(H) = 1 + p\beta(H) + q\beta(T)$$

$$\beta(T) = 1 + p\beta(H) + q\beta(T).$$

Solving, we find $\beta(S) = 2 + 3qp^{-1} + q^2p^{-2}$. (E.g., $\beta(S) = 6$ if $p = 1/2$.)

Hitting Time - Example 2



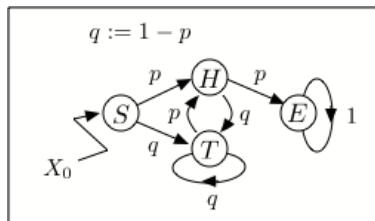
S : Start

H : Last flip = H

T : Last flip = T

E : Done

Hitting Time - Example 2



S : Start

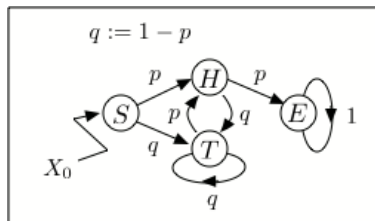
H : Last flip = H

T : Last flip = T

E : Done

Let us justify the first step equation for $\beta(T)$.

Hitting Time - Example 2



S : Start

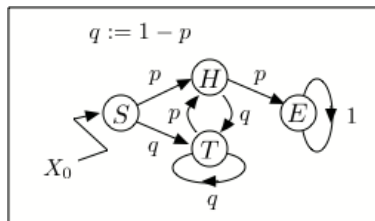
H : Last flip = H

T : Last flip = T

E : Done

Let us justify the first step equation for $\beta(T)$. The others are similar.

Hitting Time - Example 2



S : Start

H : Last flip = H

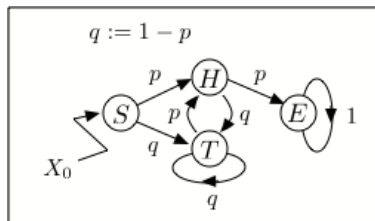
T : Last flip = T

E : Done

Let us justify the first step equation for $\beta(T)$. The others are similar.

$N(T)$ – number of steps, starting from T until the MC hits E .

Hitting Time - Example 2



S : Start

H : Last flip = H

T : Last flip = T

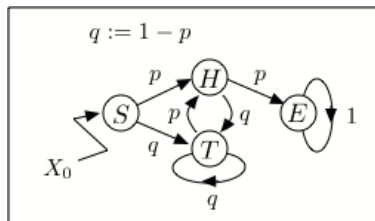
E : Done

Let us justify the first step equation for $\beta(T)$. The others are similar.

$N(T)$ – number of steps, starting from T until the MC hits E .

$N(H)$ – be defined similarly.

Hitting Time - Example 2



S : Start

H : Last flip = H

T : Last flip = T

E : Done

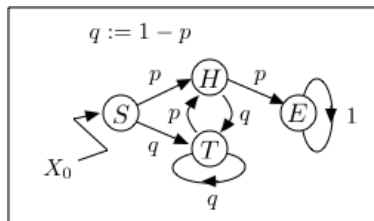
Let us justify the first step equation for $\beta(T)$. The others are similar.

$N(T)$ – number of steps, starting from T until the MC hits E .

$N(H)$ – be defined similarly.

$N'(T)$ – number of steps after the second visit to T until MC hits E .

Hitting Time - Example 2



S : Start

H : Last flip = H

T : Last flip = T

E : Done

Let us justify the first step equation for $\beta(T)$. The others are similar.

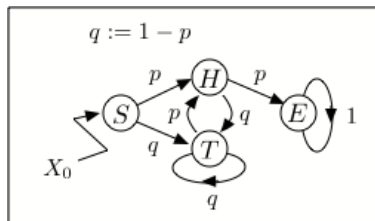
$N(T)$ – number of steps, starting from T until the MC hits E .

$N(H)$ – be defined similarly.

$N'(T)$ – number of steps after the second visit to T until MC hits E .

$$N(T) = 1 + Z \times N(H) + (1 - Z) \times N'(T)$$

Hitting Time - Example 2



S : Start

H : Last flip = H

T : Last flip = T

E : Done

Let us justify the first step equation for $\beta(T)$. The others are similar.

$N(T)$ – number of steps, starting from T until the MC hits E .

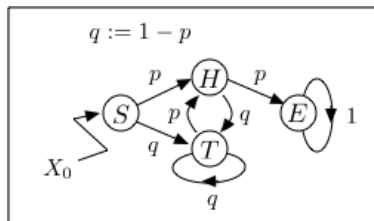
$N(H)$ – be defined similarly.

$N'(T)$ – number of steps after the second visit to T until MC hits E .

$$N(T) = 1 + Z \times N(H) + (1 - Z) \times N'(T)$$

where $Z = 1_{\{\text{first flip in } T \text{ is } H\}}$.

Hitting Time - Example 2



S : Start

H : Last flip = H

T : Last flip = T

E : Done

Let us justify the first step equation for $\beta(T)$. The others are similar.

$N(T)$ – number of steps, starting from T until the MC hits E .

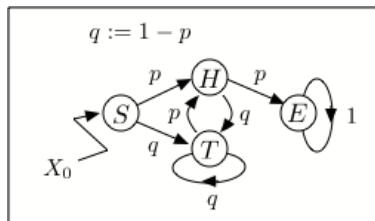
$N(H)$ – be defined similarly.

$N'(T)$ – number of steps after the second visit to T until MC hits E .

$$N(T) = 1 + Z \times N(H) + (1 - Z) \times N'(T)$$

where $Z = 1\{\text{first flip in } T \text{ is } H\}$. Since Z and $N(H)$ are independent,

Hitting Time - Example 2



S : Start

H : Last flip = H

T : Last flip = T

E : Done

Let us justify the first step equation for $\beta(T)$. The others are similar.

$N(T)$ – number of steps, starting from T until the MC hits E .

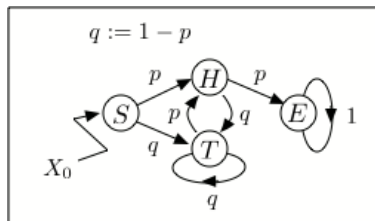
$N(H)$ – be defined similarly.

$N'(T)$ – number of steps after the second visit to T until MC hits E .

$$N(T) = 1 + Z \times N(H) + (1 - Z) \times N'(T)$$

where $Z = 1_{\{\text{first flip in } T \text{ is } H\}}$. Since Z and $N(H)$ are independent, and Z and $N'(T)$ are independent,

Hitting Time - Example 2



S : Start

H : Last flip = H

T : Last flip = T

E : Done

Let us justify the first step equation for $\beta(T)$. The others are similar.

$N(T)$ – number of steps, starting from T until the MC hits E .

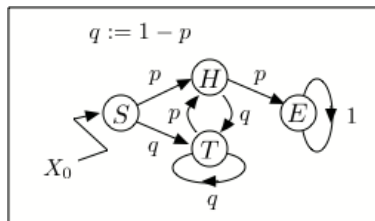
$N(H)$ – be defined similarly.

$N'(T)$ – number of steps after the second visit to T until MC hits E .

$$N(T) = 1 + Z \times N(H) + (1 - Z) \times N'(T)$$

where $Z = 1_{\{\text{first flip in } T \text{ is } H\}}$. Since Z and $N(H)$ are independent, and Z and $N'(T)$ are independent, taking expectations, we get

Hitting Time - Example 2



S : Start

H : Last flip = H

T : Last flip = T

E : Done

Let us justify the first step equation for $\beta(T)$. The others are similar.

$N(T)$ – number of steps, starting from T until the MC hits E .

$N(H)$ – be defined similarly.

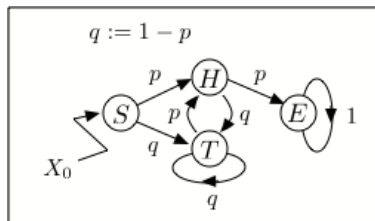
$N'(T)$ – number of steps after the second visit to T until MC hits E .

$$N(T) = 1 + Z \times N(H) + (1 - Z) \times N'(T)$$

where $Z = 1_{\{\text{first flip in } T \text{ is } H\}}$. Since Z and $N(H)$ are independent, and Z and $N'(T)$ are independent, taking expectations, we get

$$E[N(T)] = 1 + pE[N(H)] + qE[N'(T)],$$

Hitting Time - Example 2



S : Start

H : Last flip = H

T : Last flip = T

E : Done

Let us justify the first step equation for $\beta(T)$. The others are similar.

$N(T)$ – number of steps, starting from T until the MC hits E .

$N(H)$ – be defined similarly.

$N'(T)$ – number of steps after the second visit to T until MC hits E .

$$N(T) = 1 + Z \times N(H) + (1 - Z) \times N'(T)$$

where $Z = 1\{\text{first flip in } T \text{ is } H\}$. Since Z and $N(H)$ are independent, and Z and $N'(T)$ are independent, taking expectations, we get

$$E[N(T)] = 1 + pE[N(H)] + qE[N'(T)],$$

i.e.,

$$\beta(T) = 1 + p\beta(H) + q\beta(T).$$

Hitting Time - Example 3

Hitting Time - Example 3

You roll a balanced six-sided die until the sum of the last two rolls is 8.

Hitting Time - Example 3

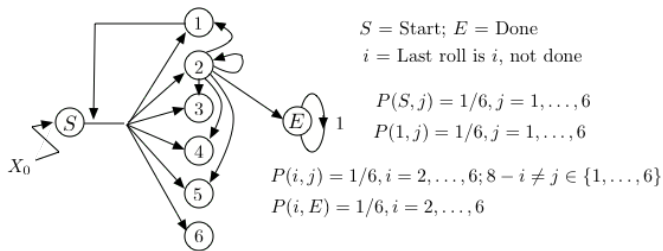
You roll a balanced six-sided die until the sum of the last two rolls is 8.
How many times do you have to roll the die,

Hitting Time - Example 3

You roll a balanced six-sided die until the sum of the last two rolls is 8. How many times do you have to roll the die, on average?

Hitting Time - Example 3

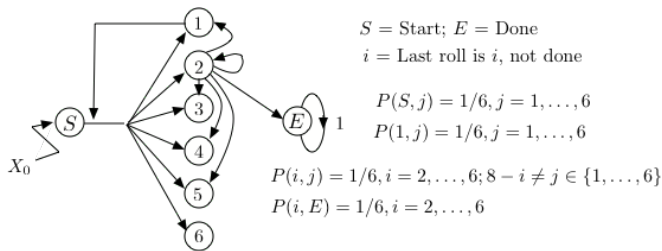
You roll a balanced six-sided die until the sum of the last two rolls is 8.
How many times do you have to roll the die, on average?



The arrows out of $3, \dots, 6$ (not shown) are similar to those out of 2.

Hitting Time - Example 3

You roll a balanced six-sided die until the sum of the last two rolls is 8.
How many times do you have to roll the die, on average?

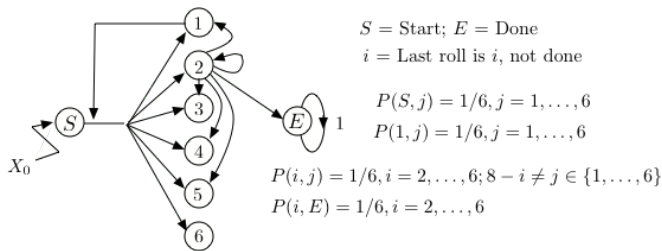


The arrows out of 3, ..., 6 (not shown) are similar to those out of 2.

$$\beta(S) = 1 + \frac{1}{6} \sum_{j=1}^6 \beta(j);$$

Hitting Time - Example 3

You roll a balanced six-sided die until the sum of the last two rolls is 8. How many times do you have to roll the die, on average?

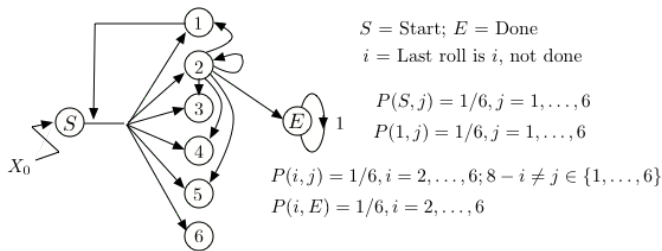


The arrows out of 3, ..., 6 (not shown) are similar to those out of 2.

$$\beta(S) = 1 + \frac{1}{6} \sum_{j=1}^6 \beta(j); \beta(1) = 1 + \frac{1}{6} \sum_{j=1}^6 \beta(j);$$

Hitting Time - Example 3

You roll a balanced six-sided die until the sum of the last two rolls is 8.
How many times do you have to roll the die, on average?

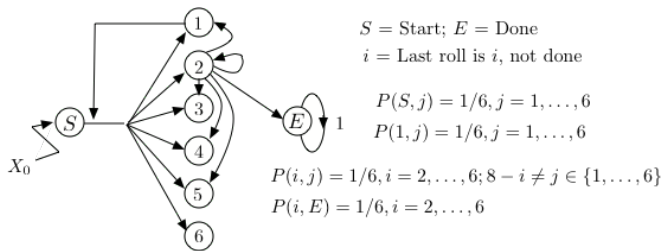


The arrows out of $3, \dots, 6$ (not shown) are similar to those out of 2.

$$\beta(S) = 1 + \frac{1}{6} \sum_{j=1}^6 \beta(j); \beta(1) = 1 + \frac{1}{6} \sum_{j=1}^6 \beta(j); \beta(i) = 1 + \frac{1}{6} \sum_{j=1, \dots, 6; j \neq 8-i} \beta(j), i = 2, \dots, 6.$$

Hitting Time - Example 3

You roll a balanced six-sided die until the sum of the last two rolls is 8.
How many times do you have to roll the die, on average?



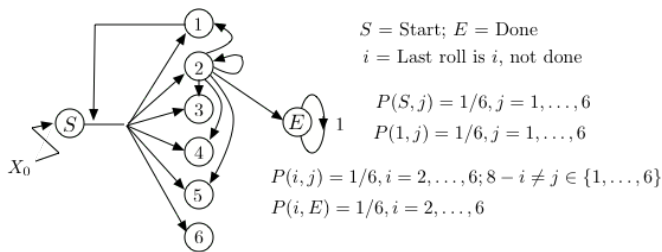
The arrows out of $3, \dots, 6$ (not shown) are similar to those out of 2.

$$\beta(S) = 1 + \frac{1}{6} \sum_{j=1}^6 \beta(j); \beta(1) = 1 + \frac{1}{6} \sum_{j=1}^6 \beta(j); \beta(i) = 1 + \frac{1}{6} \sum_{j=1, \dots, 6; j \neq 8-i} \beta(j), i = 2, \dots, 6.$$

Symmetry: $\beta(2) = \dots = \beta(6) =: \gamma$.

Hitting Time - Example 3

You roll a balanced six-sided die until the sum of the last two rolls is 8. How many times do you have to roll the die, on average?



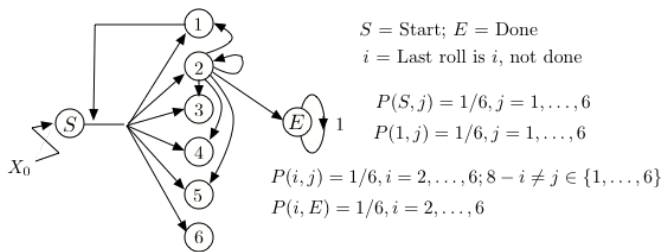
The arrows out of $3, \dots, 6$ (not shown) are similar to those out of 2.

$$\beta(S) = 1 + \frac{1}{6} \sum_{j=1}^6 \beta(j); \beta(1) = 1 + \frac{1}{6} \sum_{j=1}^6 \beta(j); \beta(i) = 1 + \frac{1}{6} \sum_{j=1, \dots, 6; j \neq 8-i} \beta(j), i = 2, \dots, 6.$$

Symmetry: $\beta(2) = \dots = \beta(6) =: \gamma$. Also, $\beta(1) = \beta(S)$.

Hitting Time - Example 3

You roll a balanced six-sided die until the sum of the last two rolls is 8.
How many times do you have to roll the die, on average?



The arrows out of 3, ..., 6 (not shown) are similar to those out of 2.

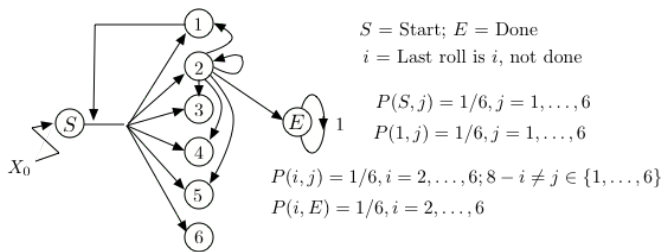
$$\beta(S) = 1 + \frac{1}{6} \sum_{j=1}^6 \beta(j); \beta(1) = 1 + \frac{1}{6} \sum_{j=1}^6 \beta(j); \beta(i) = 1 + \frac{1}{6} \sum_{j=1, \dots, 6; j \neq 8-i} \beta(j), i = 2, \dots, 6.$$

Symmetry: $\beta(2) = \dots = \beta(6) =: \gamma$. Also, $\beta(1) = \beta(S)$. Thus,

$$\beta(S) = 1 + (5/6)\gamma + \beta(S)/6;$$

Hitting Time - Example 3

You roll a balanced six-sided die until the sum of the last two rolls is 8.
How many times do you have to roll the die, on average?



The arrows out of 3, ..., 6 (not shown) are similar to those out of 2.

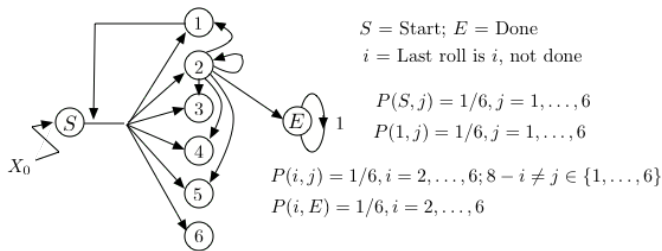
$$\beta(S) = 1 + \frac{1}{6} \sum_{j=1}^6 \beta(j); \beta(1) = 1 + \frac{1}{6} \sum_{j=1}^6 \beta(j); \beta(i) = 1 + \frac{1}{6} \sum_{j=1, \dots, 6; j \neq 8-i} \beta(j), i = 2, \dots, 6.$$

Symmetry: $\beta(2) = \dots = \beta(6) =: \gamma$. Also, $\beta(1) = \beta(S)$. Thus,

$$\beta(S) = 1 + (5/6)\gamma + \beta(S)/6; \quad \gamma = 1 + (4/6)\gamma + (1/6)\beta(S).$$

Hitting Time - Example 3

You roll a balanced six-sided die until the sum of the last two rolls is 8.
How many times do you have to roll the die, on average?



The arrows out of 3, ..., 6 (not shown) are similar to those out of 2.

$$\beta(S) = 1 + \frac{1}{6} \sum_{j=1}^6 \beta(j); \beta(1) = 1 + \frac{1}{6} \sum_{j=1}^6 \beta(j); \beta(i) = 1 + \frac{1}{6} \sum_{j=1, \dots, 6; j \neq 8-i} \beta(j), i = 2, \dots, 6.$$

Symmetry: $\beta(2) = \dots = \beta(6) =: \gamma$. Also, $\beta(1) = \beta(S)$. Thus,

$$\beta(S) = 1 + (5/6)\gamma + \beta(S)/6; \quad \gamma = 1 + (4/6)\gamma + (1/6)\beta(S).$$

$$\Rightarrow \dots \beta(S) = 8.4.$$

Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.

Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.

Start with \$10.

Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1.

Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

What is the probability that you reach \$100 before \$0?

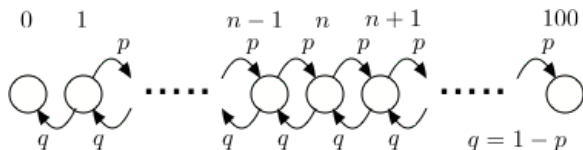
Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

What is the probability that you reach \$100 before \$0?

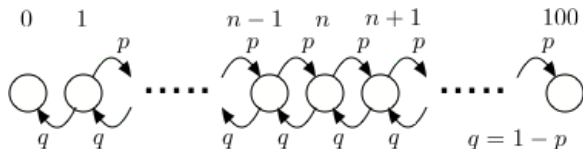


Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.
Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

What is the probability that you reach \$100 before \$0?



Let $\alpha(n)$ be the probability of reaching 100 before 0, starting from n , for $n = 0, 1, \dots, 100$.

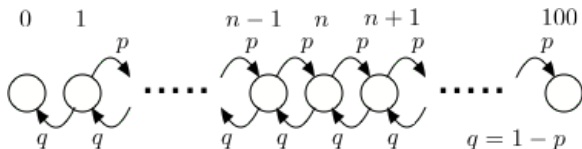
Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

What is the probability that you reach \$100 before \$0?



Let $\alpha(n)$ be the probability of reaching 100 before 0, starting from n , for $n = 0, 1, \dots, 100$.

Which equations are correct?

- (A) $\alpha(0) = 0$
- (B) $\alpha(0) = 1$.
- (C) $\alpha(100) =$

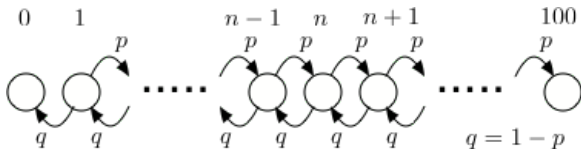
Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

What is the probability that you reach \$100 before \$0?



Let $\alpha(n)$ be the probability of reaching 100 before 0, starting from n , for $n = 0, 1, \dots, 100$.

Which equations are correct?

(A) $\alpha(0) = 0$

(B) $\alpha(0) = 1$.

(C) $\alpha(100) = 1$.

(D) $\alpha(n) = 1 + p\alpha(n+1) + q\alpha(n-1), 0 < n < 100$.

(E) $\alpha(n) =$

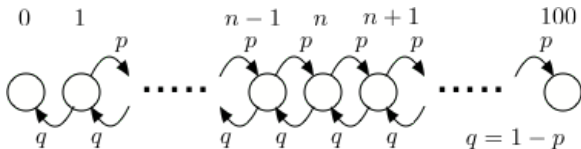
Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

What is the probability that you reach \$100 before \$0?



Let $\alpha(n)$ be the probability of reaching 100 before 0, starting from n , for $n = 0, 1, \dots, 100$.

Which equations are correct?

(A) $\alpha(0) = 0$

(B) $\alpha(0) = 1$.

(C) $\alpha(100) = 1$.

(D) $\alpha(n) = 1 + p\alpha(n+1) + q\alpha(n-1), 0 < n < 100$.

(E) $\alpha(n) = p\alpha(n+1) + q\alpha(n-1), 0 < n < 100$.

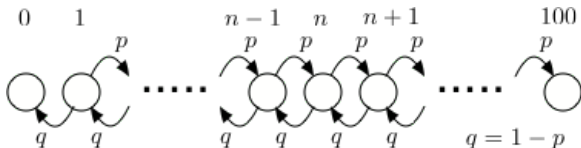
Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

What is the probability that you reach \$100 before \$0?



Let $\alpha(n)$ be the probability of reaching 100 before 0, starting from n , for $n = 0, 1, \dots, 100$.

Which equations are correct?

(A) $\alpha(0) = 0$

(B) $\alpha(0) = 1$.

(C) $\alpha(100) = 1$.

(D) $\alpha(n) = 1 + p\alpha(n+1) + q\alpha(n-1), 0 < n < 100$.

(E) $\alpha(n) = p\alpha(n+1) + q\alpha(n-1), 0 < n < 100$.

(B) is incorrect, 0 is bad.

(D) is incorrect. Confuses expected hitting time with A before B.

Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.

Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.

Start with \$10.

Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1.

Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

What is the probability that you reach \$100 before \$0?

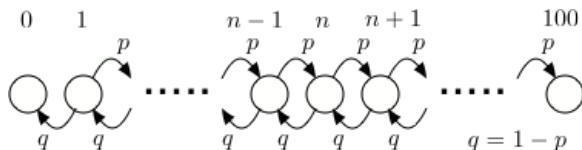
Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

What is the probability that you reach \$100 before \$0?



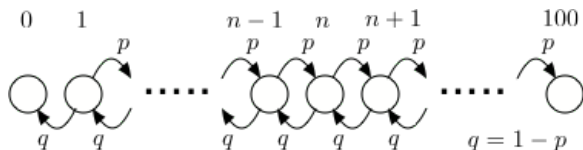
Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

What is the probability that you reach \$100 before \$0?



Let $\alpha(n)$ be the probability of reaching 100 before 0, starting from n , for $n = 0, 1, \dots, 100$.

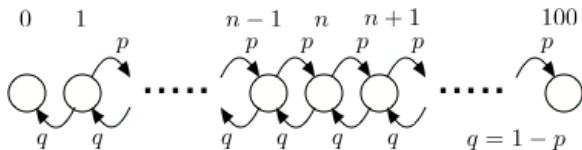
Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

What is the probability that you reach \$100 before \$0?



Let $\alpha(n)$ be the probability of reaching 100 before 0, starting from n , for $n = 0, 1, \dots, 100$.

$$\alpha(0) =$$

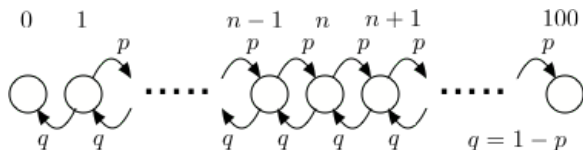
Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

What is the probability that you reach \$100 before \$0?



Let $\alpha(n)$ be the probability of reaching 100 before 0, starting from n , for $n = 0, 1, \dots, 100$.

$$\alpha(0) = 0;$$

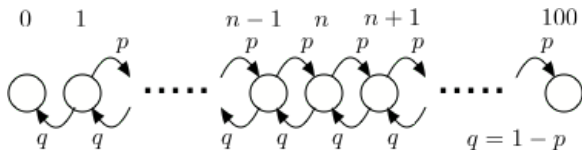
Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

What is the probability that you reach \$100 before \$0?



Let $\alpha(n)$ be the probability of reaching 100 before 0, starting from n , for $n = 0, 1, \dots, 100$.

$$\alpha(0) = 0; \alpha(100) =$$

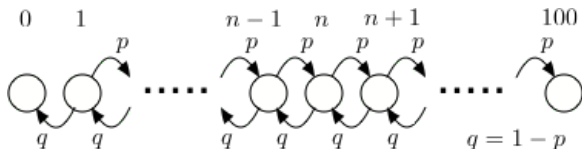
Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

What is the probability that you reach \$100 before \$0?



Let $\alpha(n)$ be the probability of reaching 100 before 0, starting from n , for $n = 0, 1, \dots, 100$.

$$\alpha(0) = 0; \alpha(100) = 1.$$

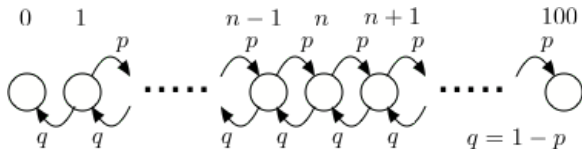
Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

What is the probability that you reach \$100 before \$0?



Let $\alpha(n)$ be the probability of reaching 100 before 0, starting from n , for $n = 0, 1, \dots, 100$.

$$\alpha(0) = 0; \alpha(100) = 1.$$

$$\alpha(n) =$$

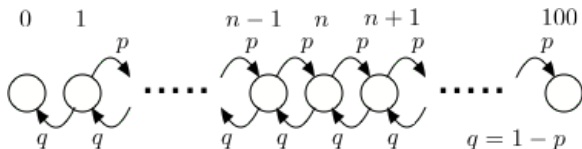
Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

What is the probability that you reach \$100 before \$0?



Let $\alpha(n)$ be the probability of reaching 100 before 0, starting from n , for $n = 0, 1, \dots, 100$.

$$\alpha(0) = 0; \alpha(100) = 1.$$

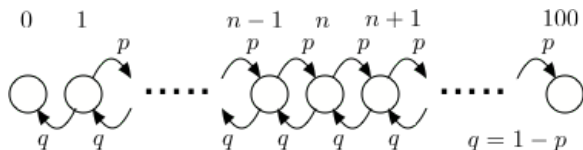
$$\alpha(n) = p\alpha(n+1) + q\alpha(n-1), 0 < n < 100.$$

Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.
Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

What is the probability that you reach \$100 before \$0?



Let $\alpha(n)$ be the probability of reaching 100 before 0, starting from n , for $n = 0, 1, \dots, 100$.

$$\alpha(0) = 0; \alpha(100) = 1.$$

$$\alpha(n) = p\alpha(n+1) + q\alpha(n-1), 0 < n < 100.$$

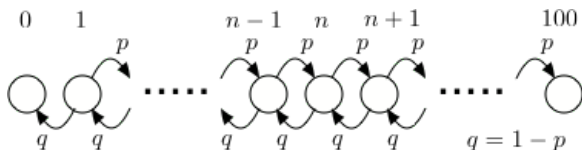
$$\Rightarrow \alpha(n) = \frac{1 - \rho^n}{1 - \rho^{100}} \text{ with } \rho = qp^{-1}.$$

Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p < 0.5$.
Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

What is the probability that you reach \$100 before \$0?



Let $\alpha(n)$ be the probability of reaching 100 before 0, starting from n , for $n = 0, 1, \dots, 100$.

$$\alpha(0) = 0; \alpha(100) = 1.$$

$$\alpha(n) = p\alpha(n+1) + q\alpha(n-1), 0 < n < 100.$$

$$\Rightarrow \alpha(n) = \frac{1 - \rho^n}{1 - \rho^{100}} \text{ with } \rho = qp^{-1}. \text{ (See LN 22)}$$

Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p = .48$.

Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p = .48$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1.

Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p = .48$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p = .48$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

What is the probability that you reach \$100 before \$0?

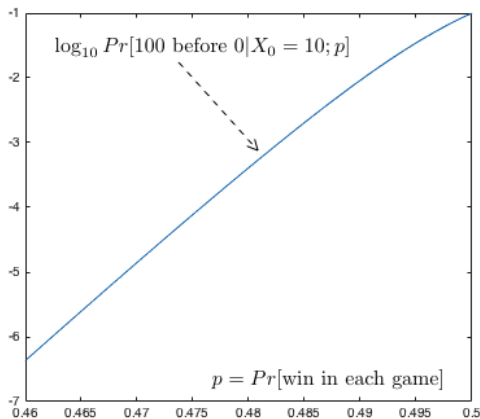
Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p = .48$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

What is the probability that you reach \$100 before \$0?



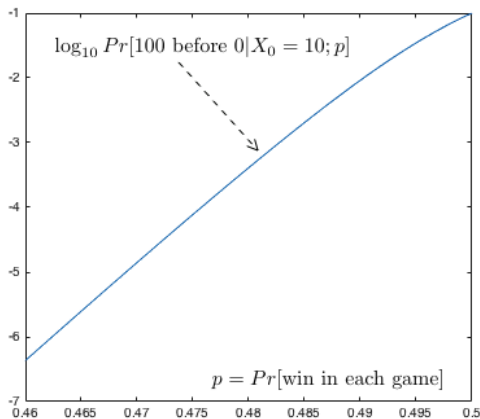
Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p = .48$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

What is the probability that you reach \$100 before \$0?



Less than 1 in a 1000.

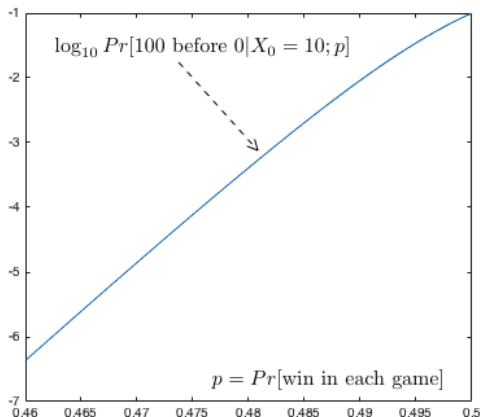
Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p = .48$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

What is the probability that you reach \$100 before \$0?



Less than 1 in a 1000. Moral of example:

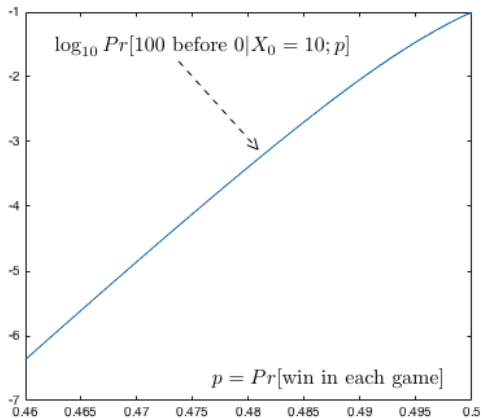
Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p = .48$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

What is the probability that you reach \$100 before \$0?



Less than 1 in a 1000. Moral of example: Money in Vegas

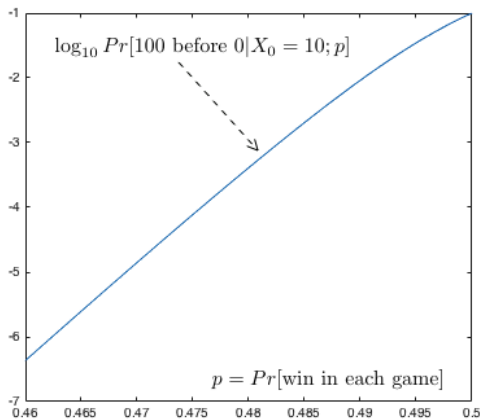
Here before There - A before B

Game of “heads or tails” using coin with ‘heads’ probability $p = .48$.

Start with \$10.

Each step, flip yields ‘heads’, earn \$1. Otherwise, lose \$1.

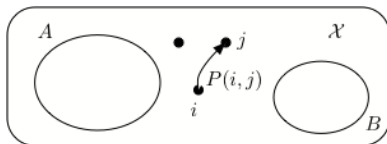
What is the probability that you reach \$100 before \$0?



Less than 1 in a 1000. Moral of example: Money in Vegas stays in Vegas.

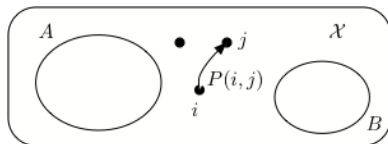
First Step Equations

First Step Equations



Let X_n be a MC on \mathcal{X} and $A, B \subset \mathcal{X}$ with $A \cap B = \emptyset$.

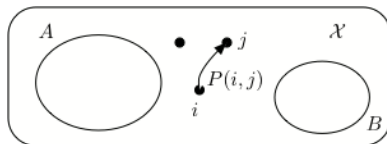
First Step Equations



Let X_n be a MC on \mathcal{X} and $A, B \subset \mathcal{X}$ with $A \cap B = \emptyset$. Define

$$T_A = \min\{n \geq 0 \mid X_n \in A\}$$

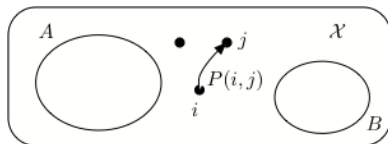
First Step Equations



Let X_n be a MC on \mathcal{X} and $A, B \subset \mathcal{X}$ with $A \cap B = \emptyset$. Define

$$T_A = \min\{n \geq 0 \mid X_n \in A\} \text{ and } T_B = \min\{n \geq 0 \mid X_n \in B\}.$$

First Step Equations

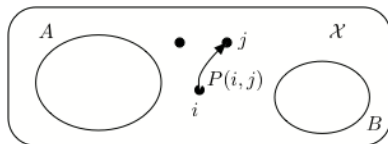


Let X_n be a MC on \mathcal{X} and $A, B \subset \mathcal{X}$ with $A \cap B = \emptyset$. Define

$$T_A = \min\{n \geq 0 \mid X_n \in A\} \text{ and } T_B = \min\{n \geq 0 \mid X_n \in B\}.$$

For $\beta(i) = E[T_A \mid X_0 = i]$, first step equations are:

First Step Equations



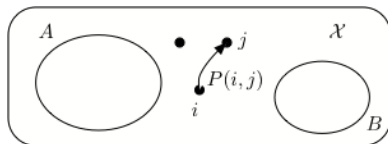
Let X_n be a MC on \mathcal{X} and $A, B \subset \mathcal{X}$ with $A \cap B = \emptyset$. Define

$$T_A = \min\{n \geq 0 \mid X_n \in A\} \text{ and } T_B = \min\{n \geq 0 \mid X_n \in B\}.$$

For $\beta(i) = E[T_A \mid X_0 = i]$, first step equations are:

$$\beta(i) = 0, i \in A$$

First Step Equations



Let X_n be a MC on \mathcal{X} and $A, B \subset \mathcal{X}$ with $A \cap B = \emptyset$. Define

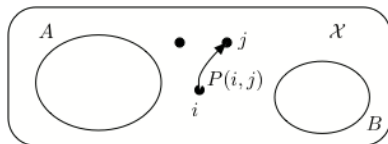
$$T_A = \min\{n \geq 0 \mid X_n \in A\} \text{ and } T_B = \min\{n \geq 0 \mid X_n \in B\}.$$

For $\beta(i) = E[T_A \mid X_0 = i]$, first step equations are:

$$\beta(i) = 0, i \in A$$

$$\beta(i) = 1 + \sum_j P(i, j)\beta(j), i \notin A$$

First Step Equations



Let X_n be a MC on \mathcal{X} and $A, B \subset \mathcal{X}$ with $A \cap B = \emptyset$. Define

$$T_A = \min\{n \geq 0 \mid X_n \in A\} \text{ and } T_B = \min\{n \geq 0 \mid X_n \in B\}.$$

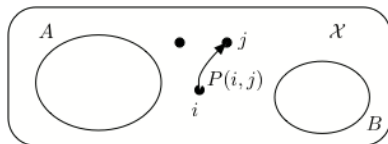
For $\beta(i) = E[T_A \mid X_0 = i]$, first step equations are:

$$\beta(i) = 0, i \in A$$

$$\beta(i) = 1 + \sum_j P(i, j) \beta(j), i \notin A$$

For $\alpha(i) = Pr[T_A < T_B \mid X_0 = i], i \in \mathcal{X}$, first step equations are:

First Step Equations



Let X_n be a MC on \mathcal{X} and $A, B \subset \mathcal{X}$ with $A \cap B = \emptyset$. Define

$$T_A = \min\{n \geq 0 \mid X_n \in A\} \text{ and } T_B = \min\{n \geq 0 \mid X_n \in B\}.$$

For $\beta(i) = E[T_A \mid X_0 = i]$, first step equations are:

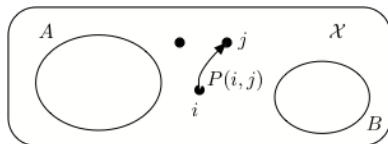
$$\beta(i) = 0, i \in A$$

$$\beta(i) = 1 + \sum_j P(i, j) \beta(j), i \notin A$$

For $\alpha(i) = Pr[T_A < T_B \mid X_0 = i], i \in \mathcal{X}$, first step equations are:

$$\alpha(i) = 1, i \in A$$

First Step Equations



Let X_n be a MC on \mathcal{X} and $A, B \subset \mathcal{X}$ with $A \cap B = \emptyset$. Define

$$T_A = \min\{n \geq 0 \mid X_n \in A\} \text{ and } T_B = \min\{n \geq 0 \mid X_n \in B\}.$$

For $\beta(i) = E[T_A \mid X_0 = i]$, first step equations are:

$$\beta(i) = 0, i \in A$$

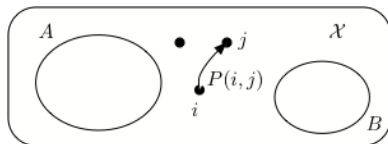
$$\beta(i) = 1 + \sum_j P(i, j) \beta(j), i \notin A$$

For $\alpha(i) = Pr[T_A < T_B \mid X_0 = i], i \in \mathcal{X}$, first step equations are:

$$\alpha(i) = 1, i \in A$$

$$\alpha(i) = 0, i \in B$$

First Step Equations



Let X_n be a MC on \mathcal{X} and $A, B \subset \mathcal{X}$ with $A \cap B = \emptyset$. Define

$$T_A = \min\{n \geq 0 \mid X_n \in A\} \text{ and } T_B = \min\{n \geq 0 \mid X_n \in B\}.$$

For $\beta(i) = E[T_A \mid X_0 = i]$, first step equations are:

$$\beta(i) = 0, i \in A$$

$$\beta(i) = 1 + \sum_j P(i, j) \beta(j), i \notin A$$

For $\alpha(i) = Pr[T_A < T_B \mid X_0 = i], i \in \mathcal{X}$, first step equations are:

$$\alpha(i) = 1, i \in A$$

$$\alpha(i) = 0, i \in B$$

$$\alpha(i) = \sum_j P(i, j) \alpha(j), i \notin A \cup B.$$

Accumulating Rewards

Accumulating Rewards

Let X_n be a Markov chain on \mathcal{X} with P .

Accumulating Rewards

Let X_n be a Markov chain on \mathcal{X} with P . Let $A \subset \mathcal{X}$

Accumulating Rewards

Let X_n be a Markov chain on \mathcal{X} with P . Let $A \subset \mathcal{X}$

Let also $g : \mathcal{X} \rightarrow \Re$ be some function.

Accumulating Rewards

Let X_n be a Markov chain on \mathcal{X} with P . Let $A \subset \mathcal{X}$

Let also $g : \mathcal{X} \rightarrow \Re$ be some function.

Define

$$\gamma(i) = E\left[\sum_{n=0}^{T_A} g(X_n) \mid X_0 = i\right], i \in \mathcal{X}.$$

Accumulating Rewards

Let X_n be a Markov chain on \mathcal{X} with P . Let $A \subset \mathcal{X}$

Let also $g : \mathcal{X} \rightarrow \Re$ be some function.

Define

$$\gamma(i) = E\left[\sum_{n=0}^{T_A} g(X_n) \mid X_0 = i\right], i \in \mathcal{X}.$$

Then

$$\gamma(i) = \begin{cases} g(i), & \text{if } i \in A \end{cases}$$

Accumulating Rewards

Let X_n be a Markov chain on \mathcal{X} with P . Let $A \subset \mathcal{X}$

Let also $g : \mathcal{X} \rightarrow \Re$ be some function.

Define

$$\gamma(i) = E\left[\sum_{n=0}^{T_A} g(X_n) \mid X_0 = i\right], i \in \mathcal{X}.$$

Then

$$\gamma(i) = \begin{cases} g(i), & \text{if } i \in A \\ g(i) + \sum_j P(i, j)\gamma(j), & \text{otherwise.} \end{cases}$$

Accumulating Rewards

Let X_n be a Markov chain on \mathcal{X} with P . Let $A \subset \mathcal{X}$

Let also $g : \mathcal{X} \rightarrow \Re$ be some function.

Define

$$\gamma(i) = E\left[\sum_{n=0}^{T_A} g(X_n) \mid X_0 = i\right], i \in \mathcal{X}.$$

Then

$$\gamma(i) = \begin{cases} g(i), & \text{if } i \in A \\ g(i) + \sum_j P(i, j)\gamma(j), & \text{otherwise.} \end{cases}$$

Example

Example

Flip a fair coin until you get two consecutive H s.

Example

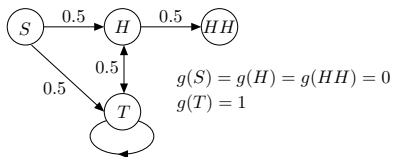
Flip a fair coin until you get two consecutive H s.

What is the expected number of T s that you see?

Example

Flip a fair coin until you get two consecutive H s.

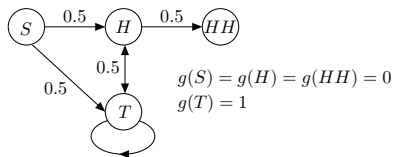
What is the expected number of T s that you see?



Example

Flip a fair coin until you get two consecutive H s.

What is the expected number of T s that you see?



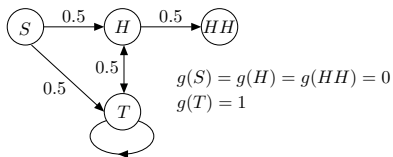
FSE:

$$\gamma(S) = 0 + 0.5\gamma(H) + 0.5\gamma(T)$$

Example

Flip a fair coin until you get two consecutive H s.

What is the expected number of T s that you see?



FSE:

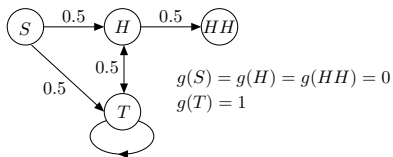
$$\gamma(S) = 0 + 0.5\gamma(H) + 0.5\gamma(T)$$

$$\gamma(H) = 0 + 0.5\gamma(HH) + 0.5\gamma(T)$$

Example

Flip a fair coin until you get two consecutive H s.

What is the expected number of T s that you see?



FSE:

$$\gamma(S) = 0 + 0.5\gamma(H) + 0.5\gamma(T)$$

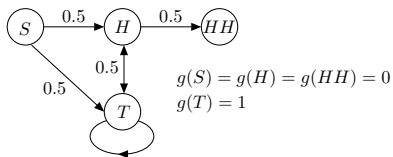
$$\gamma(H) = 0 + 0.5\gamma(HH) + 0.5\gamma(T)$$

$$\gamma(T) = 1 + 0.5\gamma(H) + 0.5\gamma(T)$$

Example

Flip a fair coin until you get two consecutive H s.

What is the expected number of T s that you see?



FSE:

$$\gamma(S) = 0 + 0.5\gamma(H) + 0.5\gamma(T)$$

$$\gamma(H) = 0 + 0.5\gamma(HH) + 0.5\gamma(T)$$

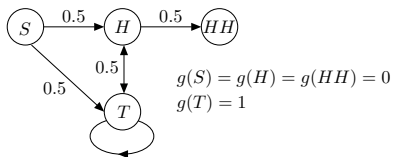
$$\gamma(T) = 1 + 0.5\gamma(H) + 0.5\gamma(T)$$

$$\gamma(HH) = 0.$$

Example

Flip a fair coin until you get two consecutive H s.

What is the expected number of T s that you see?



FSE:

$$\gamma(S) = 0 + 0.5\gamma(H) + 0.5\gamma(T)$$

$$\gamma(H) = 0 + 0.5\gamma(HH) + 0.5\gamma(T)$$

$$\gamma(T) = 1 + 0.5\gamma(H) + 0.5\gamma(T)$$

$$\gamma(HH) = 0.$$

Solving, we find $\gamma(S) = 2.5$.

Recap

Recap

- ▶ Markov Chain:

Recap

- ▶ Markov Chain:

- ▶ Finite set \mathcal{X} ; π_0 ; $P = \{P(i,j), i,j \in \mathcal{X}\}$;

- ▶ $Pr[X_0 = i] = \pi_0(i), i \in \mathcal{X}$

- ▶ $Pr[X_{n+1} = j | X_0, \dots, X_n = i] = P(i,j), i,j \in \mathcal{X}, n \geq 0.$

- ▶ Note:

- $Pr[X_0 = i_0, X_1 = i_1, \dots, X_n = i_n] =$

Recap

- ▶ Markov Chain:

- ▶ Finite set \mathcal{X} ; π_0 ; $P = \{P(i,j), i,j \in \mathcal{X}\}$;

- ▶ $Pr[X_0 = i] = \pi_0(i), i \in \mathcal{X}$

- ▶ $Pr[X_{n+1} = j | X_0, \dots, X_n = i] = P(i,j), i,j \in \mathcal{X}, n \geq 0.$

- ▶ Note:

- $$Pr[X_0 = i_0, X_1 = i_1, \dots, X_n = i_n] = \pi_0(i_0)P(i_0, i_1) \cdots P(i_{n-1}, i_n).$$

- ▶ First Passage Time:

Recap

- ▶ Markov Chain:

- ▶ Finite set \mathcal{X} ; π_0 ; $P = \{P(i,j), i,j \in \mathcal{X}\}$;

- ▶ $Pr[X_0 = i] = \pi_0(i), i \in \mathcal{X}$

- ▶ $Pr[X_{n+1} = j | X_0, \dots, X_n = i] = P(i,j), i,j \in \mathcal{X}, n \geq 0.$

- ▶ Note:

- $$Pr[X_0 = i_0, X_1 = i_1, \dots, X_n = i_n] = \pi_0(i_0)P(i_0, i_1) \cdots P(i_{n-1}, i_n).$$

- ▶ First Passage Time:

- ▶ $A \cap B = \emptyset$;

Recap

- ▶ Markov Chain:

- ▶ Finite set \mathcal{X} ; π_0 ; $P = \{P(i,j), i,j \in \mathcal{X}\}$;

- ▶ $Pr[X_0 = i] = \pi_0(i), i \in \mathcal{X}$

- ▶ $Pr[X_{n+1} = j | X_0, \dots, X_n = i] = P(i,j), i,j \in \mathcal{X}, n \geq 0.$

- ▶ Note:

- $$Pr[X_0 = i_0, X_1 = i_1, \dots, X_n = i_n] = \pi_0(i_0)P(i_0, i_1) \cdots P(i_{n-1}, i_n).$$

- ▶ First Passage Time:

- ▶ $A \cap B = \emptyset; \beta(i) = E[T_A | X_0 = i];$

Recap

▶ Markov Chain:

▶ Finite set \mathcal{X} ; π_0 ; $P = \{P(i,j), i,j \in \mathcal{X}\}$;

▶ $Pr[X_0 = i] = \pi_0(i), i \in \mathcal{X}$

▶ $Pr[X_{n+1} = j | X_0, \dots, X_n = i] = P(i,j), i,j \in \mathcal{X}, n \geq 0.$

▶ Note:

$$Pr[X_0 = i_0, X_1 = i_1, \dots, X_n = i_n] = \pi_0(i_0)P(i_0, i_1) \cdots P(i_{n-1}, i_n).$$

▶ First Passage Time:

▶ $A \cap B = \emptyset; \beta(i) = E[T_A | X_0 = i]; \alpha(i) = P[T_A < T_B | X_0 = i]$

▶ $\beta(i) = 1 + \sum_j P(i,j)\beta(j);$

▶ $\alpha(i) = \sum_j P(i,j)\alpha(j). \alpha(A) = 1, \alpha(B) = 0.$

Summary

Markov Chains

Summary

Markov Chains

Summary

Markov Chains

- ▶ Markov Chain:

Summary

Markov Chains

- ▶ Markov Chain: $Pr[X_{n+1} = j | X_0, \dots, X_n = i] = P(i, j)$

Summary

Markov Chains

- ▶ Markov Chain: $Pr[X_{n+1} = j | X_0, \dots, X_n = i] = P(i, j)$
- ▶ FSE: $\beta(i) = 1 + \sum_j P(i, j)\beta(j)$;

Summary

Markov Chains

- ▶ Markov Chain: $Pr[X_{n+1} = j | X_0, \dots, X_n = i] = P(i, j)$
- ▶ FSE: $\beta(i) = 1 + \sum_j P(i, j)\beta(j)$; $\alpha(i) = \sum_j P(i, j)\alpha(j)$.

Summary

Markov Chains

- ▶ Markov Chain: $Pr[X_{n+1} = j | X_0, \dots, X_n = i] = P(i, j)$
- ▶ FSE: $\beta(i) = 1 + \sum_j P(i, j)\beta(j)$; $\alpha(i) = \sum_j P(i, j)\alpha(j)$.
- ▶ $\pi_n = \pi_0 P^n$

Summary

Markov Chains

- ▶ Markov Chain: $Pr[X_{n+1} = j | X_0, \dots, X_n = i] = P(i, j)$
- ▶ FSE: $\beta(i) = 1 + \sum_j P(i, j)\beta(j)$; $\alpha(i) = \sum_j P(i, j)\alpha(j)$.
- ▶ $\pi_n = \pi_0 P^n$
- ▶ π is invariant iff $\pi P = \pi$

Summary

Markov Chains

- ▶ Markov Chain: $Pr[X_{n+1} = j | X_0, \dots, X_n = i] = P(i, j)$
- ▶ FSE: $\beta(i) = 1 + \sum_j P(i, j)\beta(j)$; $\alpha(i) = \sum_j P(i, j)\alpha(j)$.
- ▶ $\pi_n = \pi_0 P^n$
- ▶ π is invariant iff $\pi P = \pi$
- ▶ Irreducible \Rightarrow one and only one invariant distribution π

Summary

Markov Chains

- ▶ Markov Chain: $Pr[X_{n+1} = j | X_0, \dots, X_n = i] = P(i, j)$
- ▶ FSE: $\beta(i) = 1 + \sum_j P(i, j)\beta(j)$; $\alpha(i) = \sum_j P(i, j)\alpha(j)$.
- ▶ $\pi_n = \pi_0 P^n$
- ▶ π is invariant iff $\pi P = \pi$
- ▶ Irreducible \Rightarrow one and only one invariant distribution π
- ▶ Irreducible \Rightarrow fraction of time in state i approaches $\pi(i)$

Summary

Markov Chains

- ▶ Markov Chain: $Pr[X_{n+1} = j | X_0, \dots, X_n = i] = P(i, j)$
- ▶ FSE: $\beta(i) = 1 + \sum_j P(i, j)\beta(j)$; $\alpha(i) = \sum_j P(i, j)\alpha(j)$.
- ▶ $\pi_n = \pi_0 P^n$
- ▶ π is invariant iff $\pi P = \pi$
- ▶ Irreducible \Rightarrow one and only one invariant distribution π
- ▶ Irreducible \Rightarrow fraction of time in state i approaches $\pi(i)$
- ▶ Irreducible + Aperiodic $\Rightarrow \pi_n \rightarrow \pi$.

Summary

Markov Chains

- ▶ Markov Chain: $Pr[X_{n+1} = j | X_0, \dots, X_n = i] = P(i, j)$
- ▶ FSE: $\beta(i) = 1 + \sum_j P(i, j)\beta(j)$; $\alpha(i) = \sum_j P(i, j)\alpha(j)$.
- ▶ $\pi_n = \pi_0 P^n$
- ▶ π is invariant iff $\pi P = \pi$
- ▶ Irreducible \Rightarrow one and only one invariant distribution π
- ▶ Irreducible \Rightarrow fraction of time in state i approaches $\pi(i)$
- ▶ Irreducible + Aperiodic $\Rightarrow \pi_n \rightarrow \pi$.
- ▶ Calculating π : One finds $\pi = [0, 0, \dots, 1]Q^{-1}$ where $Q = \dots$.